## EVA Tutorial #1

## **BLOCK MAXIMA APPROACH UNDER NONSTATIONARITY**

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- (1) Traditional Methods/Rationale for Extreme Value Analysis
- (2) Max Stability/Extremal Types Theorem
- (3) Block Maxima Approach under Stationarity
- (4) Return Levels
- (5) Block Maxima Approach under Nonstationarity
- (6) Trends in Extremes
- (7) Other Forms of Covariates

# (1) Traditional Methods/Rationale for Extreme Value Analysis

- Fit models/distributions to all data
- -- Even if primary focus is on extremes
- Statistical theory for averages
- -- Ubiquitous role of normal distribution
- -- Central Limit Theorem for sums or averages

- Central Limit Theorem
- -- Given time series  $X_1, X_2, \ldots, X_n$

Assume independent and identically distributed (iid) Assume common cumulative distribution function (cdf) *F* Assume finite mean  $\mu$  and variance  $\sigma^2$ 

- -- Denote sum by  $S_n = X_1 + X_2 + \cdots + X_n$
- -- Then, no matter what shape of cdf F,

$$\Pr\{(S_n - n\mu) \mid n^{1/2} \sigma \leq x\} \to \Phi(x) \text{ as } n \to \infty$$

where  $\Phi$  denotes standard normal N(0, 1) cdf

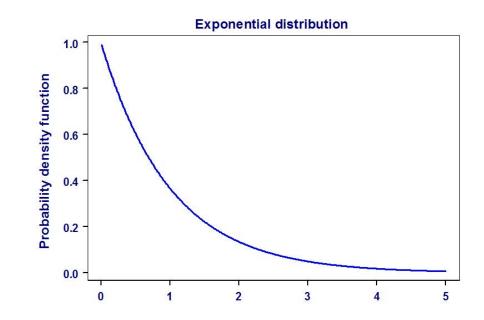
- Robustness
- -- Avoid sensitivity to extremes (outliers / contamination)
- Nonparametric Alternatives
- -- Kernel density estimation Ok for center of distribution (but not for lower & upper tails)
- -- Resampling
  - Fails for maxima
  - **Cannot extrapolate**

• Conduct sampling experiment

-- Exponential distribution with cdf

$$F(x) = 1 - \exp[-(x/\sigma)], x > 0, \sigma > 0$$

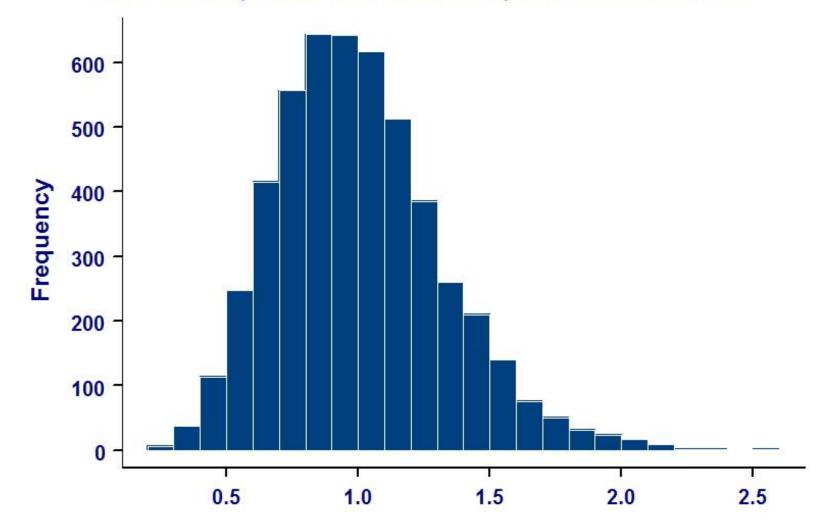
Here  $\sigma$  is scale parameter (also mean)



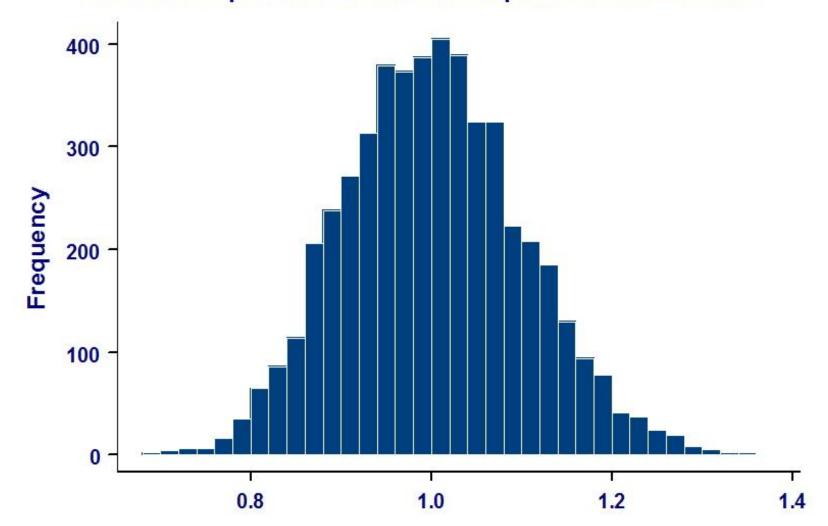
- -- Draw random samples of size n = 10 from exponential distribution (with  $\sigma = 1$ ) and calculate mean for each sample
  - (i) First pseudo random sample
    1.678, 0.607, 0.732, 1.806, 1.388, 0.630, 0.382, 0.396, 1.324, 1.148
    (Sample mean ≈ 1.009)
  - (ii) Second pseudo random sampleSample mean ≈ 0.571

(iii) Third pseudo random sampleSample mean ≈ 0.859

Repeat many more times



Mean of samples of size 10 from exponential distribution



Mean of samples of size 100 from exponential distribution

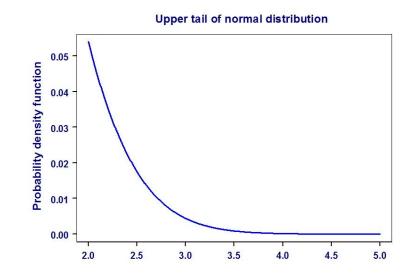
- Limited information about extremes
- -- Exploit what theory is available
- More robust/flexible approach
- -- Tail behavior of standard distributions is too restrictive Statistical theory indicates possibility of "heavy" tails Data suggest evidence of "heavy" tails Conventional distributions have "light" tails

#### -- Example

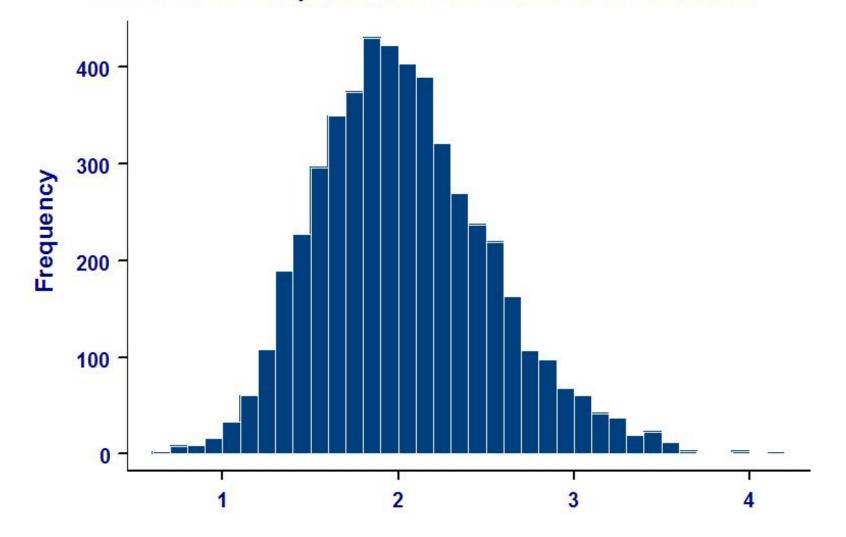
Let *X* have standard normal distribution [i. e., *N*(0, 1)] with probability density function (pdf)

 $\varphi(x) = (2\pi)^{-1/2} \exp(-x^2/2)$ 

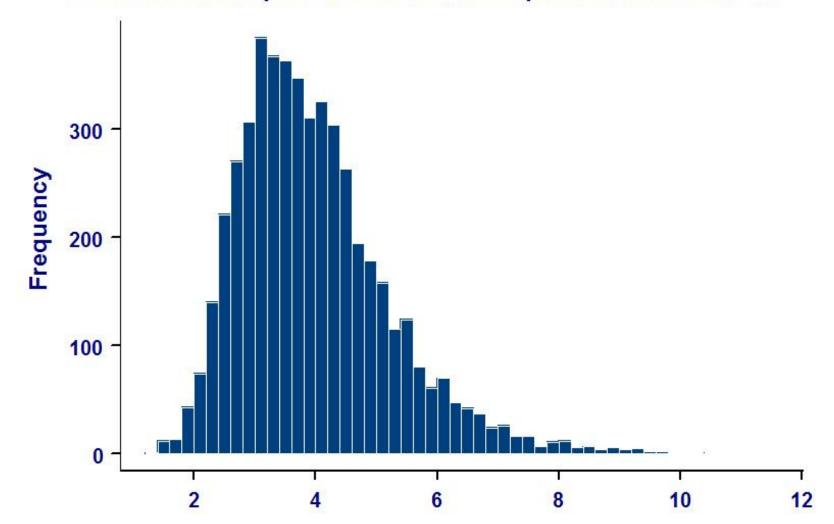
Then  $\Pr{X > x} \equiv 1 - \Phi(x) \approx \varphi(x) / x$ , for large x



- Statistical behavior of extremes
- -- Effectively no role for normal distribution
- -- What form of distribution(s) instead?
- Conduct another sampling experiment
- -- Calculate largest value of random sample (instead of mean)
  - (i) Standard normal distribution *N*(0, 1)
  - (ii) Exponential distribution ( $\sigma = 1$ )



Maximum of samples of size 30 from normal distribution



# Maximum of samples of size 30 from exponential distribution

- "Sum stability"
- -- Property of normal distribution

 $X_1, X_2, \ldots, X_n$  iid with common cdf  $N(\mu, \sigma^2)$ 

Then sum 
$$S_n = X_1 + X_2 + \cdots + X_n$$

is exactly normally distributed

In particular,  $(S_n - n\mu) / n^{1/2} \sigma$ 

has an exact N(0, 1) distribution

- "Max stability"
- -- Want to find distribution(s) for which maximum has same form as original sample

Note that

 $\max\{X_1, X_2, ..., X_{2n}\} =$ 

 $\max\{\max\{X_1, X_2, \ldots, X_n\}, \max\{X_{n+1}, X_{n+2}, \ldots, X_{2n}\}\}$ 

-- So cdf G, say, must satisfy

 $G^2(x) = G(ax + b)$ 

Here *a* > 0 and *b* are constants

Extremal Types Theorem

Time series  $X_1, X_2, \ldots, X_n$  assumed iid (for now)

Set  $M_n = \max\{X_1, X_2, ..., X_n\}$ 

Suppose that there exist constants  $a_n > 0$  and  $b_n$  such that

$$\Pr\{(M_n - b_n) \mid a_n \leq x\} \rightarrow G(x) \text{ as } n \rightarrow \infty$$

where G is a non-degenerate cdf

Then G must a generalized extreme value (GEV) cdf; that is,

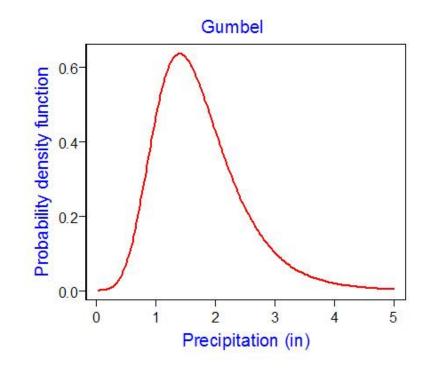
$$G(x; \mu, \sigma, \xi) = \exp \{-[1 + \xi (x - \mu)/\sigma]^{-1/\xi}\}, 1 + \xi (x - \mu)/\sigma > 0$$

 $\mu$  location parameter,  $\sigma > 0$  scale parameter,  $\xi$  shape parameter

(i)  $\xi = 0$  (*Gumbel type*, limit as  $\xi \rightarrow 0$ )

"Light" upper tail

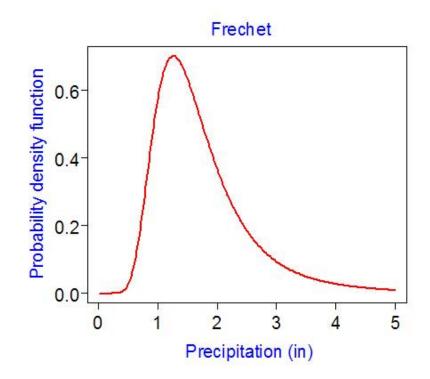
"Domain of attraction" for many common distributions (e.g., normal, exponential, gamma)



(ii)  $\xi > 0$  (*Fréchet type*)

"Heavy" upper tail with infinite *r*th-order moment if  $r \ge 1/\xi$ (e. g., infinite variance if  $\xi \ge 1/2$ )

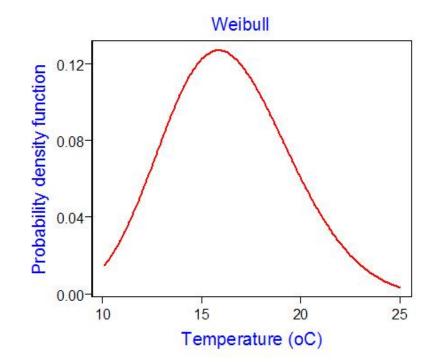
Fits precipitation, streamflow, economic damage

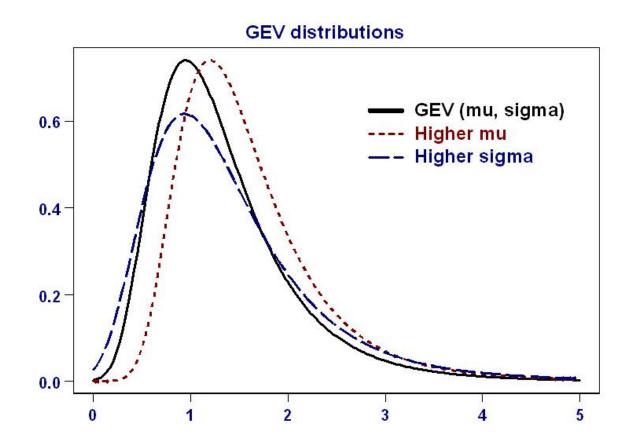


(iii)  $\xi < 0$  (Weibull type)

Bounded upper tail [ $x < \mu + \sigma / (-\xi)$ ]

Fits temperature, wind speed, sea level





Location parameter of GEV is not equivalent to mean

Scale parameter of GEV is *not* equivalent to standard deviation

- GEV distribution
- -- Fit directly to maxima (say with block size *n*)

e.g., annual maximum of daily precipitation amount or highest temperature over given year or annual peak stream flow

-- Advantages

Do not necessarily need to explicitly model annual and diurnal cycles

Do not necessarily need to explicitly model temporal dependence

Maximum likelihood estimation (mle)

-- Given observed block maxima  $X_1 = x_1, X_2 = x_2, \ldots, X_T = x_T$ 

-- Assume exact GEV dist. with pdf

$$g(x; \mu, \sigma, \xi) = G'(x; \mu, \sigma, \xi)$$

-- Likelihood function

 $L(x_1, x_2, ..., x_T; \mu, \sigma, \xi) = g(x_1; \mu, \sigma, \xi) g(x_2; \mu, \sigma, \xi) \cdots g(x_T; \mu, \sigma, \xi)$ 

#### Minimize

-In *L*(*x*<sub>1</sub>, *x*<sub>2</sub>, . . . , *x<sub>T</sub>*; μ, σ, ξ)

with respect to  $\mu$ ,  $\sigma$ ,  $\xi$ 

-- Likelihood ratio test (LRT)

For example, to test whether  $\xi = 0$  fit two models:

(i)  $-\ln L(x_1, x_2, \ldots, x_T; \mu, \sigma, \xi)$  minimized with respect to  $\mu, \sigma, \xi$ 

(ii)  $-\ln L(x_1, x_2, ..., x_T; \mu, \sigma, \xi = 0)$  minimized with respect to  $\mu, \sigma$ 

If  $\xi = 0$ , then 2 [(ii) – (i)] has approximate chi square distribution with 1 degree of freedom (df) for large *T* 

-- Confidence interval (e. g., for  $\xi$ ) based on "profile likelihood"

Minimize  $-\ln L(x_1, x_2, ..., x_T; \mu, \sigma, \xi)$  with respect to  $\mu, \sigma$  as function of  $\xi$ 

Use chi square dist. with 1 df

- Fort Collins daily precipitation amount
- -- Fort Collins, CO, USA

Time series of daily precipitation amount (in), 1900-1999

Semi-arid region

Marked annual cycle in precipitation

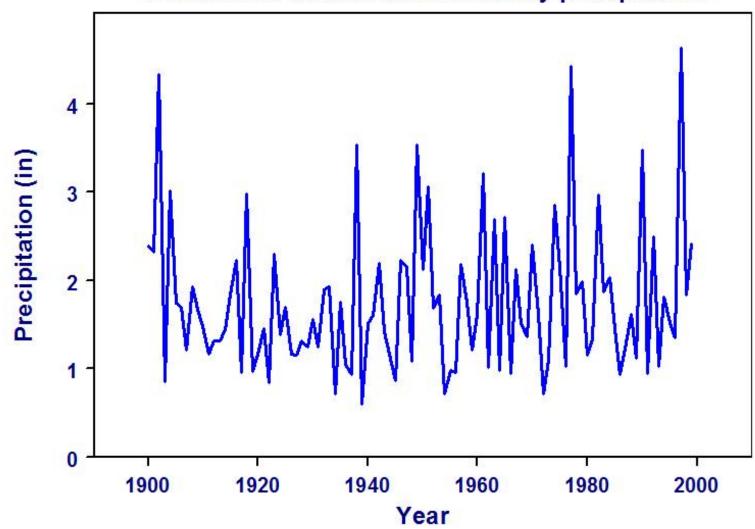
(peak in late spring/early summer, driest in winter)

Consider annual maxima (block size  $n \approx 365$ )

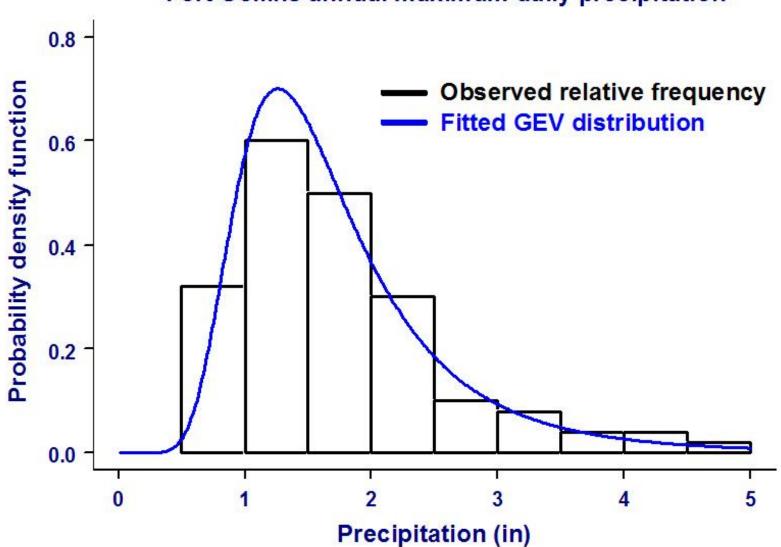
No obvious long-term trend in annual maxima (T = 100)

Flood on 28 July 1997

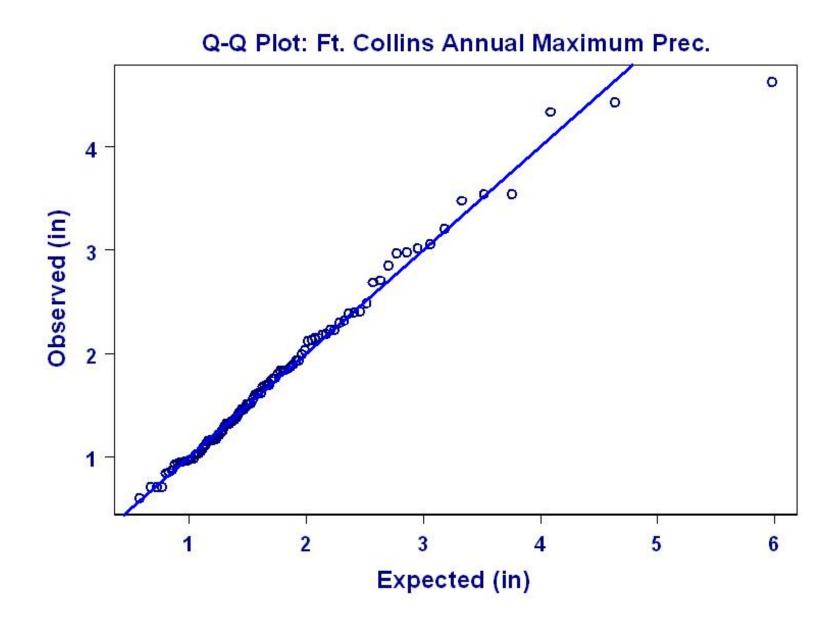
(Damaged campus of Colorado State Univ.)



Fort Collins annual maximum daily precipitation



#### Fort Collins annual maximum daily precipitation



• Parameter estimates and standard errors

Parameter <b>e</b>	<b>Estimate</b>	( <u>Std. Error)</u>
Location µ	1.347	(0.062)
Scale o	0.533	(0.049)
Shape ξ	0.174	(0.092)

-- LRT for 
$$\xi = 0$$
 (*P*-value  $\approx 0.038$ )

-- 95% confidence interval for shape parameter ξ
 (based on profile likelihood)

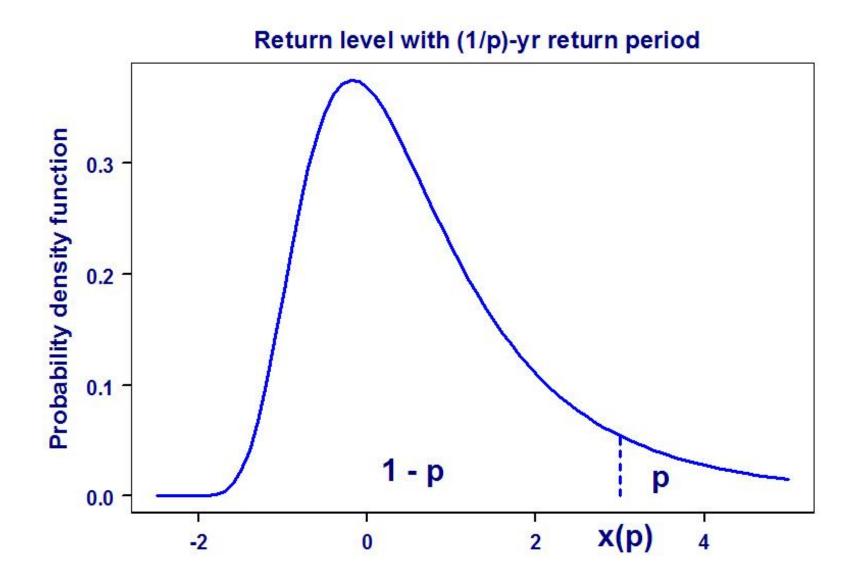
 $0.009 < \xi < 0.369$ 

- Assume stationarity
- -- i. e., unchanging climate
- Return period / Return level
- -- "Return level" with (1/p)-yr "return period"

$$x(p) = G^{-1}(1-p; \mu, \sigma, \xi), \ 0$$

Quantile of GEV cdf G

(e. g., p = 0.01 corresponds to 100-yr return period)



• GEV distribution

 $x(p) = \mu - (\sigma/\xi) \{1 - [-\ln(1 - p)]\}^{-\xi}$ 

Confidence interval: Re-parameterize replacing location parameter  $\mu$  with x(p) & use profile likelihood method

-- Fort Collins precipitation example (annual maxima)

Estimated 100-yr return level: 5.10 in

95% confidence interval (based on profile likelihood):

3.93 in < *x*(0.01) < 8.00 in

- Sources
- -- Trends

Associated with global climate change (e.g.)

-- Cycles

Annual & diurnal cycles (e.g.)

-- Physically-based

Use in statistical downscaling (e.g.)

- Theory
- -- No general extreme value theory under nonstationarity Only limited results under restrictive conditions
- Methods
- -- Introduction of covariates resembles "generalized linear models"
- -- Straightforward to extend maximum likelihood estimation
- Issues
- -- Nature of relationship between extremes & covariates Resembles that for overall / center of data?

- Trends
- -- Example (Urban heat island)

Trend in summer minimum temperature at Phoenix, AZ (i. e., block minima)

$$\min\{X_1, X_2, \ldots, X_n\} = -\max\{-X_1, -X_2, \ldots, -X_n\}$$

Assume negated summer minimum temperature in year *t* has GEV distribution with location and scale parameters:

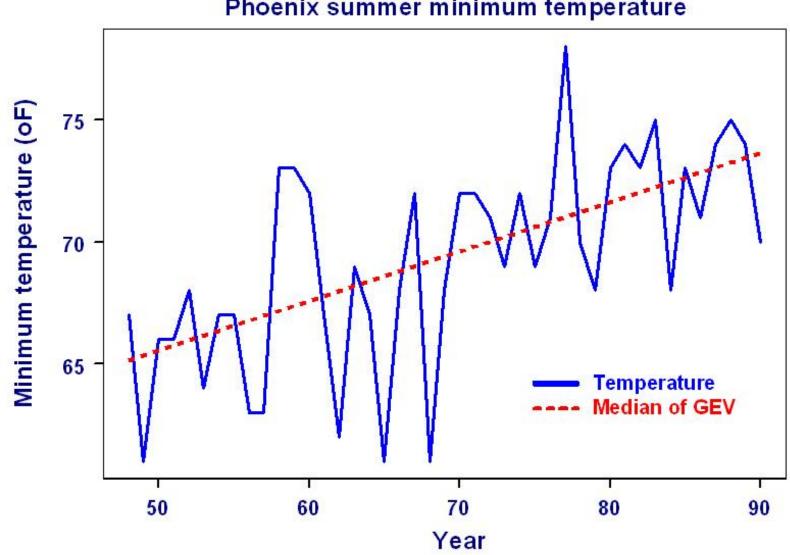
$$\mu(t) = \mu_0 + \mu_1 t$$
,  $\ln \sigma(t) = \sigma_0 + \sigma_1 t$ ,  $\xi(t) = \xi$ ,  $t = 1, 2, ...$ 

Parameter estimates and standard errors

<b>Parameter</b>		<b>Estimate</b>	( <u>Std. Error</u> )
Location:	μ <sub>0</sub>	66.17*	
	μ <sub>1</sub>	0.196*	(0.041)
Scale:	$\sigma_0$	1.338	
	<b>σ</b> 1	-0.009	(0.010)
Shape:	ξ	-0.211	

\*Sign of location parameters reversed to convert back to minima

- -- LRT for  $\mu_1 = 0$  (*P*-value <  $10^{-5}$ )
- -- LRT for  $\sigma_1 = 0$  (*P*-value  $\approx 0.366$ )



Phoenix summer minimum temperature

• Q-Q plots under non-stationarity

-- Transform to common distribution

Non-stationary GEV [ $\mu(t)$ ,  $\sigma(t)$ ,  $\xi(t)$ ]

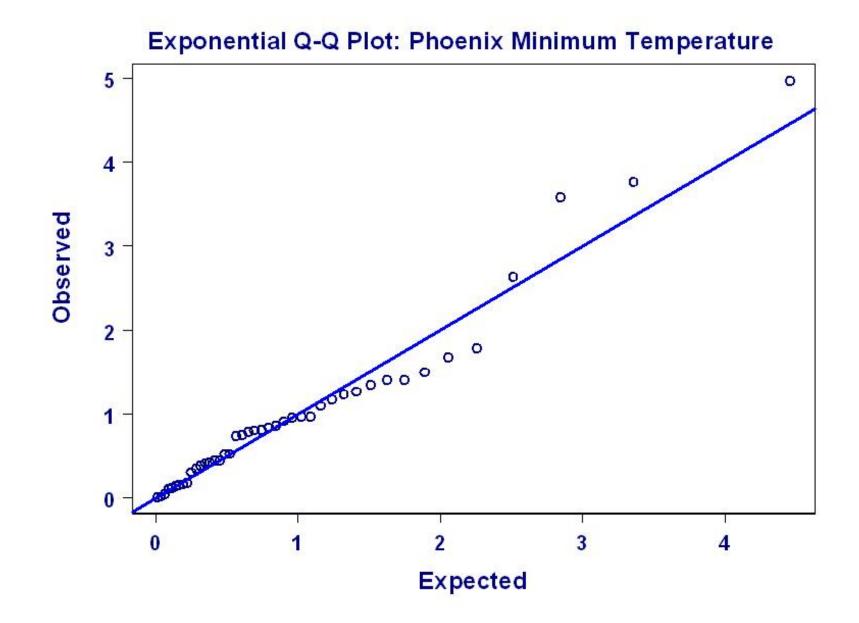
Not invariant to choice of transformation

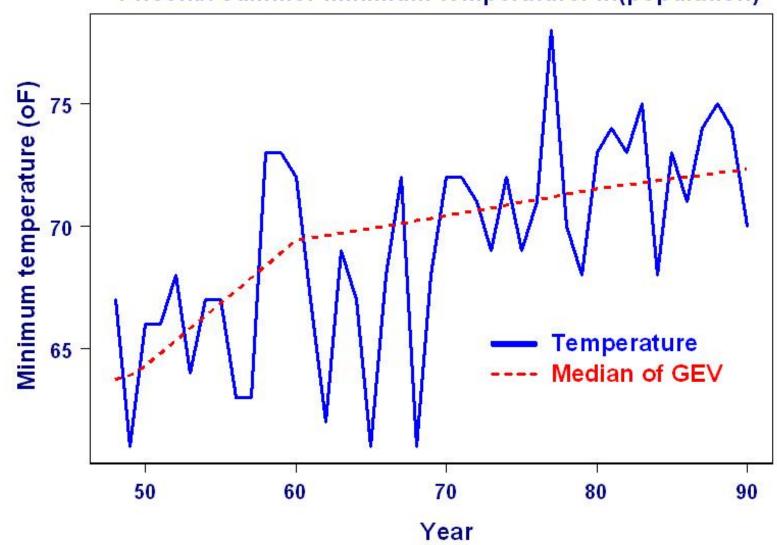
(i) Non-stationary GEV to standard exponential

$$\varepsilon_t = \{1 + \xi(t) [X_t - \mu(t)] / \sigma(t)\}^{-1/\xi(t)}$$

(ii) Non-stationary GEV to standard Gumbel (used by extRemes)

$$\varepsilon_t = [1/\xi(t)] \log \{1 + \xi(t) [X_t - \mu(t)] / \sigma(t)\}$$





Phoenix summer minimum temperature: In(population)

- Physically-based covariates
- -- Example [Arctic Oscillation (AO)]

Winter maximum temperature at Port Jervis, NY, USA

(i. e., block maxima)

Z denotes winter index of AO

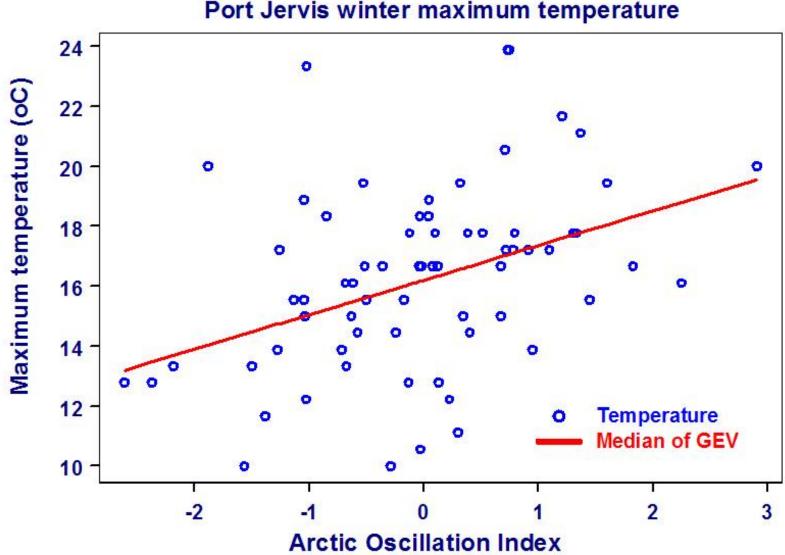
Given Z = z, assume conditional distribution of winter maximum temperature is GEV distribution with parameters:

$$\mu(z) = \mu_0 + \mu_1 z$$
,  $\ln \sigma(z) = \sigma_0 + \sigma_1 z$ ,  $\xi(z) = \xi$ 

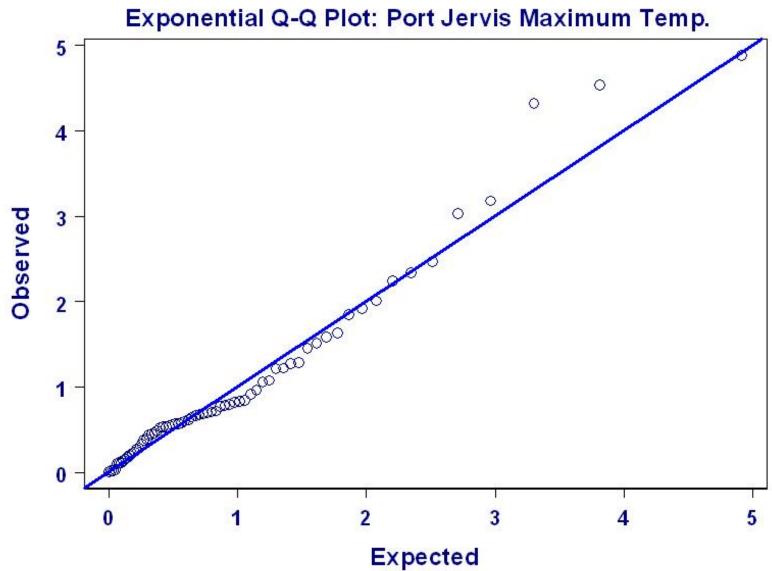
• Parameter estimates and standard errors

<u>Param</u>	<u>eter</u>	<b>Estimate</b>	( <u>Std. Error</u> )
Location:	μ <sub>0</sub>	15.26	
	μ <sub>1</sub>	1.175	(0.319)
Scale:	σ <sub>0</sub>	0.984	
	<b>σ</b> <sub>1</sub>	-0.044	(0.092)
Shape:	ξ	-0.186	

- -- LRT for  $\mu_1 = 0$  (*P*-value < 0.001)
- -- LRT for  $\sigma_1 = 0$  (*P*-value  $\approx 0.635$ )



Port Jervis winter maximum temperature



A random variable X has a *lognormal distribution* if the logtransformed variable

 $Y = \ln X$ 

has a normal distribution. Then Y is in the domain of attraction of the Gumbel type.

What is the domain of attraction of *X*? (i. e., Gumbel, Fréchet, or Weibull type?)

<u>Answer</u>: X is in the domain of attraction of the Gumbel type.