BSAN 400 Introduction to Machine Learning Lecture 3. Linear Regression and Variable Selection¹

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¹Partially based on Hastie, et al. (2009) ESL, and James, et al. (2013) ISLR

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- Supervised learning method
- It assumes the dependence of Y on X is linear
- Largely used in many disciplines
- Simple and interpretable
- Fundamental in data science

Simple linear regression

$$Y = \beta_0 + \beta_1 X + \epsilon$$

Multiple linear regression

$$Y = \beta_0 + \beta_1 X + \ldots + \beta_p X_p + \epsilon$$

- Y: dependent variable (response, outcome)
- X's: independent variable (covariates, explanatory variable)
- \blacksquare β 's: regression coefficients
- *ϵ*: random error (irreducible error)

Using matrix format

$$Y = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$$

X is called design matrix with first column being 1's

■ The estimated linear regression model is

$$\hat{Y} = \mathbb{E}(Y|X) = \mathbf{X}\hat{oldsymbol{eta}}$$

• Goal: estimate regression coefficient β

- $\mathbb{E}(Y|X)$ is a linear function of X or its basis expansion such as X_1^2, X_2^3, \ldots
- The error term $\{\epsilon_i, \ldots, \epsilon_n\} \overset{i.i.d.}{\sim} N(0, \sigma^2)$

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Least Square Solution



We want to minimize residual sum squares

$$RSS(\beta) = \sum_{i=1}^{n} (y_i - \mathbf{x}_i^{\mathsf{T}} \beta)^2$$
$$= (\mathbf{y} - \mathbf{X}\beta)^{\mathsf{T}} (\mathbf{y} - \mathbf{X}\beta)$$

 \blacksquare Take first-order derivative with respect to $oldsymbol{eta}$ and set to 0

$$0 = \frac{\partial RSS(\beta)}{\partial \beta} = -2\mathbf{X}^{T}(\mathbf{y} - \mathbf{X}\beta)$$

$$\mathbf{x}^{T}\mathbf{y} = \mathbf{x}^{T}\mathbf{X}\beta \qquad \beta = (\mathbf{x}^{T}\mathbf{X})^{-1}\mathbf{X}^{T}\mathbf{y}$$

This is called *normal equation*.

$$\mathbf{\beta} = (\mathbf{x}^{T}\mathbf{X})^{-1}\mathbf{X}^{T}\mathbf{y}$$

$$\mathbf{\beta} = \mathbf{x} \mathbf{y}^{T}\mathbf{x}\mathbf{y}^{T}\mathbf{y}$$

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Least Square Solution



- The predicted value is $\hat{\mathbf{y}} = \begin{bmatrix} \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} & \hat{\boldsymbol{\beta}} \end{bmatrix}$ $\mathbf{H} = \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T$ is called hat matrix or projection matrix



Model Assessment – R Square and MSE

• It is proportion of variation in Y explained by the model $R^{2} = \frac{TSS - RSS}{TSS} = 1 - \frac{RSS}{TSS}$

 \mathbb{R}^2 increases monotonically as number of explanatory variable increasing.

Adjusted <u>R²</u>

$$R_{adj}^2 = 1 - \frac{n-1}{n-p-1} \frac{RSS}{TSS}$$

Mean Squared Error (MSE)

$$MSE = \frac{1}{n - p - 1} \times RSS$$

It is an unbiased estimate of σ^2 for irreducible error ϵ .

Akaike information criterion (AIC), the smaller the better

$$AIC = -2\log(\hat{L}) + 2p$$

Bayesian information criteria (BIC), the smaller the better

$$BIC = -2\log(\hat{L}) + \log(n)p$$

where \hat{L} is estimated likelihood function. 3R55

- Mellow's C_p is the same as AIC for linear regression
- Cross-validation error (CV score)

Is a specific X relevant?

• Testing for individual β • $H_0: \beta_j = 0; H_1: \beta_j \neq 0$ • Using T-test since the true variance is unknown • f_j be(β_j) t $T = \frac{\hat{\beta}_j}{se(\hat{\beta}_j)} = \frac{\hat{\beta}_j}{\hat{\sigma}\sqrt{v_j}} \sim t_{n-p-1}$ where v_j is the *j*th diagonal element of $(\mathbf{X}^T \mathbf{X})^{-1}$ • Reject H_0 if p-value $< \alpha$ or $|T| > T_{1-\alpha}^{(n-p-1)}$

= Reject H_0 if p-value $< \alpha$ or $|T| > T_{1-\alpha}$ = Confidence interval: $\hat{\beta} \pm se(\hat{\beta}) \times T_{1-\alpha}^{(n-p-1)}$ $\beta \leftarrow \beta$

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$$F^* = \frac{(TSS - RSS)/p}{RSS/(n-p-1)} \sim F_{p,n-p-1}$$

where $TSS = \sum_{i=1}^{n} (y_i - \bar{y})^2$, is total sum squares

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- Resampling method
- A powerful tool to quantify uncertainty
 - standard error
 - confidence interval
- Random sampling with replacement
- Example:
 - train a model with 1000 bootstrap samples
 - store all the parameter estimates
 - calculate standard error and confidence interval



Model Diagnostics

- Check the assumptions on the error term
 - Independent normal distribution?
 - $\mathbb{E}(\epsilon_i) = 0$?
 - $Var(\epsilon_i) = \sigma^2 = \text{constant}?$

End N(0, 02)

Residual plot (an ideal residual plot looks like this)²



 ²source: Camm, et al., Essentials of Business Analytics, <≧ > <≧ > <≧ > <≧ > << ≥ > <≧ > << ≥ > <≧ > << ≥ > <≧ > << ≥ > <≧ > << ≥ > <≤ ≥ << ≥ > << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > <≤ ≥ << ≥ > << ≥ > <≤ ≥ << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > << ≥ > <<

Residual plot

Which type of assumption is violated?



 ³source: Camm, et al., Essentials of Business Analytics

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Normal Quantile-Quantile Plot

- It plots the standardized residual vs. theoretical quantiles
- An easy way to visually test the normality assumption
- If residual follows normal distribution, you should expect all dots lie on the diagonal straight line.
- Residual-Leverage Plot
 - This plot checks if there are any influential points, which could alter your analysis by excluding them
 - The points that lie outside the dashed line, Cook's distance, are considered as influential points

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• Why? Try to improve the model: exclude unnecessary predictors

- Interpretation and simplicity
- Prediction stability and accuracy
- Less computational cost
- Bias-variance tradeoff
- Common approaches
 - Subset selection
 - Shrinkage (also called *regularization*)
 - Dimension reduction (project *p* predictors into a M-dimensional subspace)
- Some times it is subjective, and needs domain knowledge so that certain variable has to be in the model.

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- Select the best subset of predictors such that the model is optimal in terms of a certain assessment metric
- Computationally expensive even infeasible
 leaps and bounds (an R package "leaps") algorithm makes it feasible for p as large as 30 or 40.
- Suppose there are 10 predictors. How many models need to be fitted and evaluated?

Example of Best Subset Selection



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- Computationally less expensive than best subset
- Iteratively adding or dropping one variable at a time
- Forward/backward is greedy procedure. That is, they won't adjust any added/dropped variables in previous step
- Stepwise: start with forward, and then iteratively add and drop variables
- Selection criteria: AIC or BIC
- R package: "step"
- An illustration: click here

- Also called penalized estimation
- Shrink the regression coefficients toward 0 by constraints (regularization)
- Shrinkage methods are always preferred over best subset or stepwise methods. Why?
- A game of bias-variance tradeoff
- We discuss two popular shrinkage methods:
 - Ridge regression
 - LASSO

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Ridge Regression

Recall least square. We solve the optimization

$$\hat{\boldsymbol{\beta}}_{LS} = \arg\min_{\boldsymbol{\beta}} \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2$$

• Ridge regression solves a (L_2) penalized least square

$$\hat{\boldsymbol{\beta}}_{Ridge} = \arg\min_{\boldsymbol{\beta}} \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{p} \beta_j^2$$

• λ is a tuning parameter, called shrinkage parameter

• Writing in matrix form, we can get the analytical solution

$$\hat{\boldsymbol{\beta}}_{\textit{Ridge}} = (\mathbf{X}^{T}\mathbf{X} + \lambda \mathbf{I})^{-1}\mathbf{X}^{T}\mathbf{y}$$
 (Exercise: Show it!)

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It is equivalent to solve a constrained optimization problem

$$\min \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 \quad \angle 5$$

s.t.
$$\sum_{j=1}^{p} \beta_j^2 = a \qquad \text{condrainf}$$

• a corresponds to the tuning parameter λ

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Ridge Regression Solution Path – Boston Housing Data



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- Least absolute shrinkage and selection operator (LASSO)
- Introduced by Tibshirani (1996)
- Shrinkage estimation
- It estimates the coefficients and perform variable selection simultaneously

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LASSO

• LASSO solves the (L₁) penalized least square

$$\hat{\beta}_{LASSO} = \arg \min_{\beta} \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{p} |\beta_j|$$

- It is a convex optimization problem
- It is equivalent to solve a constrained optimization problem

$$\min \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2$$

s.t.
$$\sum_{j=1}^{p} |\beta_j| = a$$
 Budget Constraint

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An Illustration



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LASSO Regression Solution Path – Boston Housing Data



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Tuning Parameter λ Selection

- λ controls the shrinkage level (different *lambda* associates with different estimated model)
- Cross-validation
 - In R, use the function cv.glm() in package glmnet



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- Number of predictor is very large (even larger than sample size)
- Ultra-high dimension $p \gg n$
- It is very common for gene expression and image data
- Sparsity assumption: only a few predictors are relevant
- OLS fails when *n* < *p*. Why?
- LASSO or similar methods provide sparse solution

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- Introduced by Zou and Hastie (2005)
- Combination of Ridge and LASSO

$$\hat{\boldsymbol{\beta}}_{EN} = \arg\min_{\boldsymbol{\beta}} \sum_{i=1}^{n} \left(y_i - \mathbf{x}_i^T \boldsymbol{\beta} \right)^2 + \lambda_1 \sum_{j=1}^{p} |\beta_j| + \lambda_2 \sum_{j=1}^{p} \beta_j^2$$

- Convex optimization
- Ridge and LASSO are special cases of Elastic Net
- It incorporates the advantages of both Ridge and LASSO
 - Ridge regression: lower variance; multicollinearity
 - LASSO: variable selection (selects at most *n* variables if p > n)

- It is recommended to standardize all predictors in shrinkage estimation. Why?
- Solution of Ridge regression is equivalent to the posterior of Bayesian estimates
- There are many other type of penalized estimators with different penalty functions that can perform variable selection.
 - Group Lasso (Yuan and Lin, 2006)
 - Adaptive-LASSO (Zou, 2006)
 - SCAD (Fan and Li, 2001)
 - MCP (Zhang, 2010)

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