#### Introduction to statistics: Linear mixed models

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## Summary

- 1. We know how to do simple t-tests.
- 2. We know how to fit simple linear models.
- 3. We saw that the paired t-test is identical to the varying intercepts linear mixed model.

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Now we are ready to look at linear mixed models in detail.

Returning to our SR/OR relative clause data from English (Grodner and Gibson, Expt 1). First we load the data as usual (not shown).

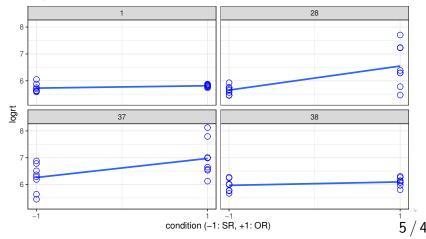
gge1crit\$so<-ifelse(gge1crit\$condition=="objgap",1,-1)</pre>

```
dat<- gge1crit
dat$logrt<-log(dat$rawRT)</pre>
```

The simple linear model (incorrect for these data):

summary(m0<-lm(logrt~so,dat))\$coefficients
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) 5.883056 0.019052 308.7841 0.000000
## so 0.062017 0.019052 3.2551 0.0011907

We can visualize the different responses of subjects (four subjects shown):



Given these differences between subjects, you could fit a separate linear model for each subject, collect together the intercepts and slopes for each subject, and then check if the intercepts and slopes are significantly different from zero.

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We will fit the model using log reading times because we want to make sure we satisfy model assumptions (e.g., normality of residuals).

There is a function in the package lme4 that computes separate linear models for each subject: lmList.

library(lme4)

## Loading required package: Matrix

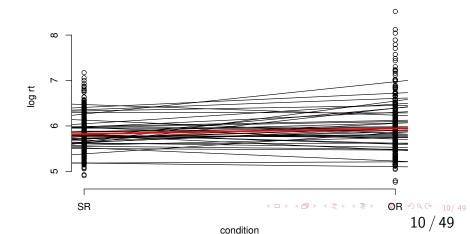
lmlist.fm1<-lmList(logrt~so|subject,dat)</pre>

Intercept and slope estimates for three subjects:

lmlist.fm1\$`1`\$coefficients ## (Intercept) SO ## 5.769617 0.043515 lmlist.fm1\$`28`\$coefficients ## (Intercept) SO ## 6.10021 0.44814 lmlist.fm1\$`37`\$coefficients ## (Intercept) SO 6.61699 0.35537 ##

One can plot the individual lines for each subject, as well as the linear model m0's line (this shows how each subject deviates in intercept and slope from the model m0's intercept and slopes).

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To find out if there is an effect of RC type, you can simply check whether the slopes of the individual subjects' fitted lines taken together are significantly different from zero.

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```
t.test(coef(lmlist.fm1)[2])
##
    One Sample t-test
##
##
## data: coef(lmlist.fm1)[2]
## t = 2.81, df = 41, p-value = 0.0076
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
## 0.017449 0.106585
## sample estimates:
## mean of x
## 0.062017
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```

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The above test is exactly the same as the paired t-test and the varying intercepts linear mixed model **on aggregated data**:

t.test(logrt~condition,bysubj,paired=TRUE)\$statistic

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## t ## 2.8102

## Estimate Std. Error t value
## -0.124033 0.044137 -2.810207

- The above ImList model we fit is called repeated measures regression. We now look at how to model unaggregated data using the linear mixed model.
- This model is now only of historical interest, and useful only for understanding the linear mixed model, which is the modern standard approach.

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└─Linear mixed models └─Model type 1: Varying intercepts models

#### Linear mixed models

- The linear mixed model does something related to the above by-subject fits, but with some crucial twists, as we see below.
- In the model shown in the next slide, the statement (1|subject) adjusts the grand mean estimates of the intercept by a term (a number) for each subject.

Model type 1: Varying intercepts models

#### Linear mixed models

#### Notice that we did not aggregate the data here.

m0.lmer<-lmer(logrt~so+(1|subject),dat)</pre>

Abbreviated output:

```
Random effects:
                  Variance Std.Dev.
 Groups Name
 subject (Intercept) 0.09983 0.3160
 Residual
                      0.14618 0.3823
Number of obs: 672, groups: subject, 42
Fixed effects:
            Estimate Std. Error t value
(Intercept) 5.88306 0.05094 115.497
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            0.06202 0.01475 4.205
SO
                                                   16 / 49
```

Model type 1: Varying intercepts models

#### Linear mixed models

One thing to notice is that the coefficients (intercept and slope) of the fixed effects of the above model are identical to those in the linear model m0 above.

The varying intercepts for each subject can be viewed by typing:

```
ranef(m0.lmer)$subject[,1][1:10]
```

## [1] -0.1039283 0.0771948 -0.2306209 0.2341978 0.0088
## [7] -0.2055713 -0.1553708 0.0759436 -0.3643671

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Model type 1: Varying intercepts models

#### Visualizing random effects

Here is another way to summarize the adjustments to the grand mean intercept by subject. The error bars represent 95% confidence intervals.

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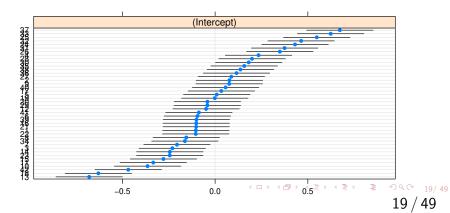
library(lattice)
print(dotplot(ranef(m0.lmer,condVar=TRUE)))

└─Model type 1: Varying intercepts models

## Visualizing random effects







Model type 1: Varying intercepts models

#### Linear mixed models

The model m0.Imer above prints out the following type of linear model. i indexes subject, and j indexes items. Once we know the subject id and the item id, we know which subject saw which condition:

subset(dat,subject==1 & item == 1)
## subject item condition rawRT so logrt
## 6 1 1 objgap 320 1 5.7683

$$y_{ij} = \beta_0 + u_{0i} + \beta_1 \times so_{ij} + \epsilon_{ij} \tag{1}$$

Model type 1: Varying intercepts models

#### Linear mixed models

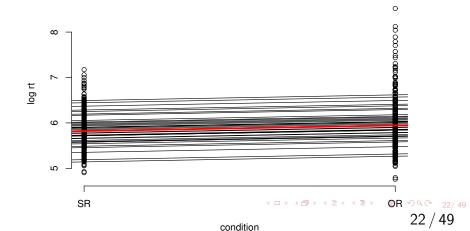
- Note that these by-subject adjustments to the intercept u<sub>0i</sub> are assumed by lmer to come from a normal distribution centered around 0: u<sub>0i</sub> ~ Normal(0, σ<sub>u0</sub>)
- The ordinary linear model m0 has one intercept β<sub>0</sub> for all subjects, whereas the linear mixed model with varying intercepts m0.Imer has a different intercept (β<sub>0</sub> + u<sub>0i</sub>) for each subject *i*.
- We can visualize the adjustments for each subject to the intercepts as shown below.

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Linear mixed models

└─Model type 1: Varying intercepts models

#### Linear mixed models



Model type 1: Varying intercepts models

# Formal statement of varying intercepts linear mixed model

i indexes subjects, j items.

$$y_{ij} = \beta_0 + u_{0i} + (\beta_1) \times so_{ij} + \epsilon_{ij}$$
(2)

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Variance components:

- $\blacktriangleright$   $u_0 \sim Normal(0, \sigma_{u0})$
- $\blacktriangleright \ \epsilon \sim Normal(0,\sigma)$

Model type 2: Varying intercepts and slopes model (no correlation)

#### Linear mixed models

Note that, unlike the figure associated with the Imlist.fm1 model above, which also involves fitting separate models for each subject, the model m0.lmer assumes **different intercepts** for each subject **but the same slope**.

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We can have Imer fit different intercepts AND slopes for each subject.

└─Model type 2: Varying intercepts and slopes model (no correlation)

#### Linear mixed models

Varying intercepts and slopes by subject

We assume now that each subject's slope is also adjusted:

$$y_{ij} = \beta_0 + u_{0i} + (\beta_1 + u_{1i}) \times so_{ij} + \epsilon_{ij}$$
(3)

That is, we additionally assume that  $u_{1i} \sim Normal(0, \sigma_{u1})$ .

m1.lmer<-lmer(logrt~so+(1+so||subject),dat)</pre>

#### Random effects:

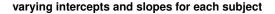
Groups	Name	Variance	Std.Dev.		
subject	(Intercept)	0.1006	0.317		
subject.1	SO	0.0121	0.110		
Residual		0.1336	0.365		
Number of d	obs: 672, gro	oups: sub	oject, 42		
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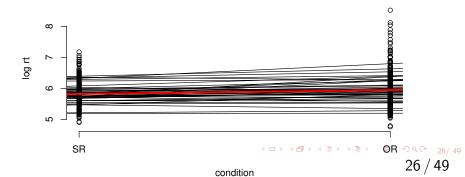
Fixed effects:

Model type 2: Varying intercepts and slopes model (no correlation)

#### Linear mixed models

These fits for each subject are visualized below (the red line shows the model with a single intercept and slope, i.e., our old model m0):





Model type 2: Varying intercepts and slopes model (no correlation)

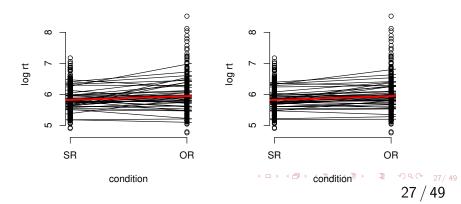
#### Linear mixed models

Comparing ImList model with varying intercepts model

Compare this model with the Imlist.fm1 model we fitted earlier:

ordinary linear model

varying intercepts and slopes



└─Model type 2: Varying intercepts and slopes model (no correlation)

## Visualizing random effects

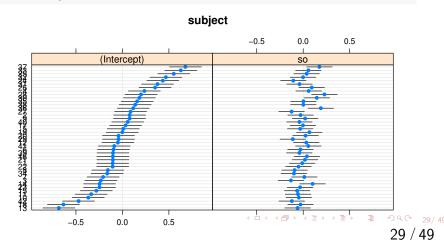
print(dotplot(ranef(m1.lmer,condVar=TRUE)))

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Model type 2: Varying intercepts and slopes model (no correlation)

#### Visualizing random effects

## \$subject



Model type 2: Varying intercepts and slopes model (no correlation)

# Formal statement of varying intercepts and varying slopes linear mixed model

i indexes subjects, j items.

$$y_{ij} = \beta_0 + u_{0i} + (\beta_1 + u_{1i}) \times so_{ij} + \epsilon_{ij}$$

$$\tag{4}$$

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Variance components:

- $\blacktriangleright$   $u_0 \sim Normal(0, \sigma_{u0})$
- $\blacktriangleright$   $u_1 \sim Normal(0, \sigma_{u1})$
- $\blacktriangleright \ \epsilon \sim Normal(0,\sigma)$

Shrinkage in linear mixed models

#### Shrinkage in linear mixed models

- The estimate of the effect by participant is smaller than when we fit a separate linear model to the subject's data.
- This is called shrinkage in linear mixed models: the individual level estimates are shunk towards the mean slope.

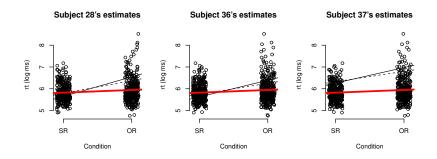
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The less data we have from a given subject, the more the shrinkage.

Shrinkage in linear mixed models

#### Shrinkage in linear mixed models



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Shrinkage in linear mixed models

# Shrinkage in linear mixed models

The effect of missing data on estimation in LMMs

Let's randomly delete some data from one subject:

set.seed(4321)
## choose some data randomly to remove:
rand<-rbinom(1,n=16,prob=0.5)</pre>



Shrinkage in linear mixed models

#### Shrinkage in linear mixed models The effect of missing data on estimation in LMMs

```
dat[which(dat$subject==37),]$rawRT
```

```
## [1] 770 536 686 578 457 487 2419 884 3365 233
## [15] 1081 971
```

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```
dat$deletedRT<-dat$rawRT
dat[which(dat$subject==37),]$deletedRT<-
    ifelse(rand,NA,
        dat[which(dat$subject==37),]$rawRT)</pre>
```

└-Shrinkage in linear mixed models

# Shrinkage in linear mixed models

The effect of missing data on estimation in LMMs

Now subject 37's estimates are going to be pretty wild:

<pre>subset(dat,subject==37)\$deletedRT</pre>										
##	[1]	770	NA	686	578	NA	NA	NA	NA 3365	233
##	[15]	NA	971							

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Shrinkage in linear mixed models

#### Shrinkage in linear mixed models The effect of missing data on estimation in LMMs

```
## original no pooling estimate:
lmList.fm1_old<-lmList(log(rawRT)~so|subject,dat)
coefs_old<-coef(lmList.fm1_old)
intercepts_old<-coefs_old[1]
colnames(intercepts_old)<-"intercept"
slopes_old<-coefs_old[2]
## subject 37's original estimates:
intercepts_old$intercept[37]
```

```
## [1] 6.617
```

slopes\_old\$so[37]

## [1] 0.35537

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Shrinkage in linear mixed models

#### Shrinkage in linear mixed models The effect of missing data on estimation in LMMs

```
## on deleted data:
lmList.fm1_deleted<-lmList(log(deletedRT)~so|subject,dat)</pre>
coefs<-coef(lmList.fm1_deleted)</pre>
intercepts<-coefs[1]
colnames(intercepts)<-"intercept"</pre>
slopes<-coefs[2]</pre>
## subject 37's new estimates on deleted data:
intercepts$intercept[37]
## [1] 6.6879
```

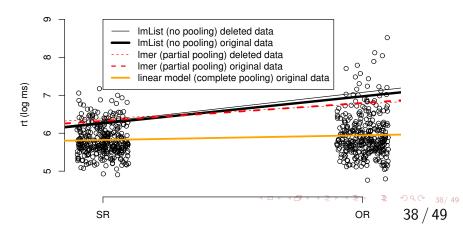
slopes\$so[37]

## [1] 0.38843

Shrinkage in linear mixed models

#### Shrinkage in linear mixed models The effect of missing data on estimation in LMMs

#### Subject 37's estimates



Shrinkage in linear mixed models

## Shrinkage in linear mixed models

The effect of missing data on estimation in LMMs

- What we see here is that the estimates from the hierarchical model are barely affected by the missingness, but the estimates from the no-pooling model are heavily affected.
- This means that linear mixed models will give you more robust estimates (think Type M error!) compared to no pooling models.

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This is one reason why linear mixed models are such a big deal.

└─Varying intercepts and slopes model, with crossed random effects for subjects and for items

#### Crossed subjects and items in LMMs

Subjects and items are fully crossed:

head(xtabs(~subject+item,dat))

##		it	ce	m															
##	subject	1	L	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	
##	1	1	L	1	1	1	1	1	1	1	1	1	1	1	1	1	1	1	
##	2	1	L	1	1	1	1	1	1	1	1	1	1	1	1	1	1	1	
##	3	1	L	1	1	1	1	1	1	1	1	1	1	1	1	1	1	1	
##	4	: 1	L	1	1	1	1	1	1	1	1	1	1	1	1	1	1	1	
##	5	1	L	1	1	1	1	1	1	1	1	1	1	1	1	1	1	1	
##	6	1	L	1	1	1	1	1	1	1	1	1	1	1	1	1	1	1	

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-Varying intercepts and slopes model, with crossed random effects for subjects and for items

#### Linear mixed models

Linear mixed model with crossed subject and items random effects.

m2.lmer<-lmer(logrt~so+(1+so||subject)+(1+so||item),dat)</pre>

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L-Varying intercepts and slopes model, with crossed random effects for subjects and for items

### Linear mixed models

#### Random effects:

Groups	Name	Variance	Std.Dev.
subject	(Intercept)	0.10090	0.3177
subject.1	SO	0.01224	0.1106
item	(Intercept)	0.00127	0.0356
item.1	SO	0.00162	0.0402
Residual		0.13063	0.3614
Number of o	obs: 672, gro	oups: sub	oject, 42; item, 16

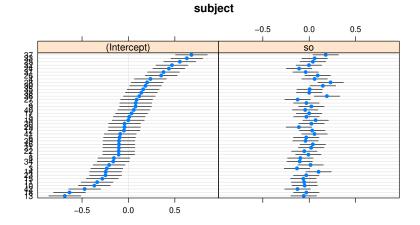
Fixed effects:

	CStimate	Stu. EIIOI	t varue				
(Intercept)	5.8831	0.0517	113.72				
SO	0.0620	0.0242	2.56	ヨト・ヨト	1	)	42/49

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L-Varying intercepts and slopes model, with crossed random effects for subjects and for items

#### Visualizing random effects

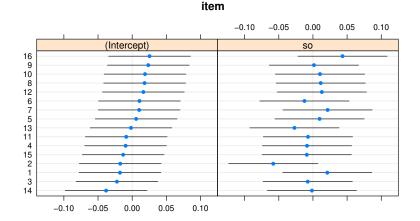


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-Varying intercepts and slopes model, with crossed random effects for subjects and for items

## Visualizing random effects



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Model type 3: Varying intercepts and varying slopes, with correlation

### Linear mixed models

Linear mixed model with crossed subject and items random effects.

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## boundary (singular) fit: see ?isSingular

Model type 3: Varying intercepts and varying slopes, with correlation

#### Linear mixed models

Linear mixed model with crossed subject and items random effects.

Random effects:

Groups	Name	Variance	<pre>Std.Dev.</pre>	Corr	
subject	(Intercept)	0.10103	0.3178		
	SO	0.01228	0.1108	0.58	
item	(Intercept)	0.00172	0.0415		
	SO	0.00196	0.0443	1.00 <=	degenerate
Residual		0.12984	0.3603		
Number of	obs: 672, gi	coups: su	ubject, 4	2; item,	16

Fixed effects: Estimate Std. Error t value (Intercept) 5.8831 0.0520 113.09 so 0.0620 0.0247 \*\*\* 2.51\*\*\*\* 2 \*\*\* 46/49 46/49

Model type 3: Varying intercepts and varying slopes, with correlation

# Formal statement of varying intercepts and varying slopes linear mixed model with correlation

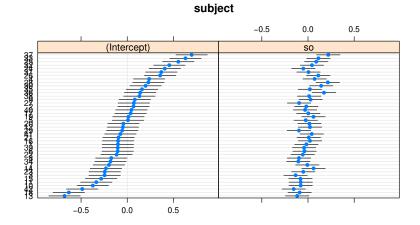
i indexes subjects, j items.

$$y_{ij} = \alpha + u_{0i} + w_{0j} + (\beta + u_{1i} + w_{1j}) * so_{ij} + \varepsilon_{ij}$$
(5)  
where  $\varepsilon_{ij} \sim Normal(0, \sigma)$  and

$$\Sigma_{u} = \begin{pmatrix} \sigma_{u0}^{2} & \rho_{u}\sigma_{u0}\sigma_{u1} \\ \rho_{u}\sigma_{u0}\sigma_{u1} & \sigma_{u1}^{2} \end{pmatrix} \qquad \Sigma_{w} = \begin{pmatrix} \sigma_{w0}^{2} & \rho_{w}\sigma_{w0}\sigma_{w1} \\ \rho_{w}\sigma_{w0}\sigma_{w1} & \sigma_{w1}^{2} \end{pmatrix}$$
(6)  
$$\begin{pmatrix} u_{0} \\ u_{1} \end{pmatrix} \sim \mathcal{N}\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \Sigma_{u} \right), \quad \begin{pmatrix} w_{0} \\ w_{1} \end{pmatrix} \sim \mathcal{N}\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \Sigma_{w} \right)$$
(7)

Model type 3: Varying intercepts and varying slopes, with correlation

## Visualizing random effects



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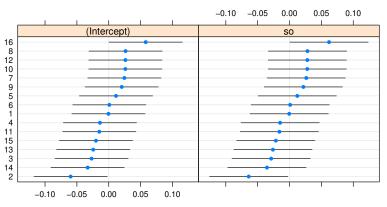
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Model type 3: Varying intercepts and varying slopes, with correlation

#### Visualizing random effects

#### These are degenerate estimates



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