

Digital Speech Processing— Lecture 13

Linear Predictive Coding (LPC)- Introduction

LPC Methods

- LPC methods are the most widely used in speech coding, speech synthesis, speech recognition, speaker recognition and verification and for speech storage
 - LPC methods provide extremely accurate estimates of speech parameters, and does it extremely efficiently
 - basic idea of Linear Prediction: current speech sample can be closely approximated as a linear combination of past samples, i.e.,

$$s(n) = \sum_{k=1}^p \alpha_k s(n-k) \text{ for some value of } p, \alpha_k \text{'s}$$

LPC Methods

- for periodic signals with period N_p , it is obvious that

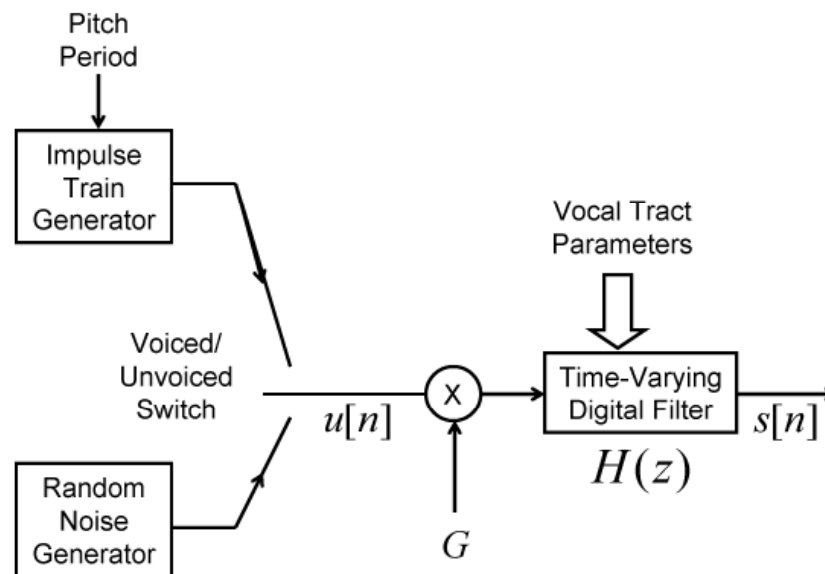
$$s(n) \approx s(n - N_p)$$

but that is not what LP is doing; it is estimating $s(n)$ from the p ($p \ll N_p$) most recent values of $s(n)$ by linearly predicting its value

- for LP, the predictor coefficients (the α_k 's) are determined (computed) by ***minimizing the sum of squared differences*** (over a finite interval) ***between the actual speech samples and the linearly predicted ones***

LPC Methods

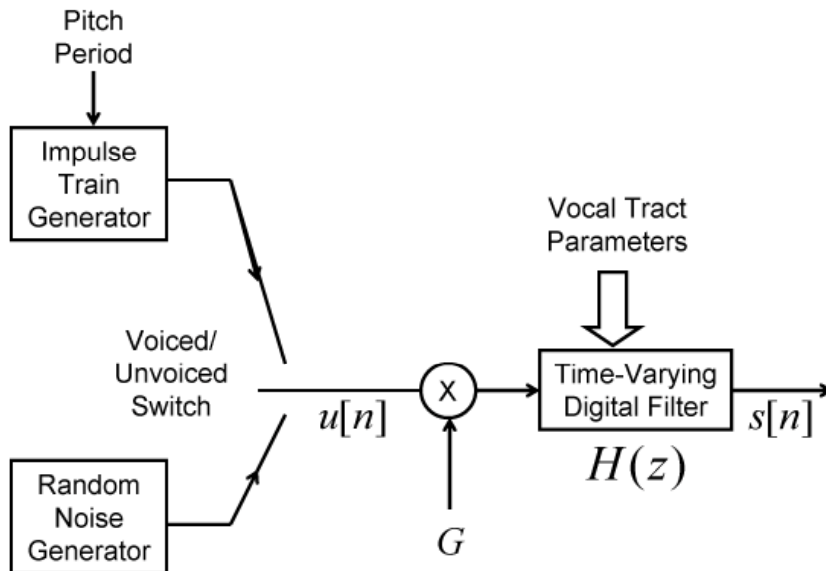
- LP is based on speech production and synthesis models
 - speech can be modeled as the output of a linear, time-varying system, excited by either quasi-periodic pulses or noise;
 - assume that the model parameters remain constant over speech analysis interval
- LP provides a **robust, reliable and accurate method** for estimating the parameters of the linear system (the combined vocal tract, glottal pulse, and radiation characteristic for voiced speech)



LPC Methods

- LP methods have been used in control and information theory—called methods of system estimation and system identification
 - used extensively in speech under group of names including
 1. covariance method
 2. autocorrelation method
 3. lattice method
 4. inverse filter formulation
 5. spectral estimation formulation
 6. maximum likelihood method
 7. inner product method

Basic Principles of LP



$$s(n) = \sum_{k=1}^p a_k s(n-k) + Gu(n)$$

$$H(z) = \frac{S(z)}{GU(z)} = \frac{1}{1 - \sum_{k=1}^p a_k z^{-k}}$$

- the time-varying digital filter represents the effects of the glottal pulse shape, the vocal tract IR, and radiation at the lips
- the system is excited by an impulse train for voiced speech, or a random noise sequence for unvoiced speech
- this 'all-pole' model is a natural representation for non-nasal voiced speech—but it also works reasonably well for nasals and unvoiced sounds

LP Basic Equations

- a p^{th} order linear predictor is a system of the form

$$\tilde{s}(n) = \sum_{k=1}^p \alpha_k s(n-k) \Leftrightarrow P(z) = \sum_{k=1}^p \alpha_k z^{-k} = \frac{\tilde{S}(z)}{S(z)}$$

- the prediction error, $e(n)$, is of the form

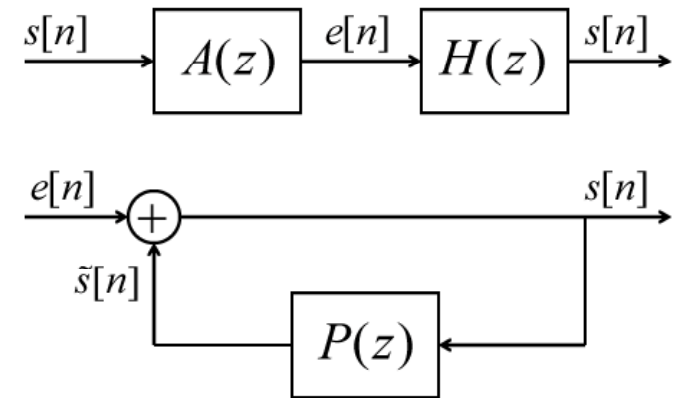
$$e(n) = s(n) - \tilde{s}(n) = s(n) - \sum_{k=1}^p \alpha_k s(n-k)$$

- the prediction error is the output of a system with transfer function

$$A(z) = \frac{E(z)}{S(z)} = 1 - P(z) = 1 - \sum_{k=1}^p \alpha_k z^{-k}$$

- if the speech signal obeys the production model exactly, and if $\alpha_k = a_k, 1 \leq k \leq p$
 $\Rightarrow e(n) = Gu(n)$ and $A(z)$ is an inverse filter for $H(z)$, i.e.,

$$H(z) = \frac{1}{A(z)}$$



LP Estimation Issues

- need to determine $\{\alpha_k\}$ directly from speech such that they give good estimates of the time-varying spectrum
 - need to estimate $\{\alpha_k\}$ from short segments of speech
 - need to minimize mean-squared prediction error over short segments of speech
 - resulting $\{\alpha_k\}$ assumed to be the actual $\{a_k\}$ in the speech production model
- => intend to show that all of this can be done ***efficiently, reliably, and accurately*** for speech

Solution for $\{\alpha_k\}$

- short-time average prediction error is defined as

$$\begin{aligned} E_{\hat{n}} &= \sum_m e_{\hat{n}}^2(m) = \sum_m (s_{\hat{n}}(m) - \tilde{s}_{\hat{n}}(m))^2 \\ &= \sum_m \left(s_{\hat{n}}(m) - \sum_{k=1}^p \alpha_k s_{\hat{n}}(m-k) \right)^2 \end{aligned}$$

- select segment of speech $s_{\hat{n}}(m) = s(m + \hat{n})$ in the vicinity of sample \hat{n}
- the key issue to resolve is the range of m for summation (to be discussed later)

Solution for $\{\alpha_k\}$

- can find values of α_k that minimize $E_{\hat{n}}$ by setting:

$$\frac{\partial E_{\hat{n}}}{\partial \alpha_i} = 0, \quad i = 1, 2, \dots, p$$

- giving the set of equations

$$-2 \sum_m s_{\hat{n}}(m-i) [s_{\hat{n}}(m) - \sum_{k=1}^p \hat{\alpha}_k s_{\hat{n}}(m-k)] = 0, \quad 1 \leq i \leq p$$

$$-2 \sum_m s_{\hat{n}}(m-i) e_{\hat{n}}(m) = 0, \quad 1 \leq i \leq p$$

- where $\hat{\alpha}_k$ are the values of α_k that minimize $E_{\hat{n}}$ (from now on just use α_k rather than $\hat{\alpha}_k$ for the optimum values)
- prediction error ($e_{\hat{n}}(m)$) is orthogonal to signal ($s_{\hat{n}}(m-i)$) for delays (i) of 1 to p

Solution for $\{\alpha_k\}$

- defining

$$\phi_{\hat{n}}(i, k) = \sum_m s_{\hat{n}}(m - i) s_{\hat{n}}(m - k)$$

- we get

$$\sum_{k=1}^p \alpha_k \phi_{\hat{n}}(i, k) = \phi_{\hat{n}}(i, 0), \quad i = 1, 2, \dots, p$$

- leading to a set of p equations in p unknowns that can be solved in an efficient manner for the $\{\alpha_k\}$

Solution for $\{\alpha_k\}$

- minimum mean-squared prediction error has the form

$$E_{\hat{n}} = \sum_m s_{\hat{n}}^2(m) - \sum_{k=1}^p \alpha_k \sum_m s_{\hat{n}}(m) s_{\hat{n}}(m-k)$$

- which can be written in the form

$$E_{\hat{n}} = \phi_{\hat{n}}(0,0) - \sum_{k=1}^p \alpha_k \phi_{\hat{n}}(0,k)$$

Process:

1. compute $\phi_{\hat{n}}(i,k)$ for $1 \leq i \leq p, 0 \leq k \leq p$
 2. solve matrix equation for α_k
- need to specify range of m to compute $\phi_{\hat{n}}(i,k)$
 - need to specify $s_{\hat{n}}(m)$

Autocorrelation Method

- assume $s_{\hat{n}}(m)$ exists for $0 \leq m \leq L - 1$ and is exactly zero everywhere else (i.e., window of length L samples) (Assumption #1)

$$s_{\hat{n}}(m) = s(m + \hat{n})w(m), \quad 0 \leq m \leq L - 1$$

- where $w(m)$ is a finite length window of length L samples



Autocorrelation Method

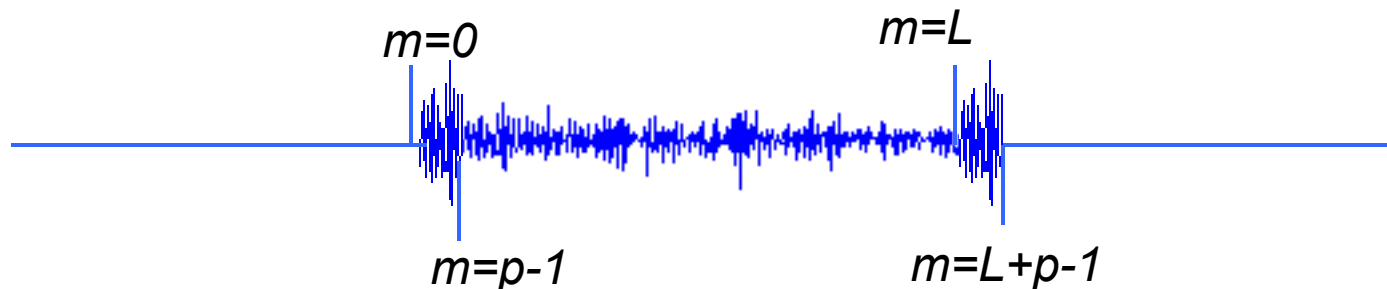
- if $s_{\hat{n}}(m)$ is non-zero only for $0 \leq m \leq L - 1$ then

$$e_{\hat{n}}(m) = s_{\hat{n}}(m) - \sum_{k=1}^p \alpha_k s_{\hat{n}}(m - k)$$

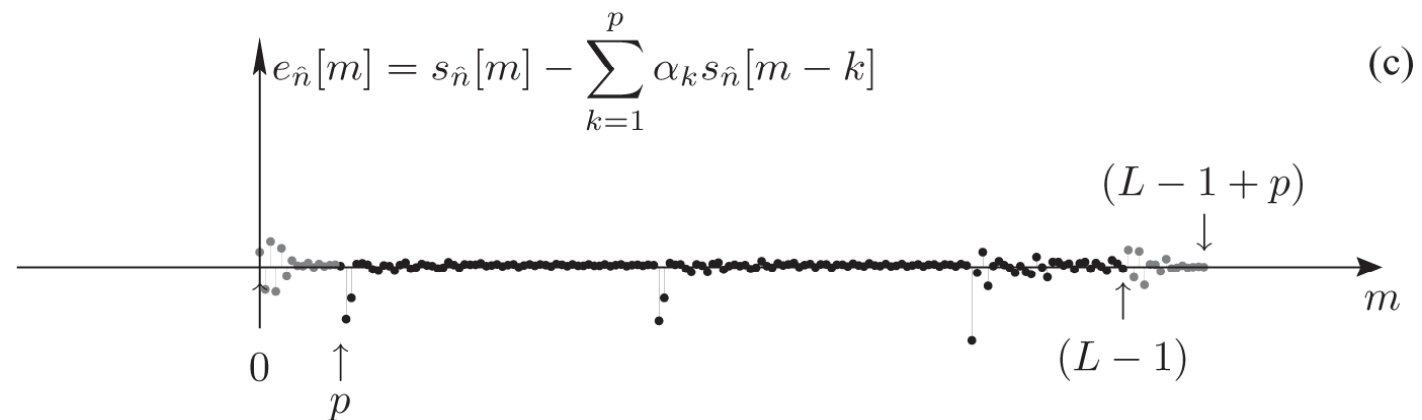
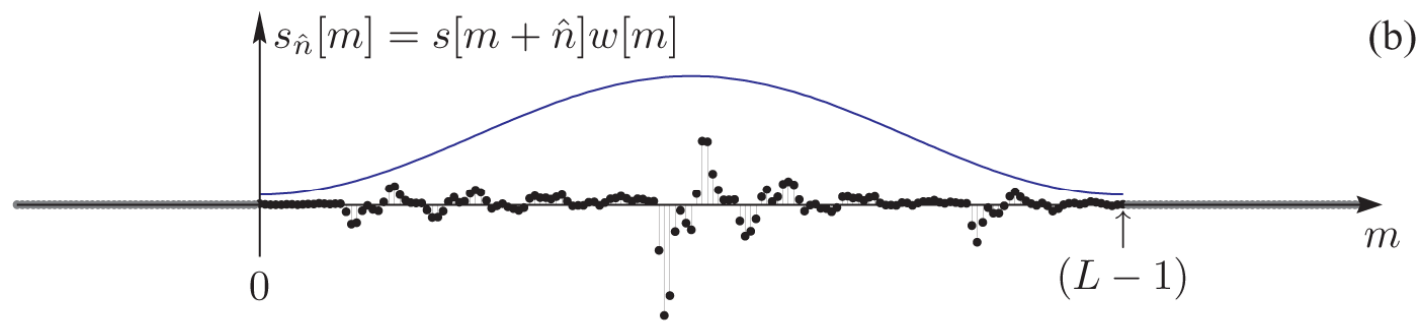
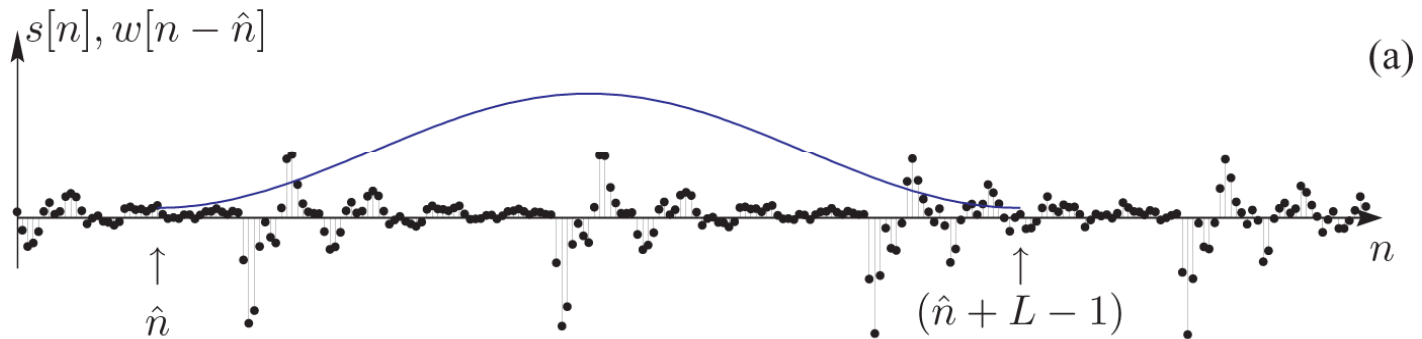
is non-zero only over the interval $0 \leq m \leq L - 1 + p$, giving

$$E_{\hat{n}} = \sum_{m=-\infty}^{\infty} e_{\hat{n}}^2(m) = \sum_{m=0}^{L-1+p} e_{\hat{n}}^2(m)$$

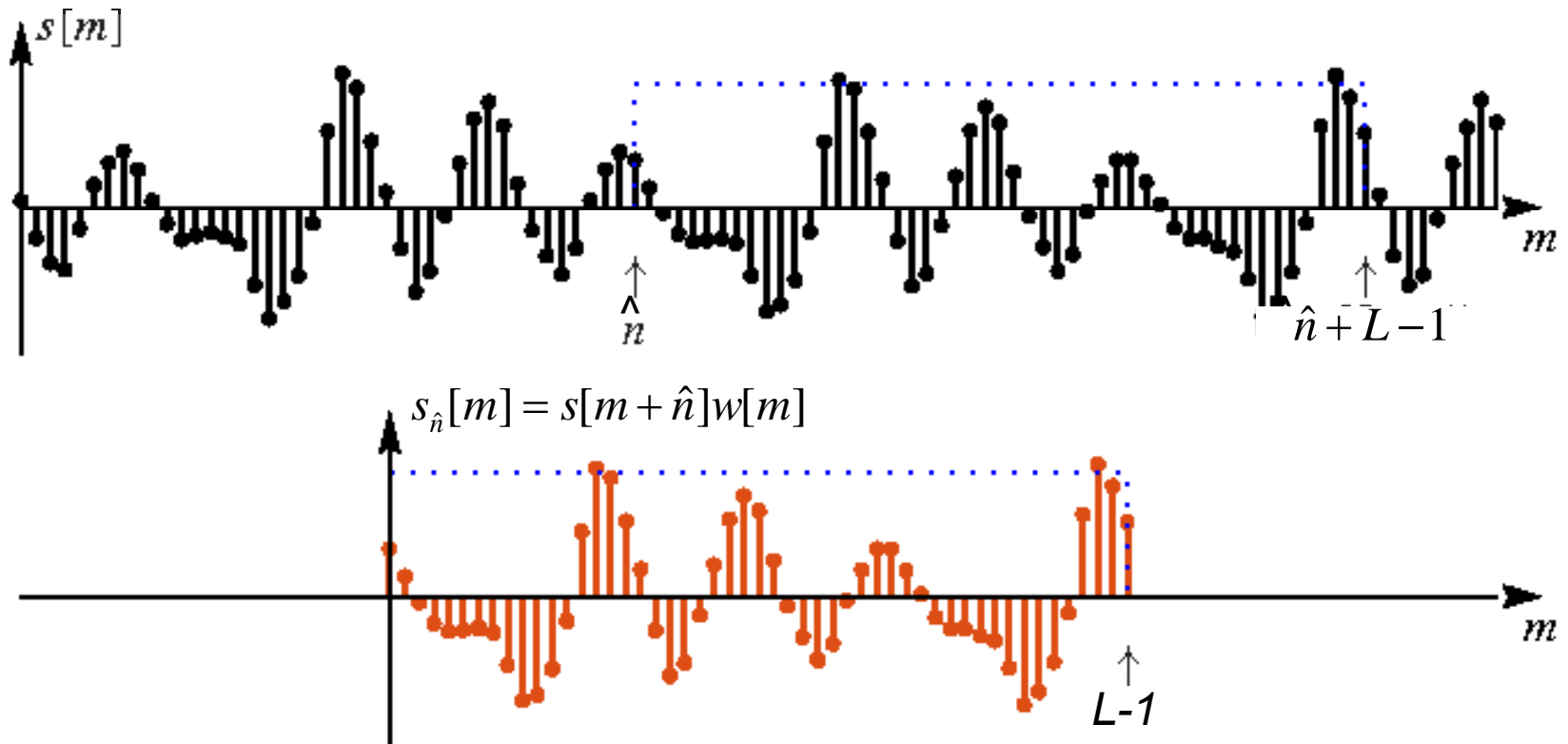
- at values of m near 0 (i.e., $m = 0, 1, \dots, p - 1$) we are predicting signal from zero-valued samples outside the window range $\Rightarrow e_{\hat{n}}(m)$ will be (relatively) large
- at values near $m = L$ (i.e., $m = L, L + 1, \dots, L + p - 1$) we are predicting zero-valued samples (outside window range) from non-zero samples $\Rightarrow e_{\hat{n}}(m)$ will be (relatively) large
- for these reasons, normally use windows that taper the segment to zero (e.g., Hamming window)



Autocorrelation Method

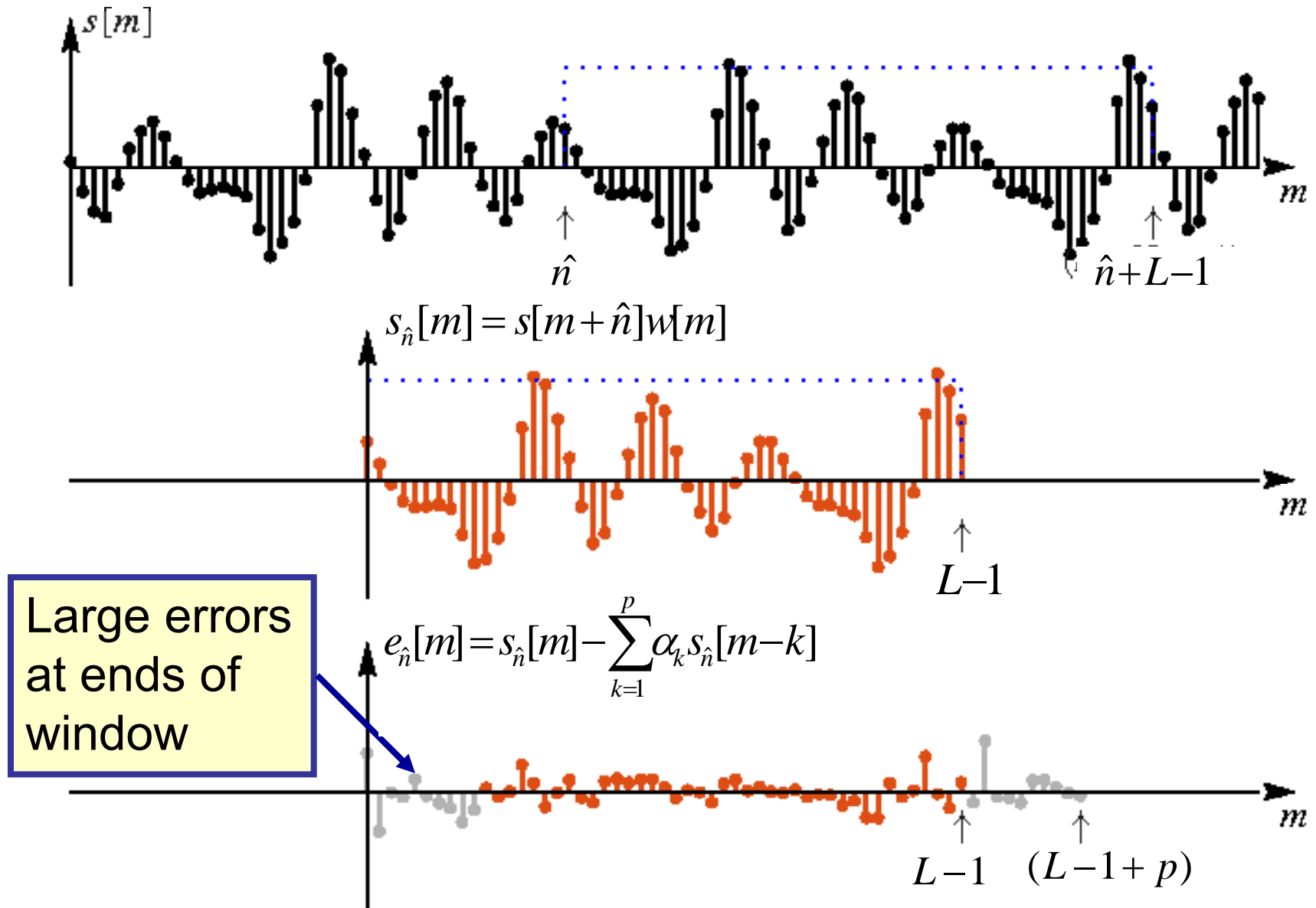


The “Autocorrelation Method”



$$R_{\hat{n}}[k] = \sum_{m=0}^{L-1-k} s_{\hat{n}}[m]s_{\hat{n}}[m+k] \quad k = 1, 2, \dots, p$$

The “Autocorrelation Method”



Autocorrelation Method

- for calculation of $\phi_{\hat{n}}(i, k)$ since $s_{\hat{n}}(m) = 0$ outside the range $0 \leq m \leq L - 1$, then

$$\phi_{\hat{n}}(i, k) = \sum_{m=0}^{L-1+p} s_{\hat{n}}(m-i)s_{\hat{n}}(m-k), \quad 1 \leq i \leq p, 0 \leq k \leq p$$

- which is equivalent to the form

$$\phi_{\hat{n}}(i, k) = \sum_{m=0}^{L-1+(i-k)} s_{\hat{n}}(m)s_{\hat{n}}(m+i-k), \quad 1 \leq i \leq p, 0 \leq k \leq p$$

- there are $L - |i - k|$ non-zero terms in the computation of $\phi_{\hat{n}}(i, k)$ for each value of i and k ; can easily show that

$$\phi_{\hat{n}}(i, k) = f(i - k) = R_{\hat{n}}(i - k), \quad 1 \leq i \leq p, 0 \leq k \leq p$$

- where $R_{\hat{n}}(i - k)$ is the short-time autocorrelation of $s_{\hat{n}}(m)$ evaluated at $i - k$ where

$$R_{\hat{n}}(k) = \sum_{m=0}^{L-1-k} s_{\hat{n}}(m)s_{\hat{n}}(m+k)$$

Autocorrelation Method

- since $R_{\hat{n}}(k)$ is even, then

$$\phi_{\hat{n}}(i, k) = R_{\hat{n}}(|i - k|), \quad 1 \leq i \leq p, \quad 0 \leq k \leq p$$

- thus the basic equation becomes

$$\sum_{k=1}^p \alpha_k \phi_{\hat{n}}(i - k) = \phi_{\hat{n}}(i, 0), \quad 1 \leq i \leq p$$

$$\sum_{k=1}^p \alpha_k R_{\hat{n}}(|i - k|) = R_{\hat{n}}(i), \quad 1 \leq i \leq p$$

- with the minimum mean-squared prediction error of the form

$$\begin{aligned} E_{\hat{n}} &= \phi_{\hat{n}}(0, 0) - \sum_{k=1}^p \alpha_k \phi_{\hat{n}}(0, k) \\ &= R_{\hat{n}}(0) - \sum_{k=1}^p \alpha_k R_{\hat{n}}(k) \end{aligned}$$

Autocorrelation Method

- as expressed in matrix form

$$\begin{bmatrix} R_{\hat{n}}(0) & R_{\hat{n}}(1) & \cdot & \cdot & R_{\hat{n}}(p-1) \\ R_{\hat{n}}(1) & R_{\hat{n}}(0) & \cdot & \cdot & R_{\hat{n}}(p-2) \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ R_{\hat{n}}(p-1) & R_{\hat{n}}(p-2) & \cdot & \cdot & R_{\hat{n}}(0) \end{bmatrix} \begin{bmatrix} \alpha_1 \\ \alpha_2 \\ \cdot \\ \cdot \\ \alpha_p \end{bmatrix} = \begin{bmatrix} R_{\hat{n}}(1) \\ R_{\hat{n}}(2) \\ \cdot \\ \cdot \\ R_{\hat{n}}(p) \end{bmatrix}$$

$$\mathfrak{R}\alpha = r$$

with solution

$$\alpha = \mathfrak{R}^{-1}r$$

- \mathfrak{R} is a $p \times p$ Toeplitz Matrix \Rightarrow symmetric with all diagonal elements equal
 \Rightarrow there exist more efficient algorithms to solve for $\{\alpha_k\}$ than simple matrix inversion

Covariance Method

- there is a second basic approach to defining the speech segment $s_{\hat{n}}(m)$ and the limits on the sums, namely **fix the interval** over which the mean-squared error is computed, giving **(Assumption #2)**:

$$E_{\hat{n}} = \sum_{m=0}^{L-1} e_{\hat{n}}^2(m) = \sum_{m=0}^{L-1} \left[s_{\hat{n}}(m) - \sum_{k=1}^p \alpha_k s_{\hat{n}}(m-k) \right]^2$$
$$\phi_{\hat{n}}(i, k) = \sum_{m=0}^{L-1} s_{\hat{n}}(m-i) s_{\hat{n}}(m-k), \quad 1 \leq i \leq p, \quad 0 \leq k \leq p$$

Covariance Method

- changing the summation index gives

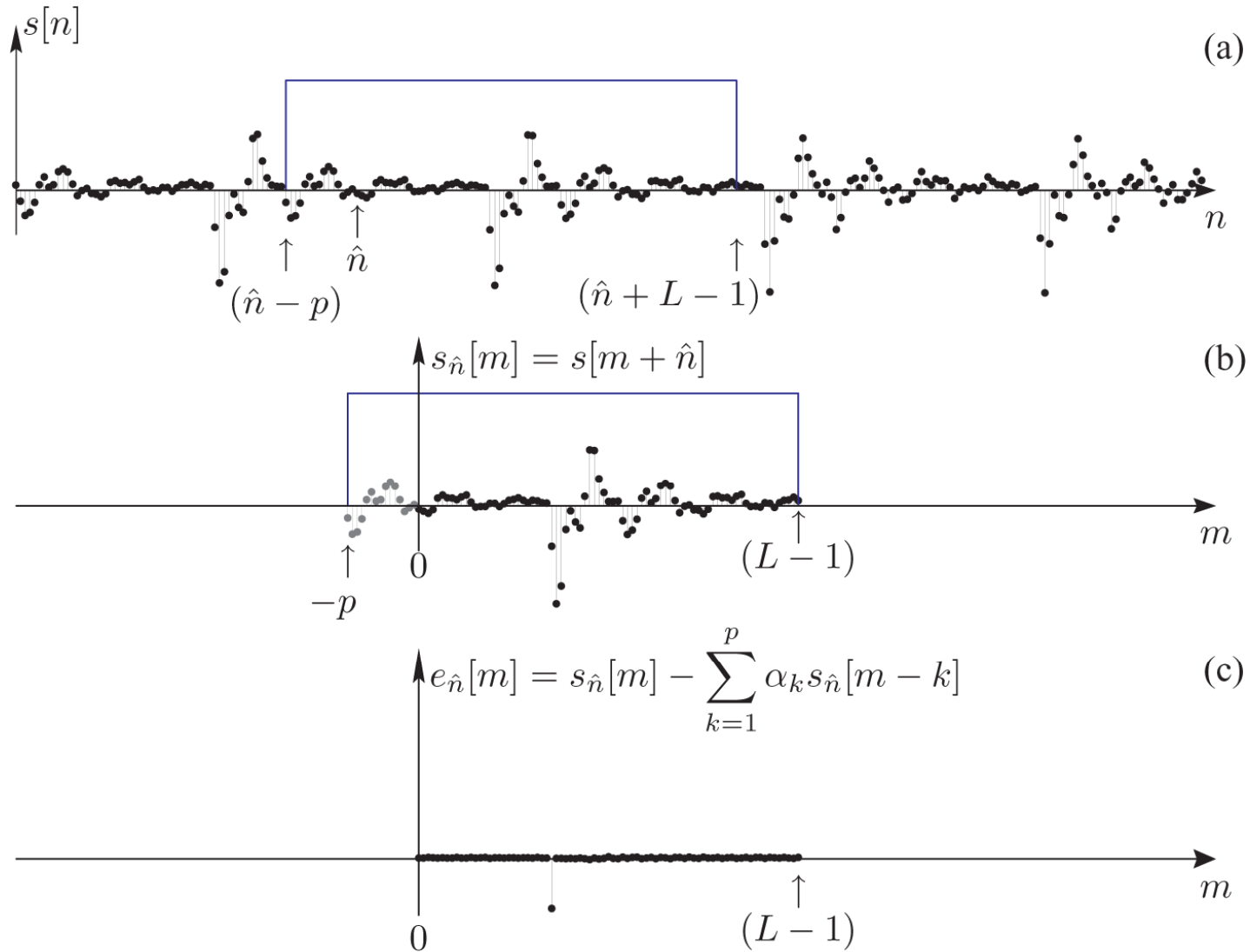
$$\phi_{\hat{n}}(i, k) = \sum_{m=-i}^{L-i-1} s_{\hat{n}}(m) s_{\hat{n}}(m+i-k), \quad 1 \leq i \leq p, \quad 0 \leq k \leq p$$

$$\phi_{\hat{n}}(i, k) = \sum_{m=-k}^{L-k-1} s_{\hat{n}}(m) s_{\hat{n}}(m+k-i), \quad 1 \leq i \leq p, \quad 0 \leq k \leq p$$

- key difference from Autocorrelation Method is that limits of summation include terms before $m = 0 \Rightarrow$ window extends p samples backwards from $s(\hat{n} - p)$ to $s(\hat{n} + L - 1)$
- since we are extending window backwards, don't need to taper it using a HW- since there is ***no transition at window edges***



Covariance Method



Covariance Method

- cannot use autocorrelation formulation => this is a true cross correlation
- need to solve set of equations of the form

$$\sum_{k=1}^p \alpha_k \phi_{\hat{n}}(i, k) = \phi_{\hat{n}}(i, 0), \quad i = 1, 2, \dots, p,$$

$$E_{\hat{n}} = \phi_{\hat{n}}(0, 0) - \sum_{k=1}^p \alpha_k \phi_{\hat{n}}(0, k)$$

$$\begin{bmatrix} \phi_{\hat{n}}(1,1) & \phi_{\hat{n}}(1,2) & \cdot & \cdot & \phi_{\hat{n}}(1,p) \\ \phi_{\hat{n}}(2,1) & \phi_{\hat{n}}(2,2) & \cdot & \cdot & \phi_{\hat{n}}(2,p) \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \phi_{\hat{n}}(p,1) & \phi_{\hat{n}}(p,2) & \cdot & \cdot & \phi_{\hat{n}}(p,p) \end{bmatrix} \begin{bmatrix} \alpha_1 \\ \alpha_2 \\ \cdot \\ \cdot \\ \alpha_p \end{bmatrix} = \begin{bmatrix} \phi_{\hat{n}}(1,0) \\ \phi_{\hat{n}}(2,0) \\ \cdot \\ \cdot \\ \phi_{\hat{n}}(p,0) \end{bmatrix}$$

$$\phi \alpha = \psi \quad \text{or} \quad \alpha = \phi^{-1} \psi$$

Covariance Method

- we have $\phi_{\hat{n}}(i, k) = \phi_{\hat{n}}(k, i) \Rightarrow$ symmetric but not Toeplitz matrix
whose diagonal elements are related as

$$\phi_{\hat{n}}(i + 1, k + 1) = \phi_{\hat{n}}(i, k) + s_{\hat{n}}(-i - 1)s_{\hat{n}}(-k - 1) - s_{\hat{n}}(L - 1 - i)s_{\hat{n}}(L - 1 - k)$$

$$\phi_{\hat{n}}(2, 2) = \phi_{\hat{n}}(1, 1) + s_{\hat{n}}(-2)s_{\hat{n}}(-2) - s_{\hat{n}}(L - 2)s_{\hat{n}}(L - 2)$$

- all terms $\phi_{\hat{n}}(i, k)$ have a fixed number of terms contributing to the computed values (L terms)
- $\phi_{\hat{n}}(i, k)$ is a covariance matrix \Rightarrow specialized solution for $\{\alpha_k\}$
called the Covariance Method

Summary of LP

- use p^{th} order linear predictor to predict $s(\hat{n})$ from p previous samples
- minimize mean-squared error, $E_{\hat{n}}$, over analysis window of duration L -samples
- solution for optimum predictor coefficients, $\{\alpha_k\}$, is based on solving a matrix equation
=> two solutions have evolved
 - **autocorrelation method** => signal is windowed by a tapering window in order to minimize discontinuities at beginning (predicting speech from zero-valued samples) and end (predicting zero-valued samples from speech samples) of the interval; the matrix $\phi_{\hat{n}}(i,k)$ is shown to be an autocorrelation function; the resulting autocorrelation matrix is Toeplitz and can be readily solved using standard matrix solutions
 - **covariance method** => the signal is extended by p samples outside the normal range of $0 \leq m \leq L - 1$ to include p samples occurring prior to $m = 0$; this eliminates large errors in computing the signal from values prior to $m = 0$ (they are available) and eliminates the need for a tapering window; resulting matrix of correlations is symmetric but not Toeplitz
=> different method of solution with somewhat different set of optimal prediction coefficients, $\{\alpha_k\}$

LPC Summary

1. Speech Production Model:

$$s(n) = \sum_{k=1}^p a_k s(n-k) + Gu(n)$$

$$H(z) = \frac{S(z)}{GU(z)} = \frac{1}{1 - \sum_{k=1}^p a_k z^{-k}}$$

2. Linear Prediction Model:

$$\tilde{s}(\hat{n}) = \sum_{k=1}^p \alpha_k s(\hat{n}-k)$$

$$P(z) = \frac{\tilde{S}(z)}{S(z)} = \sum_{k=1}^p \alpha_k z^{-k}$$

$$e(\hat{n}) = s(\hat{n}) - \tilde{s}(\hat{n}) = s(\hat{n}) - \sum_{k=1}^p \alpha_k s(\hat{n}-k)$$

$$A(z) = \frac{E(z)}{S(z)} = 1 - \sum_{k=1}^p \alpha_k z^{-k}$$

LPC Summary

3. LPC Minimization:

$$\begin{aligned} E_{\hat{n}} &= \sum_m e_{\hat{n}}^2(m) = \sum_m [s_{\hat{n}}(m) - \tilde{s}_{\hat{n}}(m)]^2 \\ &= \sum_m \left[s_{\hat{n}}(m) - \sum_{k=1}^p \alpha_k s_{\hat{n}}(m-k) \right]^2 \end{aligned}$$

$$\frac{\partial E_{\hat{n}}}{\partial \alpha_i} = 0, \quad i = 1, 2, \dots, p$$

$$\sum_m s_{\hat{n}}(m-i)s_{\hat{n}}(m) = \sum_{k=1}^p \alpha_k \sum_m s_{\hat{n}}(m-i)s_{\hat{n}}(m-k)$$

$$\phi_{\hat{n}}(i, k) = \sum_m s_{\hat{n}}(m-i)s_{\hat{n}}(m-k)$$

$$\sum_{k=1}^p \alpha_k \phi_{\hat{n}}(i, k) = \phi_{\hat{n}}(i, 0), \quad i = 1, 2, \dots, p$$

$$E_{\hat{n}} = \phi_{\hat{n}}(0, 0) - \sum_{k=1}^p \alpha_k \phi_{\hat{n}}(0, k)$$

LPC Summary

4. Autocorrelation Method:

$$s_{\hat{n}}(m) = s(m + \hat{n})w(m), \quad 0 \leq m \leq L - 1$$

$$e_{\hat{n}}(m) = s_{\hat{n}}(m) - \sum_{k=1}^p \alpha_k s_{\hat{n}}(m - k), \quad 0 \leq m \leq L - 1 + p$$

- $s_{\hat{n}}(m)$ defined for $0 \leq m \leq L - 1$; $e_{\hat{n}}(m)$ defined for $0 \leq m \leq L - 1 + p$
 \Rightarrow large errors for $0 \leq m \leq p - 1$ and for $L \leq m \leq L + p - 1$

$$E_{\hat{n}} = \sum_{m=0}^{L-1+p} e_{\hat{n}}^2(m)$$

$$\phi_{\hat{n}}(i, k) = R_{\hat{n}}(i - k) = \sum_{m=0}^{L-1-(i-k)} s_{\hat{n}}(m) s_{\hat{n}}(m + i - k) = R_{\hat{n}}(|i - k|)$$

$$\sum_{k=1}^p \alpha_k R_{\hat{n}}(|i - k|) = R_{\hat{n}}(i), \quad 1 \leq i \leq p$$

$$E_{\hat{n}} = R_{\hat{n}}(0) - \sum_{k=1}^p \alpha_k R_{\hat{n}}(k)$$

LPC Summary

4. Autocorrelation Method:

- resulting matrix equation:

$$\mathfrak{R}\alpha = r \text{ or } \alpha = \mathfrak{R}^{-1}r$$

$$\begin{bmatrix} R_{\hat{n}}(0) & R_{\hat{n}}(1) & \cdot & \cdot & R_{\hat{n}}(p-1) \\ R_{\hat{n}}(1) & R_{\hat{n}}(0) & \cdot & \cdot & R_{\hat{n}}(p-2) \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ R_{\hat{n}}(p-1) & R_{\hat{n}}(p-2) & \cdot & \cdot & R_{\hat{n}}(0) \end{bmatrix} \begin{bmatrix} \alpha_1 \\ \alpha_2 \\ \cdot \\ \cdot \\ \alpha_p \end{bmatrix} = \begin{bmatrix} R_{\hat{n}}(1) \\ R_{\hat{n}}(2) \\ \cdot \\ \cdot \\ R_{\hat{n}}(p) \end{bmatrix}$$

- matrix equation solved using Levinson or Durbin method

LPC Summary

5. Covariance Method:

- fix interval for error signal

$$E_{\hat{n}} = \sum_{m=0}^{L-1} e_{\hat{n}}^2(m) = \sum_{m=0}^{L-1} \left[s_{\hat{n}}(m) - \sum_{k=1}^p \alpha_k s_{\hat{n}}(m-k) \right]^2$$

- need signal for from $s(\hat{n} - p)$ to $s(\hat{n} + L - 1) \Rightarrow L + p$ samples

$$\sum_{k=1}^p \alpha_k \phi_{\hat{n}}(i, k) = \phi_{\hat{n}}(i, 0), i = 1, 2, \dots, p$$

$$E_{\hat{n}} = \phi_{\hat{n}}(0, 0) - \sum_{k=1}^p \alpha_k \phi_{\hat{n}}(0, k)$$

- expressed as a matrix equation:

$$\phi \alpha = \psi \quad \text{or} \quad \alpha = \phi^{-1} \psi, \quad \phi \text{ symmetric matrix}$$

$$\begin{bmatrix} \phi_{\hat{n}}(1,1) & \phi_{\hat{n}}(1,2) & \cdot & \cdot & \phi_{\hat{n}}(1,p) \\ \phi_{\hat{n}}(2,1) & \phi_{\hat{n}}(2,2) & \cdot & \cdot & \phi_{\hat{n}}(2,p) \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \phi_{\hat{n}}(p,1) & \phi_{\hat{n}}(p,2) & \cdot & \cdot & \phi_{\hat{n}}(p,p) \end{bmatrix} \begin{bmatrix} \alpha_1 \\ \alpha_2 \\ \cdot \\ \cdot \\ \alpha_p \end{bmatrix} = \begin{bmatrix} \phi_{\hat{n}}(1,0) \\ \phi_{\hat{n}}(2,0) \\ \cdot \\ \cdot \\ \phi_{\hat{n}}(p,0) \end{bmatrix}$$

Computation of Model Gain

- it is reasonable to expect the model gain, G , to be determined by matching the signal energy with the energy of the linearly predicted samples
- from the basic model equations we have

$$Gu(n) = s(n) - \sum_{k=1}^p a_k s(n-k) \Rightarrow \text{model}$$

- whereas for the prediction error we have

$$e(n) = s(n) - \sum_{k=1}^p \alpha_k s(n-k) \Rightarrow \text{best fit to model}$$

- when $\alpha_k = a_k$ (i.e., perfect match to model), then

$$e(n) = Gu(n)$$

- since it is virtually impossible to guarantee that $\alpha_k = a_k$, cannot use this simple matching property for determining the gain; instead use energy matching criterion (energy in error signal = energy in excitation)

$$G^2 \sum_{m=0}^{L-1+p} u^2(m) = \sum_{m=0}^{L-1+p} e^2(m) = E_{\hat{n}}$$

Gain Assumptions

- assumptions about excitation to solve for G
 - **voiced speech**-- $u(n) = \delta(n) \Rightarrow L$ order of a single pitch period; predictor order, p , large enough to model glottal pulse shape, vocal tract IR, and radiation
 - **unvoiced speech**-- $u(n)$ -zero mean, unity variance, stationary white noise process

Solution for Gain (Voiced)

- for voiced speech the excitation is $G\delta(n)$ with output $\tilde{h}(n)$ (since it is the IR of the system),

$$\tilde{h}(n) = \sum_{k=1}^p \alpha_k \tilde{h}(n-k) + G\delta(n); \quad \tilde{H}(z) = \frac{G}{A(z)} = \frac{G}{1 - \sum_{k=1}^p \alpha_k z^{-k}}$$

- with autocorrelation $\tilde{R}(m)$ (of the impulse response) satisfying the relation shown below

$$\tilde{R}(m) = \sum_{n=0}^{\infty} \tilde{h}(n)\tilde{h}(m+n) = \tilde{R}[-m], \quad 0 \leq m < \infty$$

$$\tilde{R}(m) = \sum_{k=1}^p \alpha_k \tilde{R}(|m-k|), \quad 1 \leq m < \infty$$

$$\tilde{R}(0) = \sum_{k=1}^p \alpha_k \tilde{R}(k) + G^2, \quad m = 0$$

Solution for Gain (Voiced)

- Since $\tilde{R}(m)$ and $R_{\hat{n}}(m)$ have the identical form, it follows that

$$\tilde{R}(m) = c \cdot R_{\hat{n}}(m), \quad 0 \leq m \leq p$$

where c is a constant to be determined.

- Since the total energies in the signal ($R(0)$) and the impulse response ($\tilde{R}(0)$) must be equal, the constant c must be 1, and we obtain the relation

$$G^2 = R_{\hat{n}}(0) - \sum_{k=1}^p \alpha_k R_{\hat{n}}(k) = E_{\hat{n}}$$

- since $\tilde{R}(m) = R_{\hat{n}}(m)$, $0 \leq m \leq p$, and the energy of the impulse response=energy of the signal => first $p + 1$ coefficients of the autocorrelation of the impulse response of the model are identical to the first $p + 1$ coefficients of the autocorrelation function of the speech signal. This condition called the *autocorrelation matching property* of the autocorrelation method.

Solution for Gain (Unvoiced)

- for unvoiced speech the input is white noise with zero mean and unity variance, i.e.,

$$E[u(n)u(n-m)] = \delta(m)$$

- if we excite the system with input $Gu(n)$ and call the output $\tilde{g}(n)$ then

$$\tilde{g}(n) = \sum_{k=1}^p \alpha_k \tilde{g}(n-k) + Gu(n)$$

- Since the autocorrelation function for the output is the convolution of the autocorrelation function of the impulse response with the autocorrelation function of the white noise input, then

$$E[\tilde{g}[n]\tilde{g}[n-m]] = \tilde{R}[m] * \delta[m] = \tilde{R}[m]$$

- letting $\tilde{R}(m)$ denote the autocorrelation of $\tilde{g}(n)$ gives

$$\begin{aligned} \tilde{R}(m) &= E[\tilde{g}(n)\tilde{g}(n-m)] = \sum_{k=1}^p \alpha_k E[\tilde{g}(n-k)\tilde{g}(n-m)] + E[Gu(n)\tilde{g}(n-m)] \\ &= \sum_{k=1}^p \alpha_k \tilde{R}(m-k), \quad m \neq 0 \end{aligned}$$

- since $E[Gu(n)\tilde{g}(n-m)] = 0$ for $m > 0$ because $u(n)$ is uncorrelated with any signal prior to $u(n)$

Solution for Gain (Unvoiced)

- for $m = 0$ we get

$$\begin{aligned}\tilde{R}(0) &= \sum_{k=1}^p \alpha_k \tilde{R}(k) + GE[u(n)\tilde{g}(n)] \\ &= \sum_{k=1}^p \alpha_k \tilde{R}(k) + G^2\end{aligned}$$

- since $E[u(n)\tilde{g}(n)] = E[u(n)(Gu(n) + \text{terms prior to } n)] = G^2$
- since the energy in the signal must equal the energy in the response to $Gu(n)$ we get

$$\tilde{R}(m) = R_{\hat{n}}(m)$$

$$G^2 = R_{\hat{n}}(0) - \sum_{k=1}^p \alpha_k R_{\hat{n}}(k) = E_{\hat{n}}$$

**Frequency
Domain
Interpretations
of Linear
Predictive
Analysis**

The Resulting LPC Model

- The final LPC model consists of the LPC parameters, $\{\alpha_k\}$, $k = 1, 2, \dots, p$, and the gain, G , which together define the system function

$$\tilde{H}(z) = \frac{G}{1 - \sum_{k=1}^p \alpha_k z^{-k}}$$

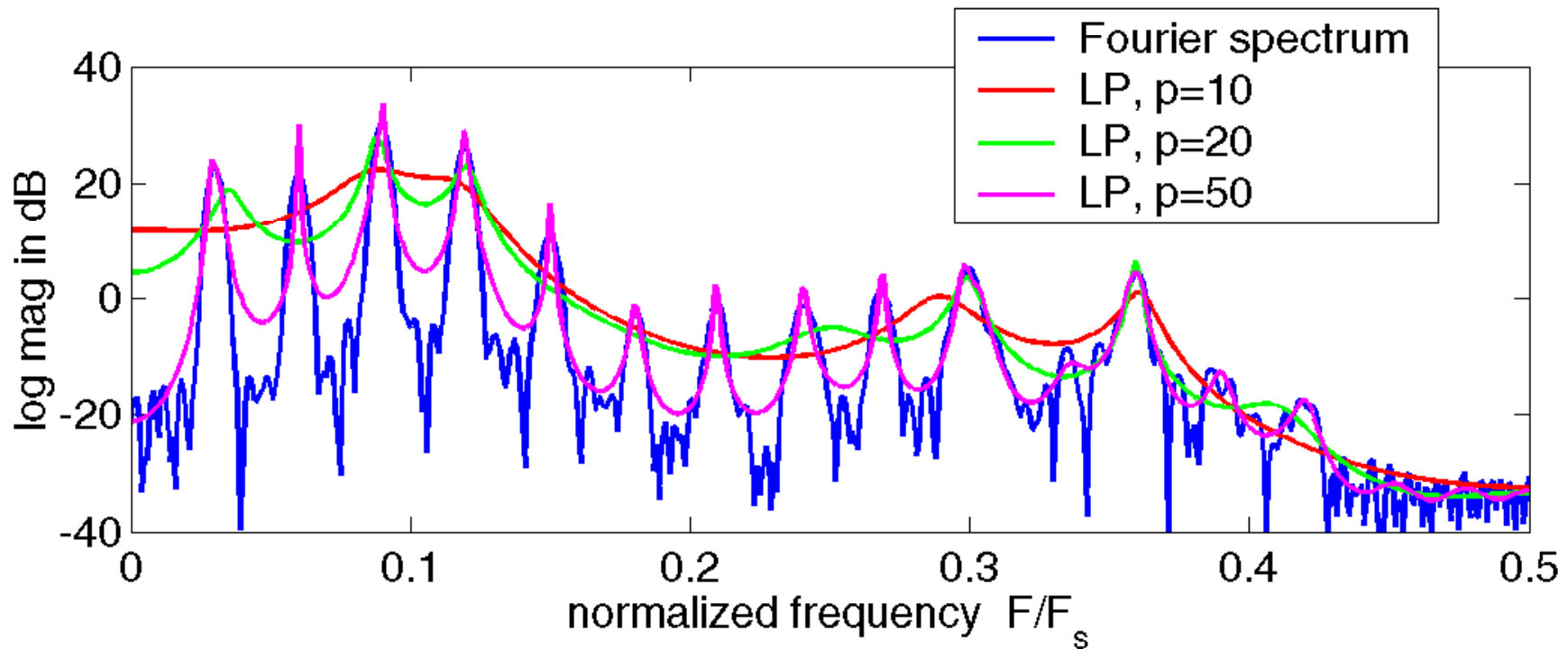
- with frequency response

$$\tilde{H}(e^{j\omega}) = \frac{G}{1 - \sum_{k=1}^p \alpha_k e^{-j\omega k}} = \frac{G}{A(e^{j\omega})}$$

- with the gain determined by matching the energy of the model to the short-time energy of the speech signal, i.e.,

$$G^2 = E_{\hat{n}} = \sum_m (e_{\hat{n}}(m))^2 = R_{\hat{n}}(0) - \sum_{k=1}^p \alpha_k R_{\hat{n}}(k)$$

LPC Spectrum



$$\tilde{H}(e^{j\omega}) = \frac{G}{1 - \sum_{k=1}^p \alpha_k e^{-j\omega k}}$$

```
x = s .* hamming(301);  
X = fft( x , 1000 )  
[A , G , r ] = autolpc( x , 10 )  
H = G ./ fft(A,1000);
```

LP Analysis is seen to be a method of short-time spectrum estimation with removal of excitation fine structure (a form of wideband spectrum analysis)

LP Short-Time Spectrum Analysis

- Defined speech segment as:

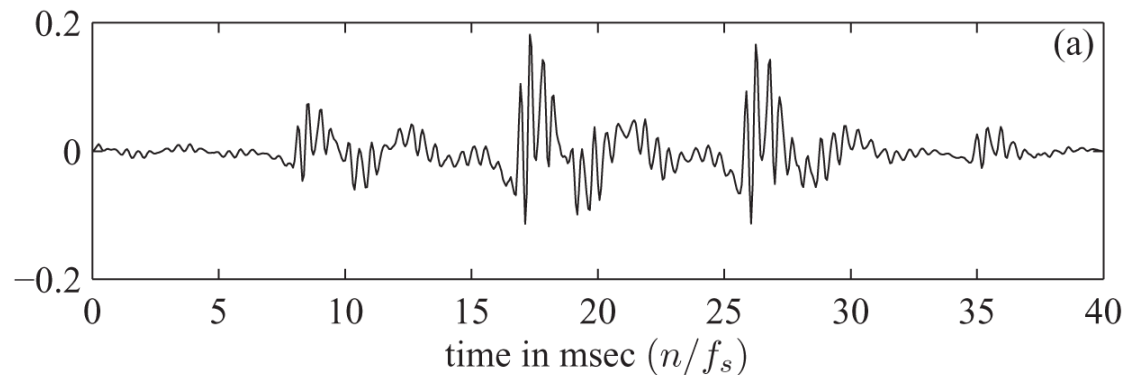
$$s_{\hat{n}}[m] = s[m + \hat{n}]w[m]$$

- The discrete-time Fourier transform of this windowed segment is:

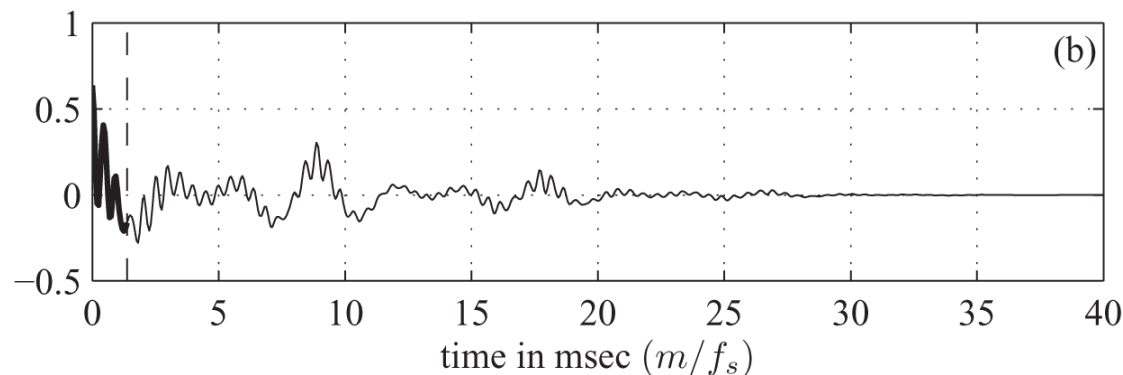
$$S_{\hat{n}}(e^{j\omega}) = \sum_{m=-\infty}^{\infty} s[m + \hat{n}]w[m]e^{-j\omega m}$$

- Short-time FT and the LP spectrum are linked via short-time autocorrelation

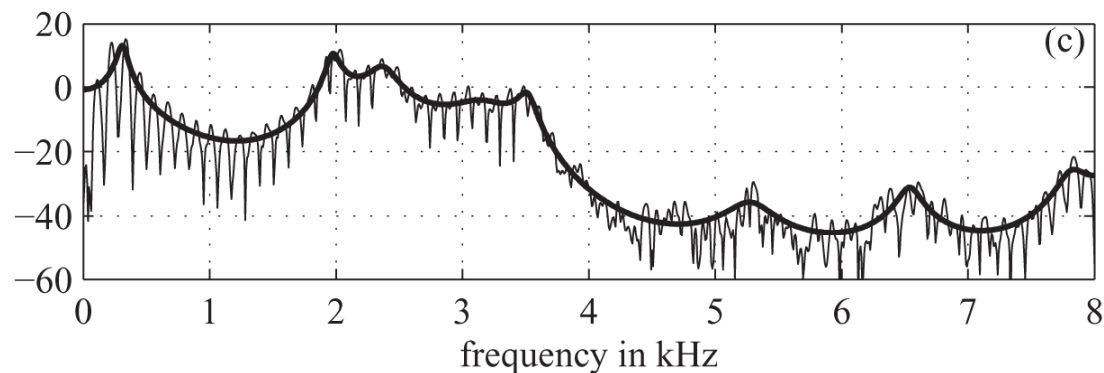
LP Short-Time Spectrum Analysis



(a) Voiced speech segment obtained using a Hamming window

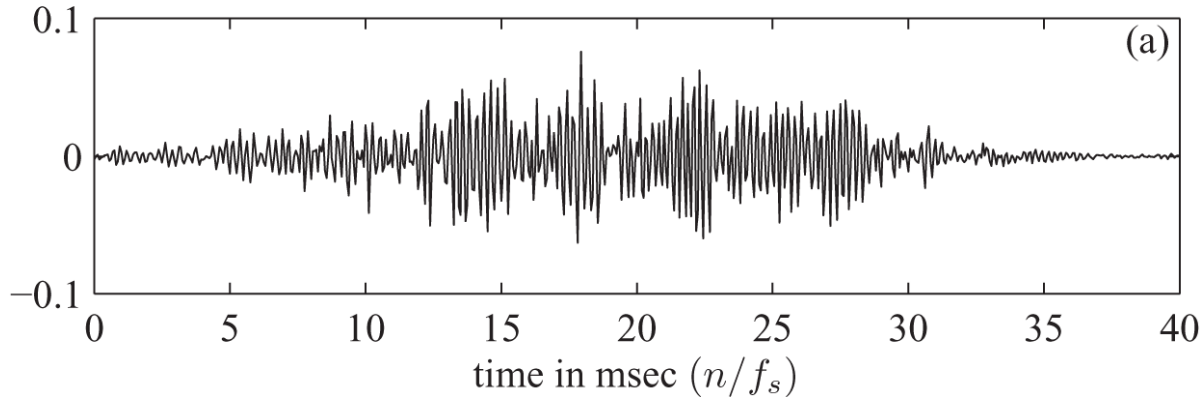


(b) Corresponding short-time autocorrelation function used in LP analysis (heavy line shows values used in LP analysis)

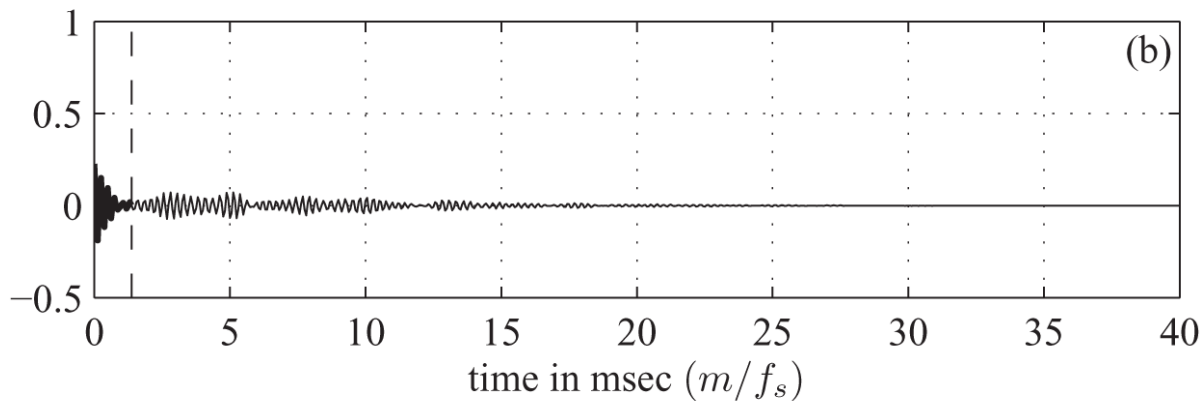


(c) Corresponding short-time log magnitude Fourier transform and short-time log magnitude LPC spectrum ($F_s=16$ kHz)

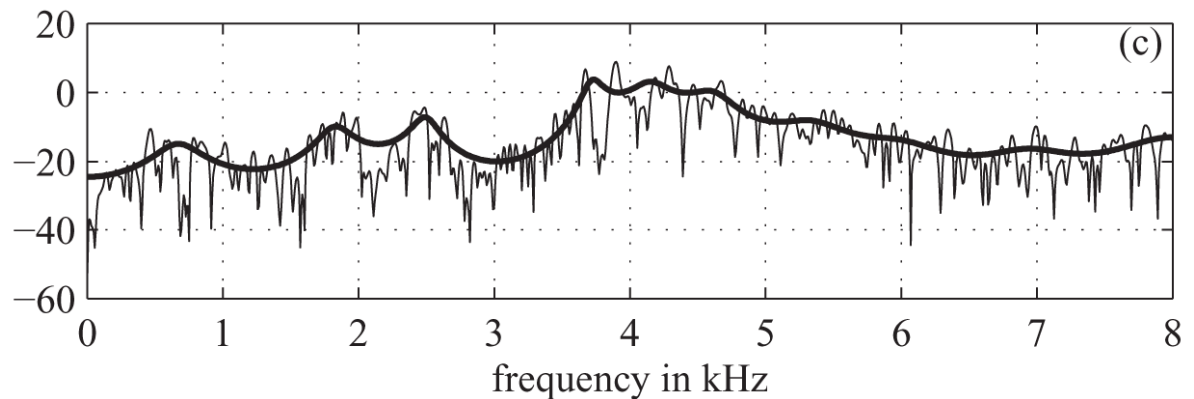
LP Short-Time Spectrum Analysis



(a) Unvoiced speech segment obtained using a Hamming window



(b) Corresponding short-time autocorrelation function used in LP analysis (heavy line shows values used in LP analysis)



(c) Corresponding short-time log magnitude Fourier transform and short-time log magnitude LPC spectrum ($F_s=16$ kHz)

Frequency Domain Interpretation of Mean-Squared Prediction Error

- The LP spectrum provides a basis for examining the properties of the prediction error (or equivalently the excitation of the VT)
- The mean-squared prediction error at sample \hat{n} is:

$$E_{\hat{n}} = \sum_{m=0}^{L+p-1} e_{\hat{n}}^2[m]$$

- which, by Parseval's Theorem, can be expressed as:

$$E_{\hat{n}} = \frac{1}{2\pi} \int_{-\pi}^{\pi} |\mathbf{E}_{\hat{n}}(e^{j\omega})|^2 d\omega = \frac{1}{2\pi} \int_{-\pi}^{\pi} |S_{\hat{n}}(e^{j\omega})|^2 |A(e^{j\omega})|^2 d\omega = G^2$$

- where $S_{\hat{n}}(e^{j\omega})$ is the FT of $s_{\hat{n}}[m]$ and $A(e^{j\omega})$ is the corresponding prediction error frequency response

$$A(e^{j\omega}) = 1 - \sum_{k=1}^p \alpha_k e^{-j\omega k}$$

Frequency Domain Interpretation of Mean-Squared Prediction Error

- The LP spectrum is of the form:

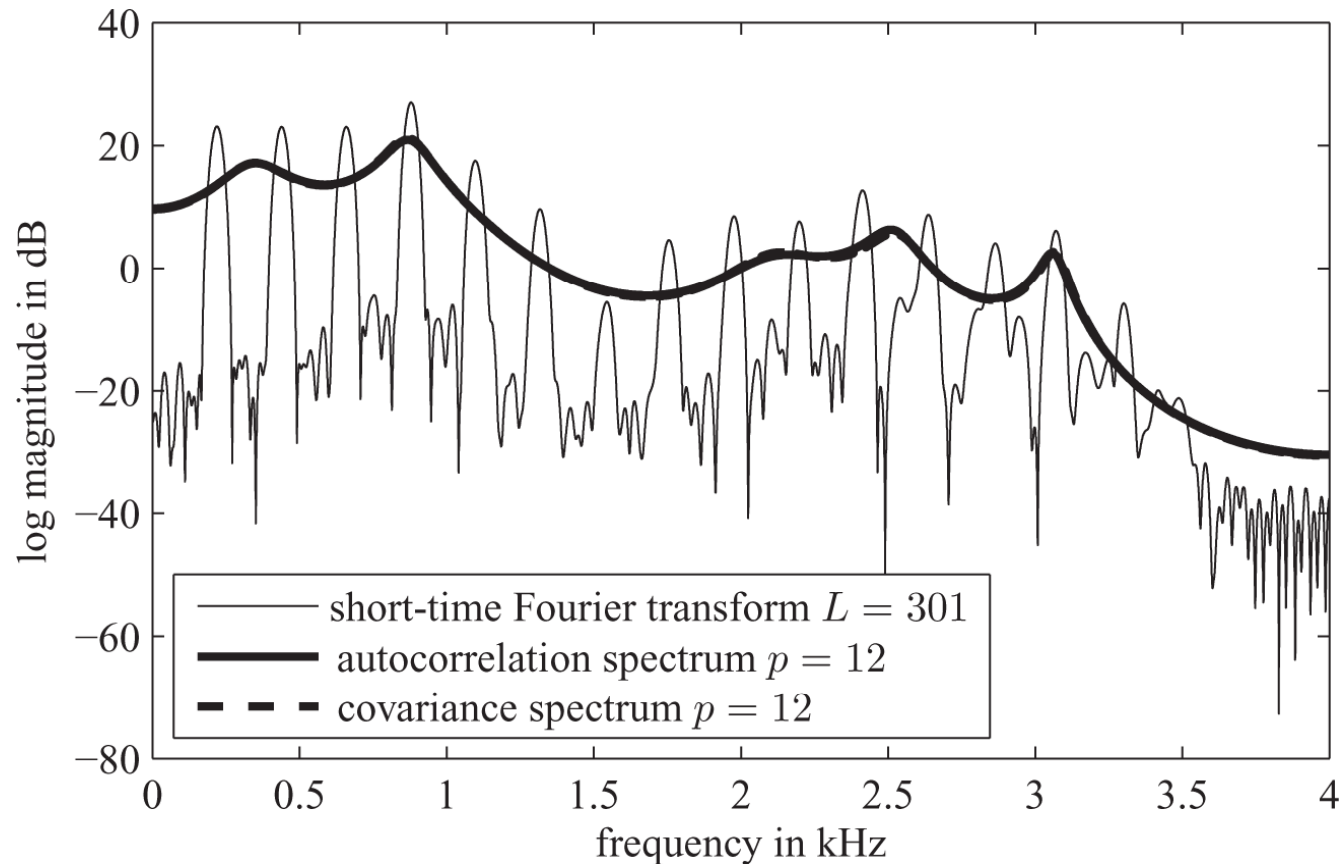
$$\tilde{H}(e^{j\omega}) = \frac{G}{A(e^{j\omega})}$$

- Thus we can express the mean-squared error as:

$$E_{\hat{n}} = \frac{G^2}{2\pi} \int_{-\pi}^{\pi} \frac{|S_{\hat{n}}(e^{j\omega})|^2}{|\tilde{H}(e^{j\omega})|^2} d\omega = G^2$$

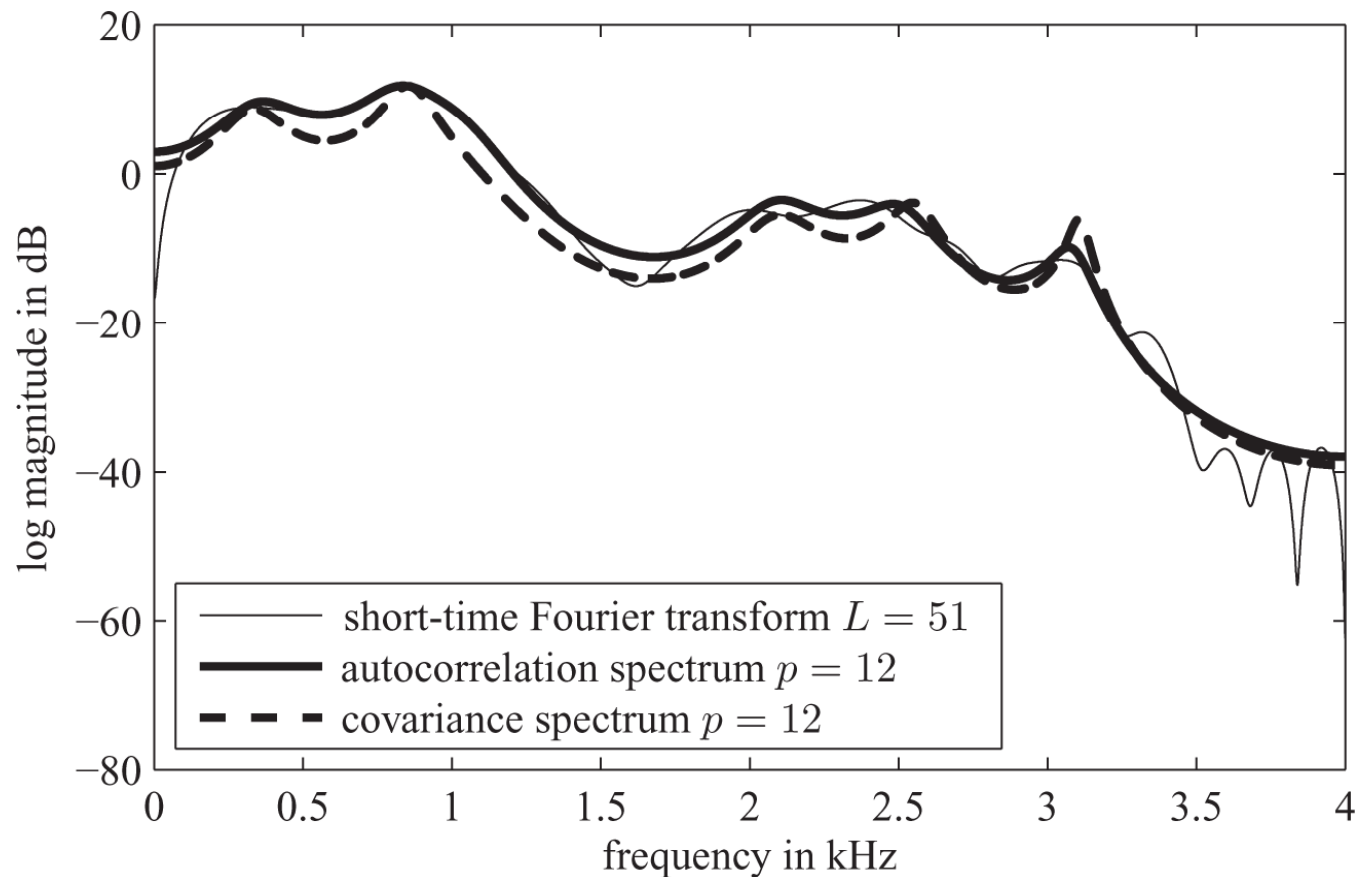
- We see that minimizing total squared prediction error is equivalent to finding gain and predictor coefficients such that the integral of the ratio of the energy spectrum of the speech segment to the magnitude squared of the frequency response of the model linear system is unity.
- Thus $|S_{\hat{n}}(e^{j\omega})|^2$ can be interpreted as a frequency-domain weighting function \Rightarrow LP weights frequencies where $|S_{\hat{n}}(e^{j\omega})|^2$ is large more heavily than when $|S_{\hat{n}}(e^{j\omega})|^2$ is small.

LP Interpretation Example 1



Much better spectral matches to STFT spectral peaks than to STFT spectral valleys as predicted by spectral interpretation of error minimization.

LP Interpretation Example2



Note small differences in spectral shape between STFT, autocorrelation spectrum and covariance spectrum when using short window duration ($L=51$ samples).

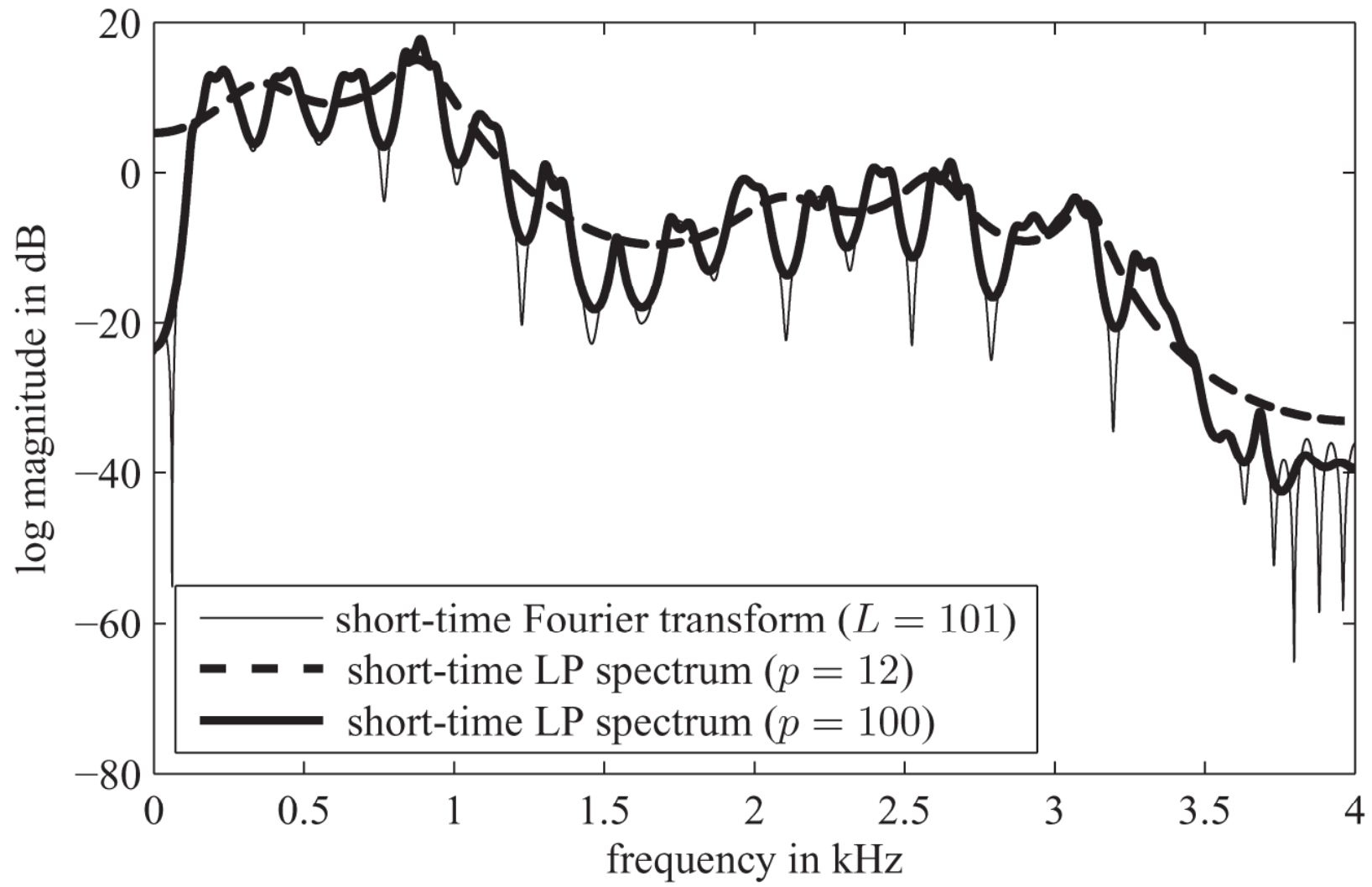
Effects of Model Order

- The AC function, $R_{\hat{n}}[m]$ of the speech segment, $s_{\hat{n}}[m]$, and the AC function, $\tilde{R}[m]$, of the impulse response, $\tilde{h}[m]$, corresponding to the system function, $\tilde{H}(z)$, are equal for the first $(p+1)$ values. Thus, as $p \rightarrow \infty$, the AC functions are equal for all values and thus:

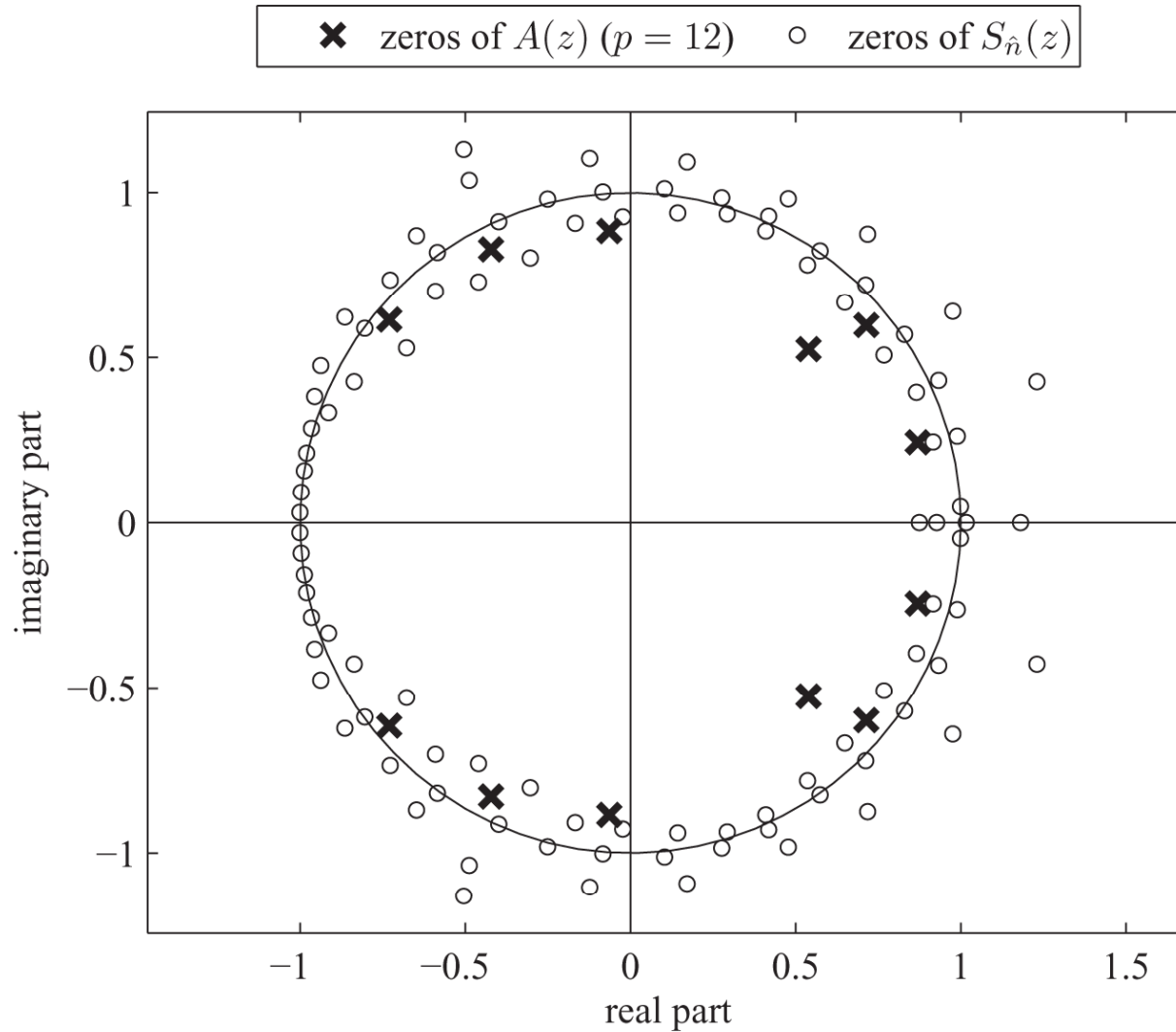
$$\lim_{p \rightarrow \infty} |\tilde{H}(e^{j\omega})|^2 = |S_{\hat{n}}(e^{j\omega})|^2$$

- Thus if p is large enough, the FR of the all-pole model, $\tilde{H}(e^{j\omega})$, can approximate the signal spectrum with arbitrarily small error.

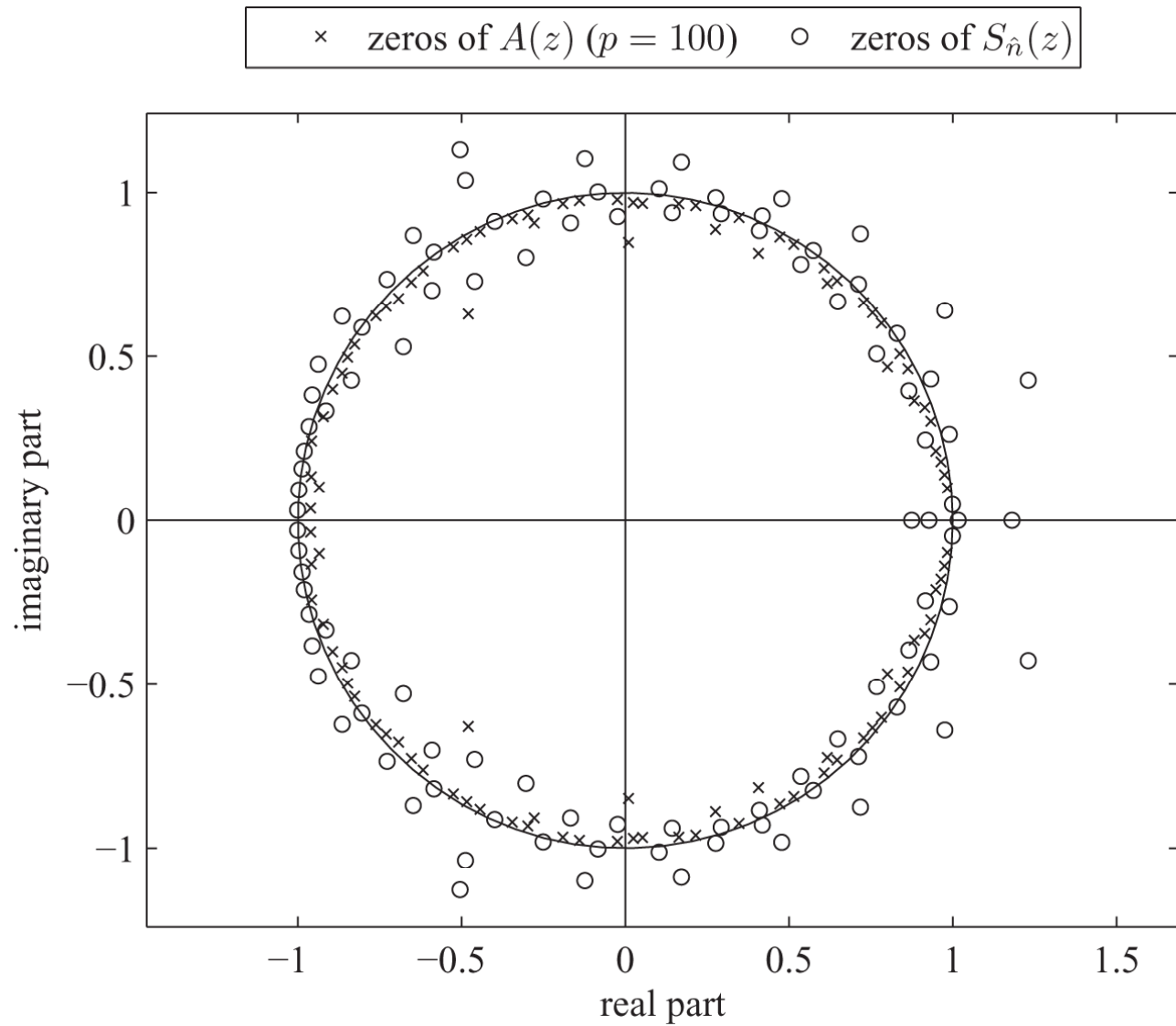
Effects of Model Order



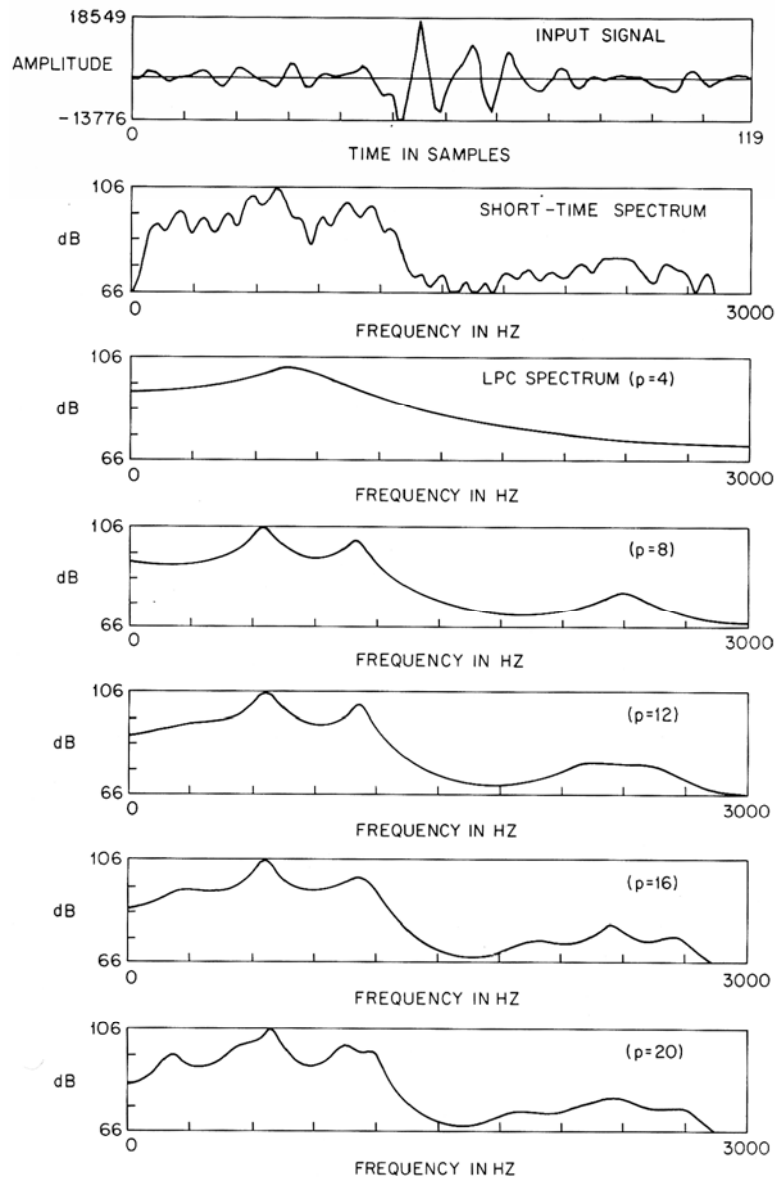
Effects of Model Order



Effects of Model Order



Effects of Model Order



- plots show Fourier transform of segment and LP spectra for various orders
 - as p increases, more details of the spectrum are preserved
 - need to choose a value of p that represents the spectral effects of the glottal pulse, vocal tract and radiation--nothing else

Linear Prediction Spectrogram

- Speech spectrogram previously defined as:

$$20 \log |S_r[k]| = 20 \log \left| \sum_{m=0}^{L-1} s[rR + m] w[m] e^{-j(2\pi/N)km} \right|$$

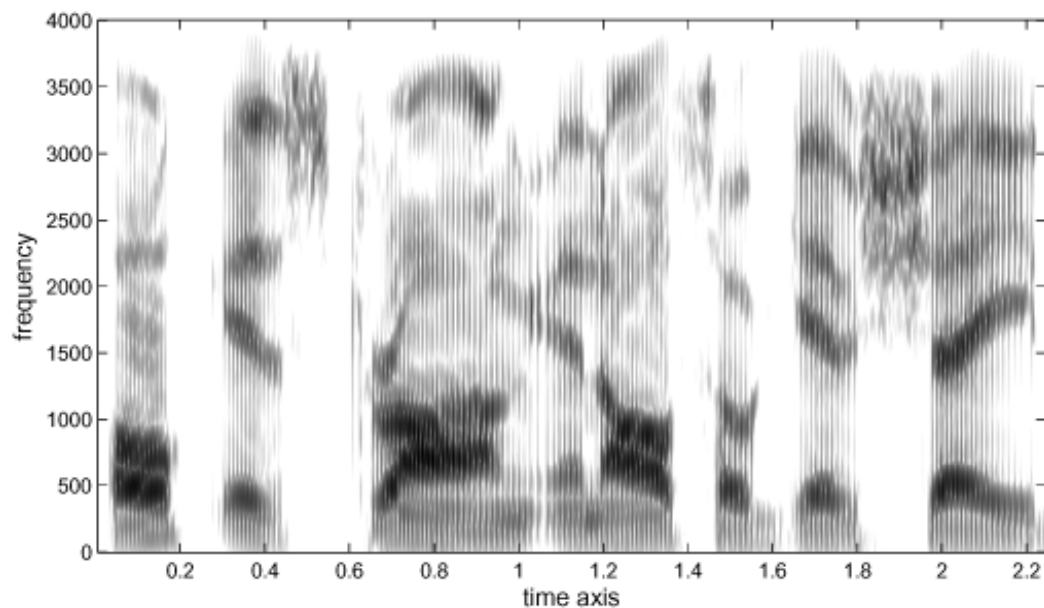
for set of times, $t_r = rRT$, and set of frequencies, $F_k = kF_s / N$, $k = 1, 2, \dots, N/2$ where R is the time shift (in samples) between adjacent STFTs, T is the sampling period, $F_s = 1/T$ is the sampling frequency, and N is the size of the discrete Fourier transform used to compute each STFT estimate.

- Similarly we can define the LP spectrogram as an image plot of:

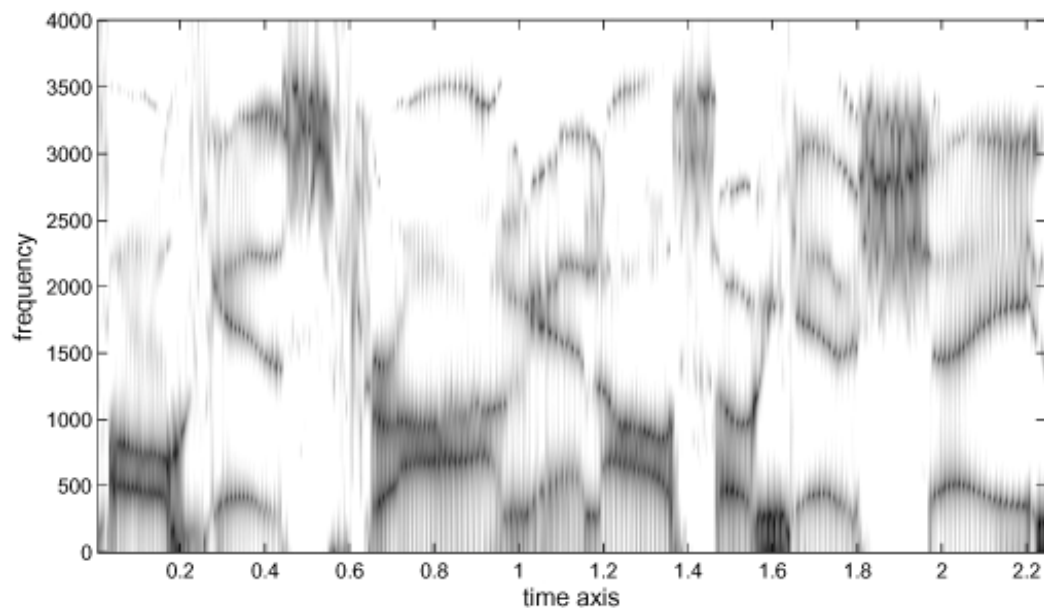
$$20 \log |\tilde{H}_r[k]| = 20 \log \left| \frac{G_r}{A_r(e^{j(2\pi/N)k})} \right|$$

where G_r and $A_r(e^{j(2\pi/N)k})$ are the gain and prediction error polynomial at analysis time rR .

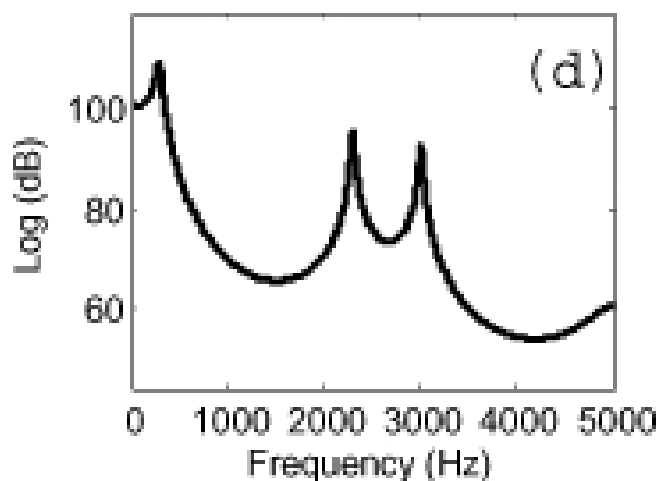
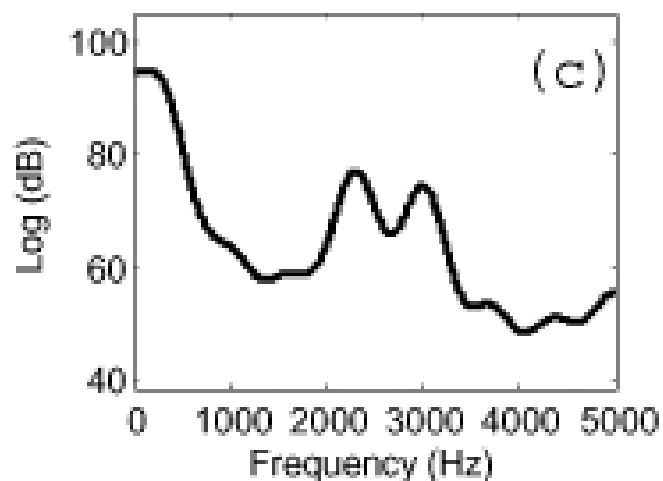
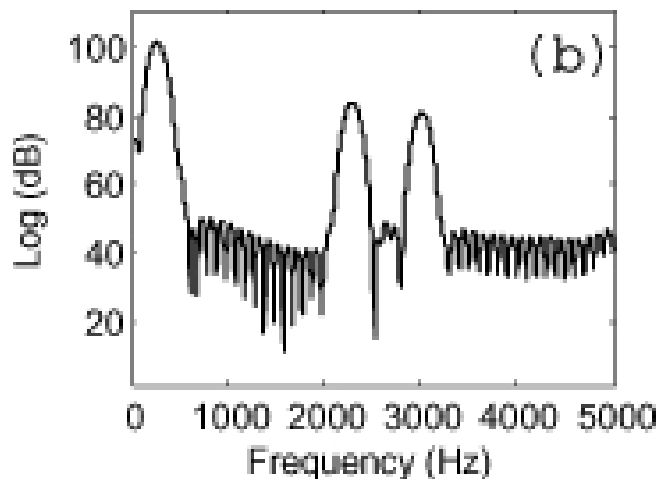
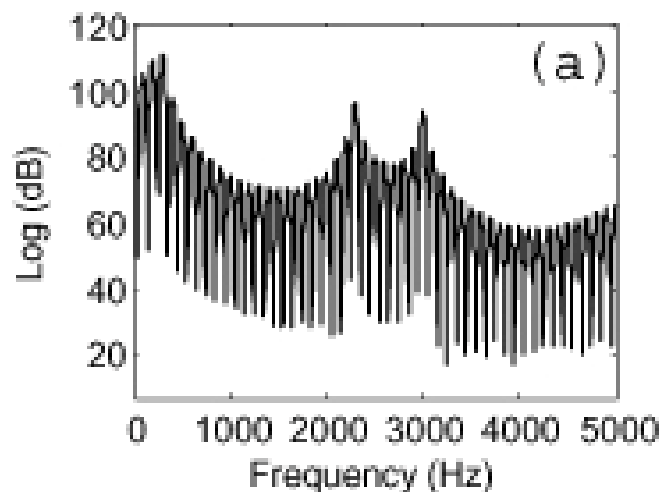
Linear Prediction Spectrogram



$L=81$, $R=3$, $N=1000$,
40 db dynamic range



Comparison to Other Spectrum Analysis Methods



Spectra of synthetic vowel /IY/

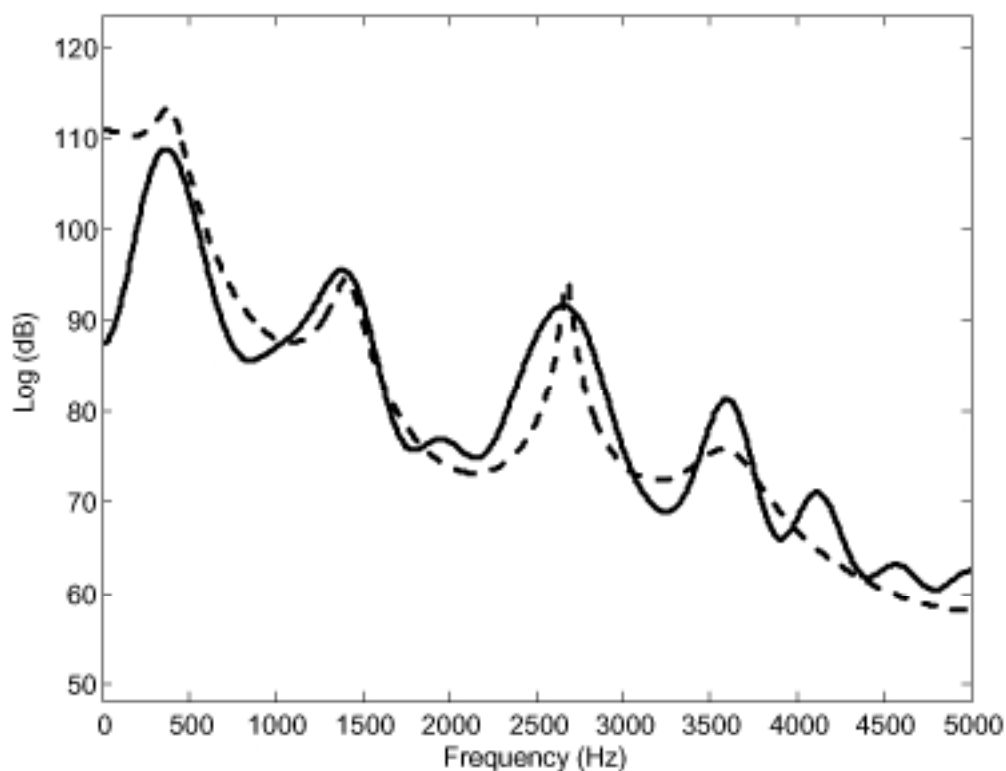
(a) Narrowband spectrum using 40 msec window

(b) Wideband spectrum using a 10 msec window

(c) Cepstrally smoothed spectrum

(d) LPC spectrum from a 40 msec section using a $p=12$ order LPC analysis

Comparison to Other Spectrum Analysis Methods



Natural speech spectral estimates using cepstral smoothing (solid line) and linear prediction analysis (dashed line).

Note the fewer (spurious) peaks in the LP analysis spectrum since LP used $p=12$ which restricted the spectral match to a maximum of 6 resonance peaks.

Note the narrow bandwidths of the LP resonances versus the cepstrally smoothed resonances.

Selective Linear Prediction

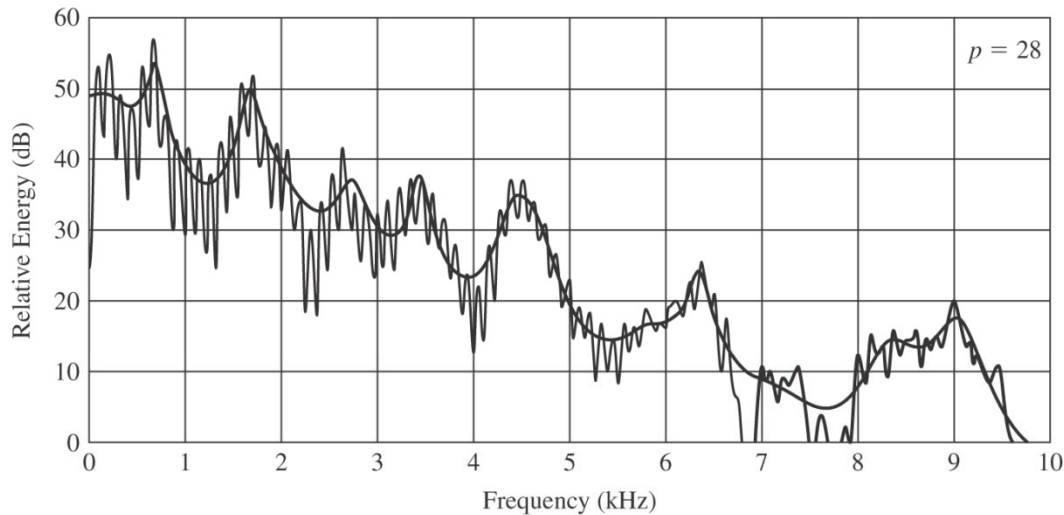
- it is possible to apply LP methods to selected parts of spectrum
 - 0-4 kHz for voiced sounds \Rightarrow use a predictor of order p_1
 - 4-8 kHz for unvoiced sounds \Rightarrow use a predictor of order p_2
- the key idea is to map the frequency region $\{f_A, f_B\}$ linearly to $\{0, .5\}$ or, equivalently, the region $\{2\pi f_A, 2\pi f_B\}$ maps linearly to $\{0, \pi\}$ via the transformation

$$\omega' = \frac{\omega - 2\pi f_A}{2\pi f_B - 2\pi f_A} \cdot 2\pi f_B$$

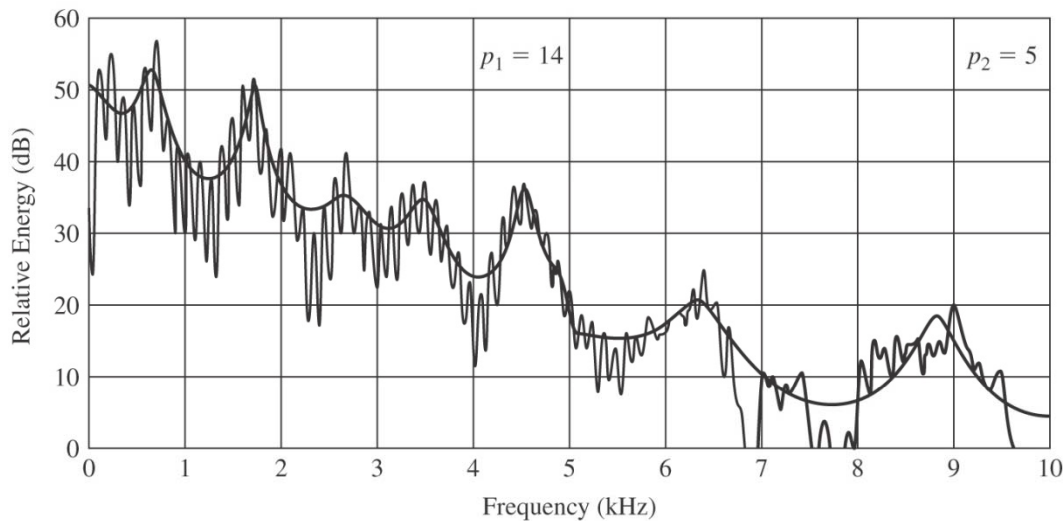
- we must modify the calculation for the autocorrelation to give:

$$R'(m) = \frac{1}{2\pi} \int_{-\pi}^{\pi} |S_{\hat{n}}(e^{j\omega'})|^2 e^{j\omega' m} d\omega'$$

Selective Linear Prediction



(a)



(b)

- 0-10 kHz region modeled using $p=28$
- no discontinuity in model spectra at 5 kHz

- 0-5 kHz region modeled using $p_1=14$
- 5-10 kHz region modeled using $p_2=5$
- discontinuity in model spectra at 5 kHz

Solutions of LPC Equations

**Covariance Method (Cholesky
Decomposition Method)**

LPC Solutions-Covariance Method

- for the covariance method we need to solve the matrix equation

$$\sum_{k=1}^p \alpha_k \phi_{\hat{n}}(i, k) = \phi_{\hat{n}}(i, 0), \quad i = 1, 2, \dots, p$$

$\phi\alpha = \psi$ (in matrix notation)

- ϕ is a positive definite, symmetric matrix with (i, j) element $\phi_{\hat{n}}(i, j)$, and α and ψ are column vectors with elements α_i and $\phi_{\hat{n}}(i, 0)$
- the solution of the matrix equation is called the Cholesky decomposition, or square root method

$\phi = \mathbf{V}\mathbf{D}\mathbf{V}^t$; \mathbf{V} = lower triangular matrix with 1's on the main diagonal
 \mathbf{D} =diagonal matrix

LPC Solutions-Covariance Method

- can readily determine elements of V and D by solving for (i, j) elements of the matrix equation, as follows

$$\phi_{\hat{n}}(i, j) = \sum_{k=1}^j V_{ik} d_k V_{jk}, \quad 1 \leq j \leq i-1$$

- giving

$$V_{ij} d_j = \phi_{\hat{n}}(i, j) - \sum_{k=1}^{j-1} V_{ik} d_k V_{jk}, \quad 1 \leq j \leq i-1$$

- and for the diagonal elements

$$\phi_{\hat{n}}(i, i) = \sum_{k=1}^i V_{ik} d_k V_{ik}$$

- giving

$$d_i = \phi_{\hat{n}}(i, i) - \sum_{k=1}^{i-1} V_{ik}^2 d_k, \quad i \geq 2$$

- with

$$d_1 = \phi_{\hat{n}}(1, 1)$$

Cholesky Decomposition Example

- consider example with $p = 4$, and matrix elements $\phi_{\hat{n}}(i, j) = \phi_{ij}$

$$\begin{bmatrix} \phi_{11} & \phi_{21} & \phi_{31} & \phi_{41} \\ \phi_{21} & \phi_{22} & \phi_{32} & \phi_{42} \\ \phi_{31} & \phi_{32} & \phi_{33} & \phi_{43} \\ \phi_{41} & \phi_{42} & \phi_{34} & \phi_{44} \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 0 \\ V_{21} & 1 & 0 & 0 \\ V_{31} & V_{32} & 1 & 0 \\ V_{41} & V_{42} & V_{43} & 1 \end{bmatrix} \begin{bmatrix} d_1 & 0 & 0 & 0 \\ 0 & d_2 & 0 & 0 \\ 0 & 0 & d_3 & 0 \\ 0 & 0 & 0 & d_4 \end{bmatrix} \begin{bmatrix} 1 & V_{21} & V_{31} & V_{41} \\ 0 & 1 & V_{32} & V_{42} \\ 0 & 0 & 1 & V_{43} \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

Cholesky Decomposition Example

- solve matrix for $d_1, V_{21}, V_{31}, V_{41}, d_2, V_{32}, V_{42}, d_3, V_{43}, d_4$

$$d_1 = \phi_{11} \quad \text{step 1}$$

$$V_{21}d_1 = \phi_{21}; \quad V_{31}d_1 = \phi_{31}; \quad V_{41}d_1 = \phi_{41}$$

$$V_{21} = \phi_{21} / d_1; \quad V_{31} = \phi_{31} / d_1; \quad V_{41} = \phi_{41} / d_1 \quad \text{step 2}$$

$$d_2 = \phi_{22} - V_{21}^2 d_1 \quad \text{step 3}$$

$$V_{32}d_2 = \phi_{32} - V_{31}d_1V_{21} \Rightarrow V_{32} = (\phi_{32} - V_{31}d_1V_{21}) / d_2$$

$$V_{42}d_2 = \phi_{42} - V_{41}d_1V_{21} \Rightarrow V_{42} = (\phi_{42} - V_{41}d_1V_{21}) / d_2 \quad \text{step 4}$$

- iterate procedure to solve for d_3, V_{43}, d_4

LPC Solutions-Covariance Method

- now need to solve for α using a 2-step procedure

$$VDV^t\alpha = \psi$$

- writing this as

$$VY = \psi \quad \text{with}$$

$$DV^t\alpha = Y \quad \text{or}$$

$$V^t\alpha = D^{-1}Y$$

- from V (which is now known) solve for column vector Y using a simple recursion of the form

$$Y_i = \psi_i - \sum_{j=1}^{i-1} V_{ij} Y_j, \quad p \geq i \geq 2$$

- with initial condition

$$Y_1 = \psi_1$$

LPC Solutions-Covariance Method

- now can solve for α using the recursion

$$\alpha_i = Y_i / d_i - \sum_{j=i+1}^p V_{ji} \alpha_j, \quad 1 \leq i \leq p-1$$

- with initial condition

$$\alpha_p = Y_p / d_p$$

- calculation proceeds backwards from $i = p-1$
down to $i = 1$

Cholesky Decomposition Example

- continuing the example we solve for Y

$$\begin{bmatrix} 1 & 0 & 0 & 0 \\ V_{21} & 1 & 0 & 0 \\ V_{31} & V_{32} & 1 & 0 \\ V_{41} & V_{42} & V_{43} & 1 \end{bmatrix} \begin{bmatrix} Y_1 \\ Y_2 \\ Y_3 \\ Y_4 \end{bmatrix} = \begin{bmatrix} \psi_1 \\ \psi_2 \\ \psi_3 \\ \psi_4 \end{bmatrix}$$

- first solving for $Y_1 - Y_4$ we get

$$Y_1 = \psi_1$$

$$Y_2 = \psi_2 - V_{21}Y_1$$

$$Y_3 = \psi_3 - V_{31}Y_1 - V_{32}Y_2$$

$$Y_4 = \psi_4 - V_{41}Y_1 - V_{42}Y_2 - V_{43}Y_3$$

Cholesky Decomposition Example

- next solve for α from equation

$$\begin{bmatrix} 1 & V_{21} & V_{31} & V_{41} \\ 0 & 1 & V_{32} & V_{42} \\ 0 & 0 & 1 & V_{43} \\ 0 & 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} \alpha_1 \\ \alpha_2 \\ \alpha_3 \\ \alpha_4 \end{bmatrix} = \begin{bmatrix} 1/d_1 & 0 & 0 & 0 \\ 0 & 1/d_2 & 0 & 0 \\ 0 & 0 & 1/d_3 & 0 \\ 0 & 0 & 0 & 1/d_4 \end{bmatrix} \begin{bmatrix} Y_1 \\ Y_2 \\ Y_3 \\ Y_4 \end{bmatrix} = \begin{bmatrix} Y_1/d_1 \\ Y_2/d_2 \\ Y_3/d_3 \\ Y_4/d_4 \end{bmatrix}$$

- giving the results

$$\alpha_4 = Y_4 / d_4$$

$$\alpha_3 = Y_3 / d_3 - V_{43} \alpha_4$$

$$\alpha_2 = Y_2 / d_2 - V_{32} \alpha_3 - V_{42} \alpha_4$$

$$\alpha_1 = Y_1 / d_1 - V_{21} \alpha_2 - V_{31} \alpha_3 - V_{41} \alpha_4$$

- completing the solution

Covariance Method Minimum Error

- the minimum mean squared error can be written in the form

$$\begin{aligned} E_{\hat{n}} &= \phi_{\hat{n}}(0,0) - \sum_{k=1}^p \alpha_k \phi_{\hat{n}}(0,k) \\ &= \phi_{\hat{n}}(0,0) - \alpha^t \psi \end{aligned}$$

- since $\alpha^t = Y^t D^{-1} V^{-1}$ can write this as

$$\begin{aligned} E_{\hat{n}} &= \phi_{\hat{n}}(0,0) - Y^t D^{-1} Y \\ &= \phi_{\hat{n}}(0,0) - \sum_{k=1}^p Y_k^2 / d_k \end{aligned}$$

- this computation for $E_{\hat{n}}$ can be used for all values of LP order from 1 to $p \Rightarrow$ can understand how LP order reduces mean-squared error

Cholesky Matrix Inversion Algorithm

```
%% Find first column of  $V$ 
 $d_1 = \phi_{11}$  (9.66)
```

```
for  $i = 2, 3, \dots, p$ 
     $V_{i1} = \phi_{i1}/d_1$  (9.67)
```

```
end
%% Find  $D$  and remaining columns of  $V$ 
for  $j = 2, 3, \dots, p - 1$ 
     $d_j = \phi_{jj} - \sum_{k=1}^{j-1} V_{jk}^2 d_k$  (9.68)
```

```
    for  $i = j + 1, \dots, p$ 
         $V_{ij} = (\phi_{ij} - \sum_{k=1}^{j-1} V_{ik} d_k V_{jk})/d_j$  (9.69)
```

```
    end
end
 $d_p = \phi_{pp} - \sum_{k=1}^{p-1} V_{pk}^2 d_k$  (9.68)
```

```
%% Find  $Y = DV^T \alpha$ 
 $Y_1 = \psi_1$  (9.72)
```

```
for  $i = 2, 3, \dots, p$ 
     $Y_i = \psi_i - \sum_{j=1}^{i-1} V_{ij} Y_j$  (9.73)
```

```
end
%% Find  $\alpha$  from  $Y$ 
 $\alpha_p = Y_p/d_p$  (9.76)
```

```
for  $i = p - 1, p - 2, \dots, 1$ 
     $\alpha_i = Y_i/d_i - \sum_{j=i+1}^p V_{ij} \alpha_j$  (9.77)
```

```
end
```

Solutions of LPC Equations

**Autocorrelation Method via
Levinson-Durbin Algorithm**

Levinson-Durbin Algorithm 1

- Autocorrelation equations (at each frame \hat{n}):

$$\sum_{k=1}^p \alpha_k R[|i-k|] = R[i] \quad 1 \leq i \leq p$$

$$\mathbf{R}\alpha = \mathbf{r}$$

- \mathbf{R} is a positive definite symmetric Toeplitz matrix
- The set of optimum predictor coefficients satisfy:

$$R[i] - \sum_{k=1}^p \alpha_k R[|i-k|] = 0, \quad 1 \leq i \leq p$$

- with minimum mean-squared prediction error of:

$$R[0] - \sum_{k=1}^p \alpha_k R[k] = E^{(p)}$$

Levinson-Durbin Algorithm 2

- By combining the last two equations we get a larger matrix equation of the form:

$$\begin{bmatrix} R[0] & R[1] & R[2] & \dots & R[p] \\ R[1] & R[0] & R[1] & \dots & R[p-1] \\ R[2] & R[1] & R[0] & \dots & R[p-2] \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ R[p] & R[p-1] & R[p-2] & \dots & R[0] \end{bmatrix} \begin{bmatrix} 1 \\ -\alpha_1^{(p)} \\ -\alpha_2^{(p)} \\ \cdot \\ -\alpha_p^{(p)} \end{bmatrix} = \begin{bmatrix} E^{(p)} \\ 0 \\ 0 \\ \cdot \\ 0 \end{bmatrix}$$

- expanded matrix is still Toeplitz and can be solved iteratively by incorporating new correlation value at each iteration and solving for next higher order predictor in terms of new correlation value and previous predictor

Levinson-Durbin Algorithm 3

- Show how i^{th} order solution can be derived from $(i-1)^{st}$ order solution; i.e., given $\alpha^{(i-1)}$, the solution to $R^{(i-1)}\alpha^{(i-1)} = e^{(i-1)}$ we derive solution to $R^{(i)}\alpha^{(i)} = e^{(i)}$
- The $(i-1)^{st}$ solution can be expressed as:

$$\begin{bmatrix} R[0] & R[1] & R[2] & \dots & R[i-1] \\ R[1] & R[0] & R[1] & \dots & R[i-2] \\ R[2] & R[1] & R[0] & \dots & R[i-3] \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ R[i-1] & R[i-2] & R[i-3] & \dots & R[0] \end{bmatrix} \begin{bmatrix} 1 \\ -\alpha_1^{(i-1)} \\ -\alpha_2^{(i-1)} \\ \cdot \\ -\alpha_{i-1}^{(i-1)} \end{bmatrix} = \begin{bmatrix} E^{(i-1)} \\ 0 \\ 0 \\ \cdot \\ 0 \end{bmatrix}$$

Levinson-Durbin Algorithm 4

- Appending a 0 to vector $\alpha^{(i-1)}$ and multiplying by the matrix $R^{(i)}$ gives:

$$\begin{bmatrix} R[0] & R[1] & R[2] & \dots & R[i] \\ R[1] & R[0] & R[1] & \dots & R[i-1] \\ R[2] & R[1] & R[0] & \dots & R[i-2] \\ \cdot & \cdot & \cdot & \dots & \cdot \\ R[i-1] & R[i-2] & R[i-3] & \dots & R[1] \\ R[i] & R[i-1] & R[i-2] & \dots & R[0] \end{bmatrix} \begin{bmatrix} 1 \\ -\alpha_1^{(i-1)} \\ -\alpha_2^{(i-1)} \\ \cdot \\ -\alpha_{i-1}^{(i-1)} \\ 0 \end{bmatrix} = \begin{bmatrix} E^{(i-1)} \\ 0 \\ 0 \\ \cdot \\ 0 \\ \gamma^{(i-1)} \end{bmatrix}$$

- where $\gamma^{(i-1)} = R[i] - \sum_{j=1}^{i-1} \alpha_j^{(i-1)} R[i-j]$ and $R[i]$ are introduced

Levinson-Durbin Algorithm 5

- Key step is that since Toeplitz matrix has special symmetry we can reverse the order of the equations (first equation last, last equation first), giving:

$$\begin{bmatrix} R[0] & R[1] & R[2] & \dots & R[i] \\ R[1] & R[0] & R[1] & \dots & R[i-1] \\ R[2] & R[1] & R[0] & \dots & R[i-2] \\ \cdot & \cdot & \cdot & \dots & \cdot \\ R[i-1] & R[i-2] & R[i-3] & \dots & R[1] \\ R[i] & R[i-1] & R[i-2] & \dots & R[0] \end{bmatrix} \begin{bmatrix} 0 \\ -\alpha_1^{(i-1)} \\ -\alpha_2^{(i-1)} \\ \cdot \\ -\alpha_{i-1}^{(i-1)} \\ 1 \end{bmatrix} = \begin{bmatrix} \gamma^{(i-1)} \\ 0 \\ 0 \\ \cdot \\ 0 \\ E^{(i-1)} \end{bmatrix}$$

Levinson-Durbin Algorithm 6

- To get the equation into the desired form (a single component in the vector $e^{(i)}$) we combine the two sets of matrices (with a multiplicative factor k_i) giving:

$$R^{(i)} \left[\begin{bmatrix} 1 \\ -\alpha_1^{(i-1)} \\ -\alpha_2^{(i-1)} \\ \cdot \\ -\alpha_{i-1}^{(i-1)} \\ 0 \end{bmatrix} - k_i \begin{bmatrix} 0 \\ -\alpha_1^{(i-1)} \\ -\alpha_2^{(i-1)} \\ \cdot \\ -\alpha_{i-1}^{(i-1)} \\ 1 \end{bmatrix} \right] = \left[\begin{bmatrix} E^{(i-1)} \\ 0 \\ 0 \\ \cdot \\ 0 \\ \gamma^{(i-1)} \end{bmatrix} - k_i \begin{bmatrix} \gamma^{(i-1)} \\ 0 \\ 0 \\ \cdot \\ 0 \\ E^{(i-1)} \end{bmatrix} \right]$$

- Choose $\gamma^{(i-1)}$ so that vector on right has only a single non-zero entry, i.e.,

$$k_i = \frac{\gamma^{(i-1)}}{E^{(i-1)}} = \frac{R[i] - \sum_{j=1}^{i-1} \alpha_j^{(i-1)} R[i-j]}{E^{(i-1)}}$$

Levinson-Durbin Algorithm 7

- The first element of the right hand side vector is now:

$$E^{(i)} = E^{(i-1)} - k_i \gamma^{(i-1)} = E^{(i-1)} (1 - k_i^2)$$

- The k_i parameters are called PARCOR coefficients.
- With this choice of $\gamma^{(i-1)}$, the vector of i^{th} order predictor coefficients is:

$$\begin{bmatrix} 1 \\ -\alpha_1^{(i)} \\ -\alpha_2^{(i)} \\ \cdot \\ -\alpha_{i-1}^{(i)} \\ -\alpha_i^{(i)} \end{bmatrix} = \begin{bmatrix} 1 \\ -\alpha_1^{(i-1)} \\ -\alpha_2^{(i-1)} \\ \cdot \\ -\alpha_{i-1}^{(i-1)} \\ 0 \end{bmatrix} - k_i \begin{bmatrix} 0 \\ -\alpha_{i-1}^{(i-1)} \\ -\alpha_{i-2}^{(i-1)} \\ \cdot \\ -\alpha_1^{(i-1)} \\ 1 \end{bmatrix}$$

- yielding the updating procedure

$$\alpha_j^{(i)} = \alpha_j^{(i-1)} - k_i \alpha_{i-j}^{(i-1)}, \quad j = 1, 2, \dots, p$$

$$\alpha_i^{(i)} = k_i$$

Levinson-Durbin Algorithm 7

- The final solution for order p is:

$$\alpha_j = \alpha_j^{(p)} \quad 1 \leq j \leq p$$

- with prediction error

$$E^{(p)} = E[0] \prod_{m=1}^p (1 - k_m^2) = R[0] \prod_{m=1}^p (1 - k_m^2)$$

- If we use normalized autocorrelation coefficients:

$$r[k] = R[k] / R[0]$$

- we get normalized errors of the form:

$$\nu^{(i)} = \frac{E^{(i)}}{R[0]} = 1 - \sum_{k=1}^i \alpha_k^{(i)} r[k] = \prod_{m=1}^i (1 - k_m^2)$$

- where $0 < \nu^{(i)} \leq 1$ or $-1 < k_i < 1$

Levinson-Durbin Algorithm

Levinson-Durbin Algorithm

$$\mathcal{E}^{(0)} = R[0] \quad (9.98)$$

for $i = 1, 2, \dots, p$

$$k_i = \left(R[i] - \sum_{j=1}^{i-1} \alpha_j^{(i-1)} R[i-j] \right) / \mathcal{E}^{(i-1)} \quad (9.93)$$

$$\alpha_i^{(i)} = k_i \quad (9.96b)$$

if $i > 1$ then for $j = 1, 2, \dots, i-1$

$$\alpha_j^{(i)} = \alpha_j^{(i-1)} - k_i \alpha_{i-j}^{(i-1)} \quad (9.96a)$$

end

$$\mathcal{E}^{(i)} = (1 - k_i^2) \mathcal{E}^{(i-1)} \quad (9.94)$$

end

$$\alpha_j = \alpha_j^{(p)} \quad j = 1, 2, \dots, p \quad (9.97)$$

$$\Rightarrow A^{(i)}(z) = A^{(i-1)}(z) - k_i z^{-i} A^{(i-1)}(z^{-1})$$

Autocorrelation Example

- consider a simple $p = 2$ solution of the form

$$\begin{bmatrix} R(0) & R(1) \\ R(1) & R(0) \end{bmatrix} \begin{bmatrix} \alpha_1 \\ \alpha_2 \end{bmatrix} = \begin{bmatrix} R(1) \\ R(2) \end{bmatrix}$$

- with solution

$$E^{(0)} = R(0)$$

$$k_1 = R(1)/R(0)$$

$$\alpha_1^{(1)} = R(1)/R(0)$$

$$E^{(1)} = \frac{R^2(0) - R^2(1)}{R(0)}$$

Autocorrelation Example

$$k_2 = \frac{R(2)R(0) - R^2(1)}{R^2(0) - R^2(1)}$$

$$\alpha_2^{(2)} = \frac{R(2)R(0) - R^2(1)}{R^2(0) - R^2(1)}$$

$$\alpha_1^{(2)} = \frac{R(1)R(0) - R(1)R(2)}{R^2(0) - R^2(1)}$$

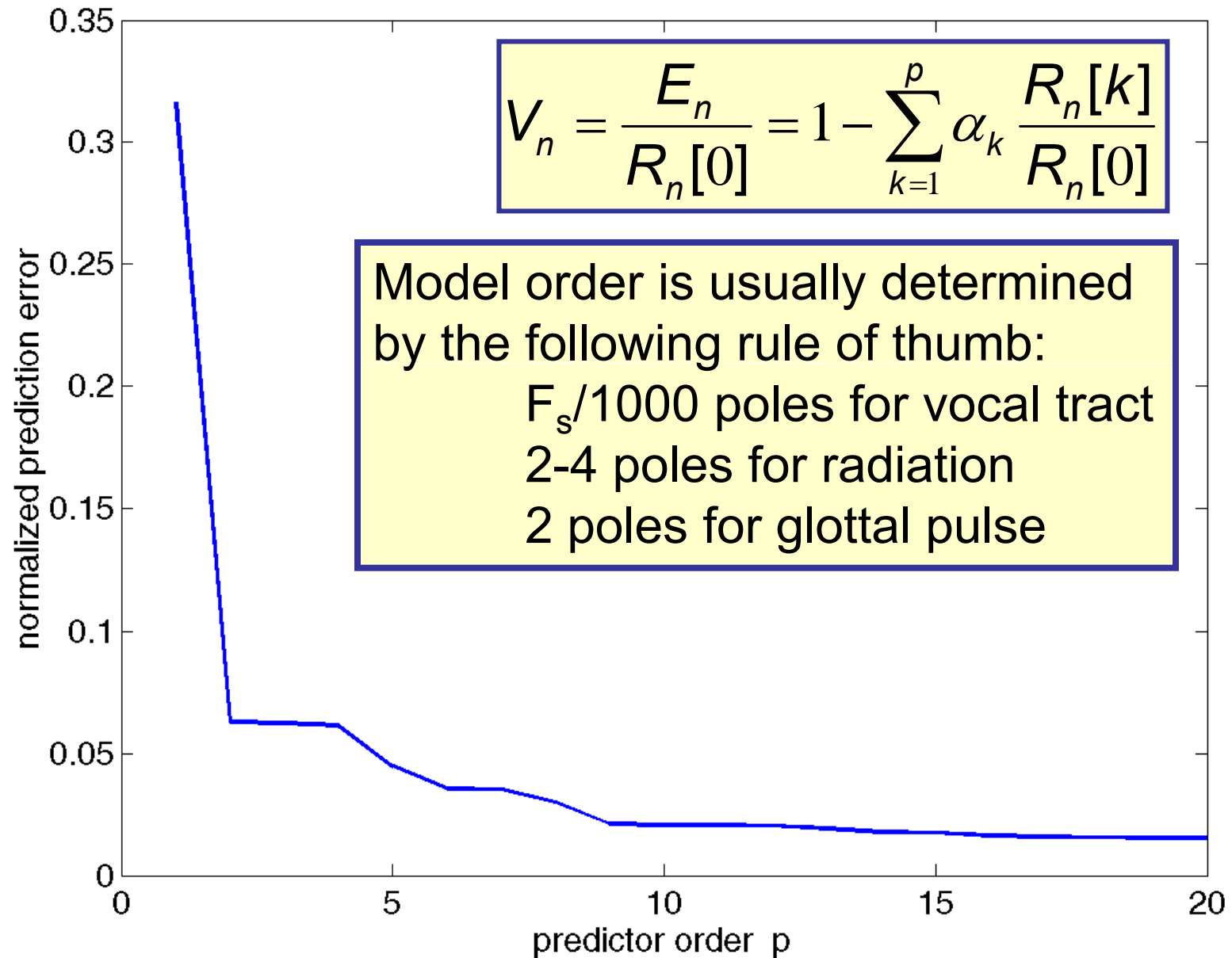
- with final coefficients

$$\alpha_1 = \alpha_1^{(2)}$$

$$\alpha_2 = \alpha_2^{(2)}$$

$E^{(i)}$ = prediction error for predictor of order i

Prediction Error as a Function of p



Autocorrelation Method Properties

- mean-squared prediction error always non-zero
 - decreases monotonically with increasing model order
- autocorrelation matching property
 - model and data match up to order p
- spectrum matching property
 - favors peaks of short-time FT
- minimum-phase property
 - zeros of $A(z)$ are inside the unit circle
- Levinson-Durbin recursion
 - efficient algorithm for finding prediction coefficients
 - PARCOR coefficients and MSE are by-products