

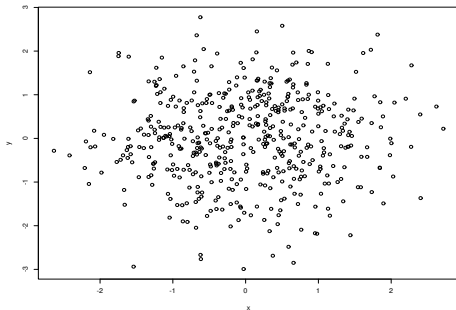


## A Short Course in Linear and Logistic Regression

Suman Guha  
Assistant Professor  
Department of Statistics  
Presidency University, Kolkata  
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- Observations taken on two features - say height ( $x$ ) and weight ( $y$ ) of individuals.
- Situations when ( $x$ ) and ( $y$ ) show no interrelationships - no point doing regression.

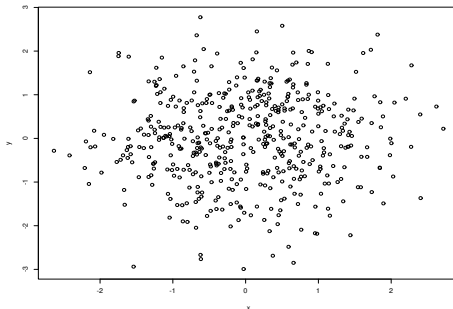


(a)

Figure: Artificially simulated dataset showing no dependence between  $x$  and  $y$ .



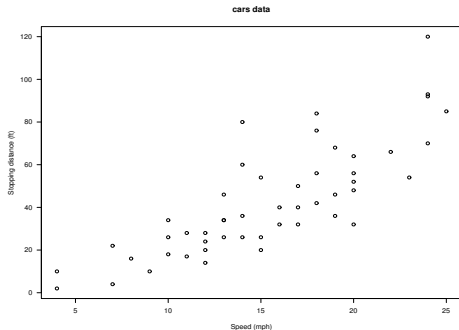
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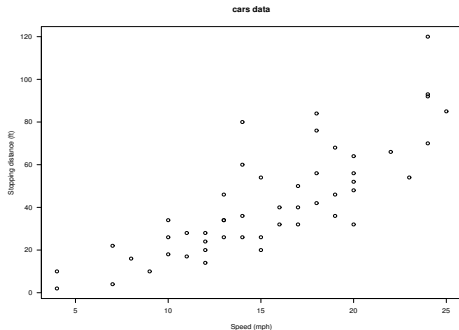


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Figure: (x) car speed in miles per hour vs (y) stopping distance in feet.

- Want an approximate formula ( $y \approx f(x)$ ) of stopping distance ( $y$ ) in terms of car speed ( $x$ ) - regression problem.
- Why?

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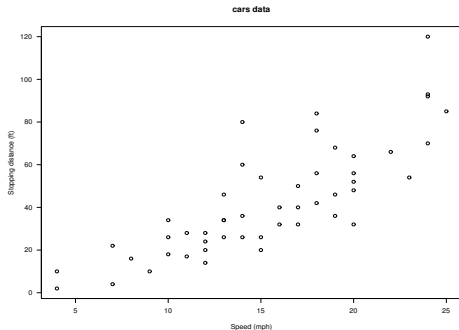
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- To understand the nature of dependence between  $(x)$  and  $(y)$ .
- Sometimes  $(y)$  may be costly/difficult to measure (total annual income) but  $(x)$  may be measured easily (total annual expenditure) - can use the formula to predict  $y^*$  using  $x^*$ .
- What type of formula? -  $f(x) = ax^3 + b\sqrt{x} + c$ ?
- No, we want a formula of form  $f(x) = a + bx$  - equation of a straight line.
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Reason?(i) Mathematically simple.

(ii) Most of the time linear regression perform quite well!

- How to get the value of  $a, b$ ? - a line that pass through the most middle - obtained by minimizing  $\sum_{i=1}^n (y_i - a - bx_i)^2$ .
- Closed form solution available -  

$$f(x) = (\bar{y} - \frac{\text{Cov}(x,y)}{\text{Var}(x)} \bar{x}) + \frac{\text{Cov}(x,y)}{\text{Var}(x)} x = \bar{y} + \frac{\text{Cov}(x,y)}{\text{Var}(x)} (x - \bar{x}).$$



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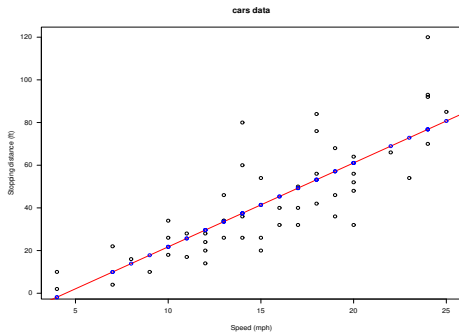
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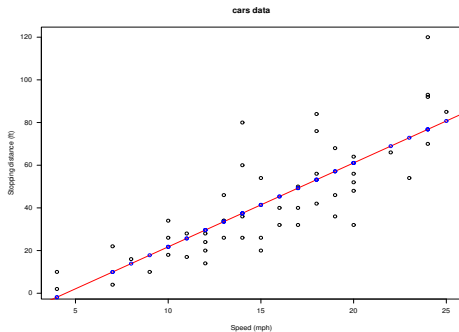


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Figure: Scatter plot with the regression line, fitted values and residuals.

- Minimizing  $\sum_{i=1}^n (y_i - a - bx_i)^2$  wrt  $a, b$  - Principle of least squares (LS) - LS regression line.
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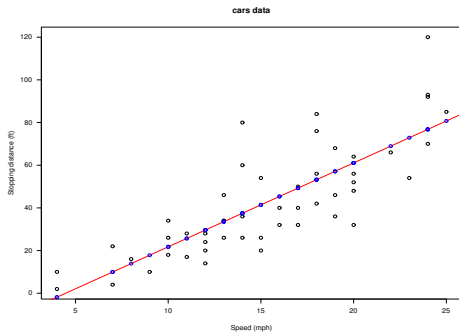


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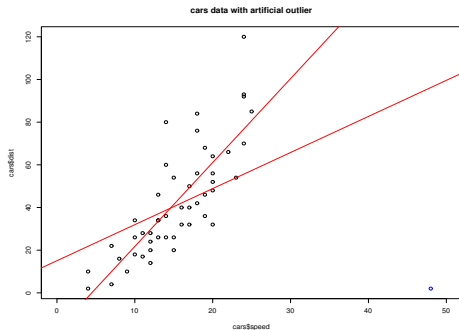


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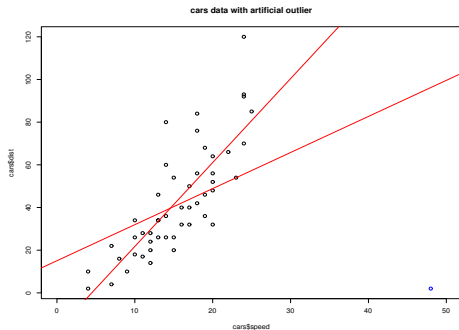




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Figure: Effect of a single outlier on LS regression line.

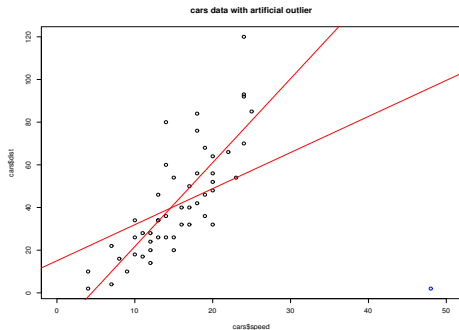
- Two possibilities : (i) detect and drop the outlier (ii) apply an outliers resistant regression.
- Minimizing  $\sum_{i=1}^n (y_i - a - bx_i)^2$  wrt  $a, b$  equivalent minimizing  $\frac{1}{n} \sum_{i=1}^n (y_i - a - bx_i)^2$  (mean of  $(y_i - a - bx_i)^2$ ) wrt  $a, b$ .
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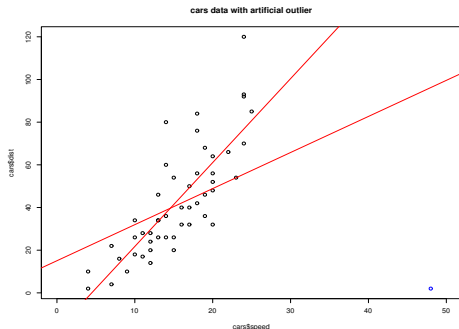
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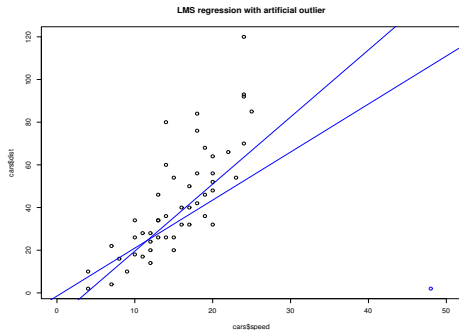


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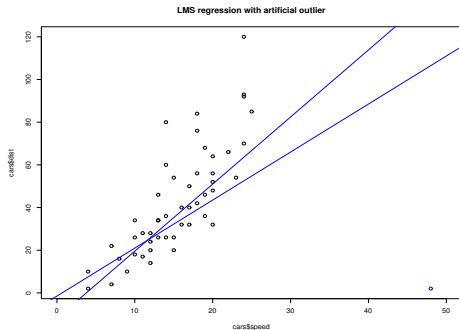


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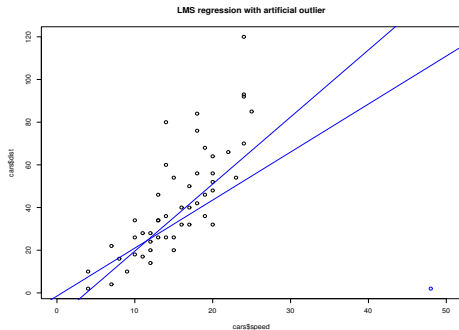


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- Model parameters -  $a, b, \sigma_\epsilon$ .
- The model looks unfamiliar?
- The model is nothing but a family of MVN distributions indexed by unknown parameters  $a, b, \sigma_\epsilon$ .
- More familiar specification -  $\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$ ;  $\boldsymbol{\epsilon} \sim MVN(\mathbf{0}, \sigma_\epsilon^2 \mathbf{I}_n)$ .

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- The model is fitted using maximum likelihood method.
- Inferential goal - estimating  $\beta$  and  $\sigma_\epsilon^2$ .
- mle of  $\beta$  is given by  $\hat{\beta} = \mathbf{Q}_X \mathbf{y} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$  - same as LS regression values.
- mle of  $\sigma_\epsilon^2$  is given by  $\hat{\sigma}_\epsilon^2 = \frac{\sum_{i=1}^n e_i^2}{n}$  - biased.
- An unbiased estimator  $\tilde{\sigma}_\epsilon^2 = \frac{\sum_{i=1}^n e_i^2}{n-2}$ .
- Only concentrate on  $\hat{\beta}$  from now on.
- How good/reliable are these estimates? - calculate standard errors.
- $Var(\hat{\beta}) = Var(\mathbf{Q}_X \mathbf{y}) = \mathbf{Q}_X Var(\mathbf{y}) \mathbf{Q}_X' = \mathbf{Q}_X \sigma_\epsilon^2 \mathbf{I}_n \mathbf{Q}_X' = \sigma_\epsilon^2 \mathbf{Q}_X \mathbf{Q}_X' = \sigma_\epsilon^2 (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{X} (\mathbf{X}'\mathbf{X})^{-1} = \sigma_\epsilon^2 (\mathbf{X}'\mathbf{X})^{-1}$ .
- Estimate of  $Var(\hat{\beta})$  is  $\tilde{\sigma}_\epsilon^2 (\mathbf{X}'\mathbf{X})^{-1}$  (we use the unbiased estimator  $\tilde{\sigma}_\epsilon^2$  not mle  $\hat{\sigma}_\epsilon^2$ ).
- Its diagonal entries - estimate of standard error  $\widehat{se}(\hat{\beta}_0)$  and  $\widehat{se}(\hat{\beta}_1)$ .



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(i) Linearity: The relationship between  $\mathbf{X}$  and the mean of  $\mathbf{Y}$  is linear  
( $E(\mathbf{Y}|\mathbf{X}) = \mathbf{X}\beta$ ).

(ii) Homoscedasticity: The variance of residual is the same for  $x_1, x_2, \dots, x_n$ .

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■ Normality: For any fixed value  $x_i$ ,  $[Y_i|X_i = x_i]$  is normally distributed.

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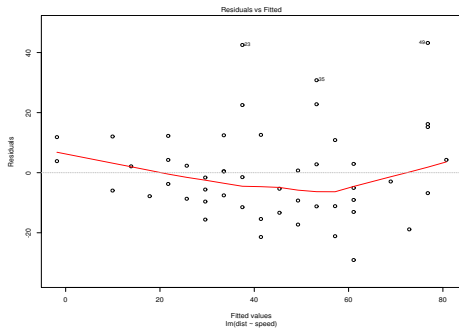
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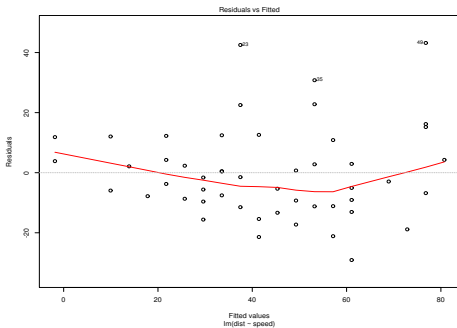


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Figure: Clear indication of nonlinearity and heteroscedasticity.



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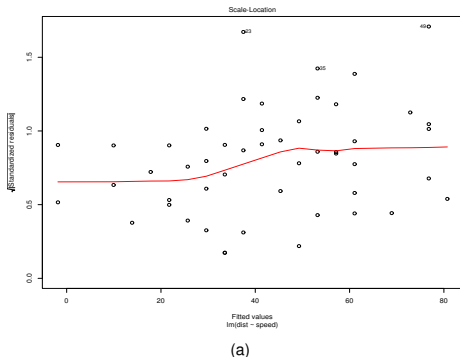


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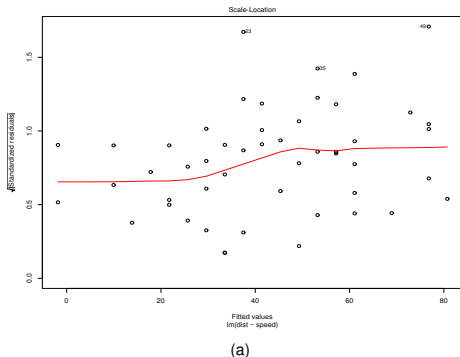


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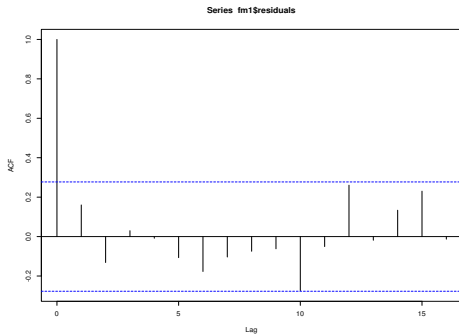




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- $\mathbf{P_X}$  (some refers it as hat-matrix  $\mathbf{H}$ ) is an orthogonal projection matrix - idempotent and symmetric - also,  $\hat{\mathbf{y}} = \mathbf{P_X y}$ .
- One can use Breusch-Pagan Test for checking homoscedasticity - asymptotically  $\chi^2$  distributed.
- Uncorrelatedness: Plot the sample autocorrelation function of the residuals.



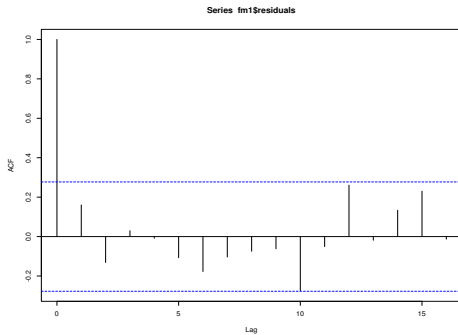
- $Var(e_i) = \sigma_\epsilon^2(1 - h_{ii})$  - so,  $\widehat{Var}(e_i) = \tilde{\sigma}_\epsilon^2(1 - h_{ii})$ .
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Figure: Indication of uncorrelatedness.

- Also can perform Durbin-Watson test and Box-Pierce test for checking whether there is any autocorrelation.



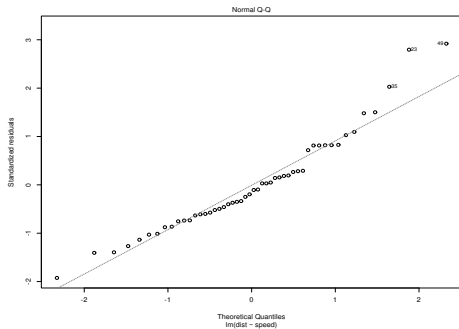
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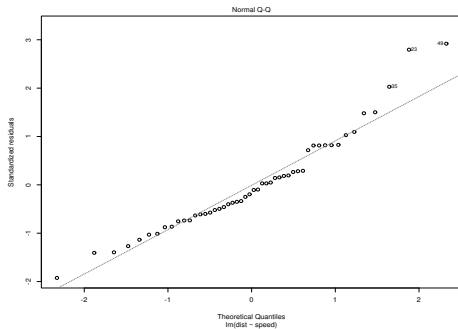
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Figure: Indication of non-normality.

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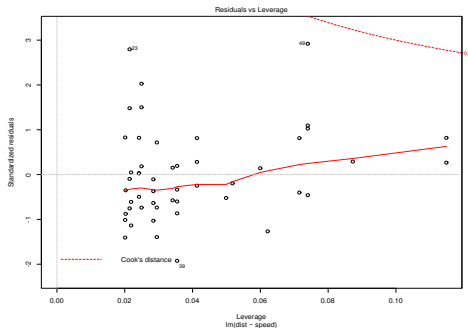
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- Outliers : Check the fitted value  $\hat{Y}_i$  vs residual  $e_i$  plot for large values - potential outliers.
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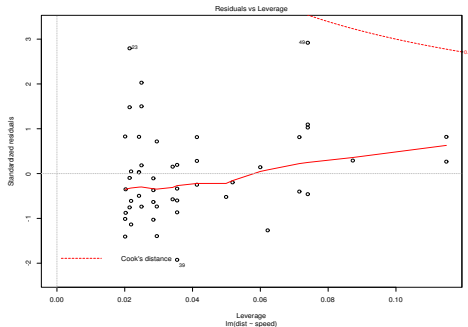


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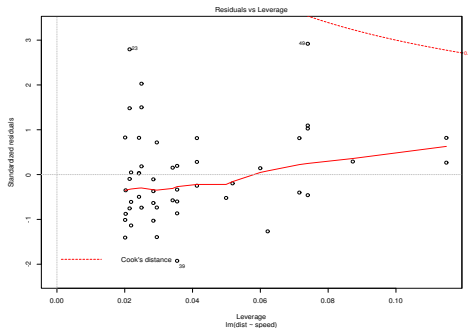
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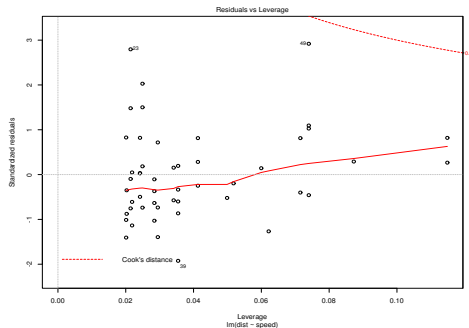


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- Also, some more numerical diagnostic measures are there for detection of potentially influential observations.

- Cook's distance :

$$D_i = \frac{1}{2} \left( \frac{e_i}{\hat{\sigma}_e \sqrt{1-h_{ii}}} \right)^2 \frac{h_{ii}}{1-h_{ii}} = \frac{1}{2} (\text{standardized residual})^2 \frac{h_{ii}}{1-h_{ii}}.$$

- So, Cook's D is a function of studentized residual and leverage value - can be plotted as a nonlinear contours in the residuals vs leverage plot.
- High leverage values (close to 1) means Cook's distance very large - **highly influential observation**.
- DFFIT :  $DFFIT_i$  = difference in fit as we drop the  $i$ th observation.
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- These measures are useful in selection of a single best model among several competing models.
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- Problem of  $R^2$  - tend to select overfitting models.
- Adjusted R-squared -  $R_{adj}^2 = 1 - \frac{(n-1)(1-R^2)}{(n-2)}$  - higher the better - can be negative!
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More familiar specification -  $\mathbf{Y} = \mathbf{X}\beta + \epsilon$ ;  $\epsilon \sim MVN(\mathbf{0}, \sigma_\epsilon^2 \mathbf{I}_n)$ .

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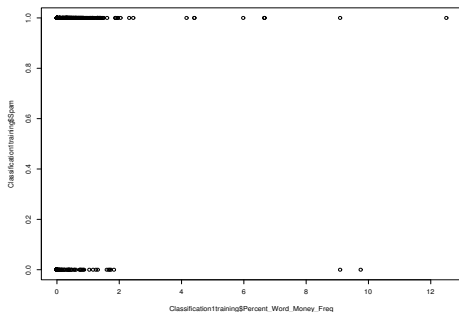
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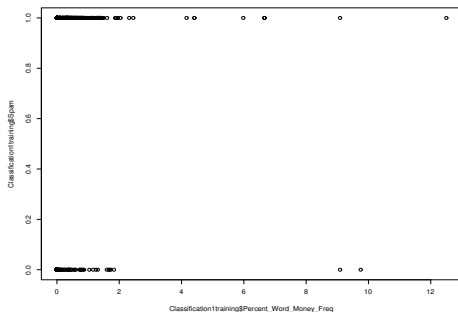


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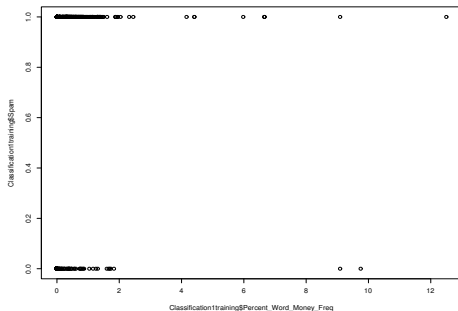
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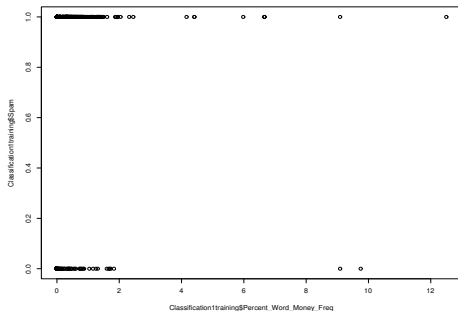
Figure: Scatter plot of  $x$  and  $y(0/1)$  - not useful.

- Not much of descriptive statistics can be done.
- Still - need some motivation!





- Observations taken on two features - covariate is continuous say dosage of a drug ( $x$ ) and response ( $y$ ) is binary subject is alive/dead (we code it as 0/1).
- Scatter plot of ( $x$ ) and ( $y$ ) does not give much insight!



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Figure: Scatter plot of  $x$  and  $y(0/1)$  - not useful.

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- In simple linear regression model we have assumption  $E(Y|X = x) = a + bx$ .
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  - (a) spam detection based on certain words and characters.
  - (b) malignant tumor detection based on certain cell profiles.
  - (c) loan defaulters detection based on personal/socio-economic and demographic profiles.
- Difference with linear regression - no closed form solution available.
- Simple logistic regression model :

$$\begin{aligned}
 & [Y_1 = y_1, \dots, Y_n = y_n | X_1 = x_1, \dots, X_n = x_n] \sim \\
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- Model parameters -  $a, b$ .
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- mle of  $\beta$  is denoted by  $\hat{\beta}$  - unlike linear regression **no closed form expression**.
- mle is calculated using numerical algorithm - Fisher's scoring algorithm.
- Often the algorithm may not converge - **multicollinearity, sparseness and complete separation**.
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- Individual test of significance  $H_0 : \beta_0 = 0$  vs  $H_1 : \beta_0 \neq 0$  (test of intercept).
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- Model parameters -  $\beta_0, \beta_1, \beta_2, \dots, \beta_p$ .
- Everything is same as simple logistic regression.
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- Perform a multiple linear regression of the  $j$  th covariate on the remaining  $(p - 1)$  covariates - calculate the  $R_j^2$  (R-squared).
- $VIF_j = \frac{1}{1 - R_j^2}$
- High VIF means highly correlated covariate -  $VIF_j > 5$  is high (thumb rule).
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