# Introduction to statistics: Linear mixed models 

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## Summary

1. We know how to do simple t-tests.
2. We know how to fit simple linear models.
3. We saw that the paired t-test is identical to the varying intercepts linear mixed model.
Now we are ready to look at linear mixed models in detail.

## Linear models

Returning to our SR/OR relative clause data from English (Grodner and Gibson, Expt 1). First we load the data as usual (not shown).

```
gge1crit<-read.table("data/grodnergibson05data.txt",
    header=TRUE)
```

gge1crit\$so<-ifelse(gge1crit\$condition=="objgap", 1, -1)
dat<- gge1crit
dat\$logrt<-log(dat\$rawRT)
bysubj<-aggregate(logrt~subject+condition,
mean,data=dat)

## Linear models

The simple linear model (incorrect for these data):
summary(m0<-lm(logrt~so,dat))\$coefficients

| \#\# | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|t\|)$ |
| :--- | ---: | ---: | ---: | ---: |
| \#\# (Intercept) | 5.883056 | 0.019052 | 308.7841 | 0.0000000 |
| \#\# so | 0.062017 | 0.019052 | 3.2551 | 0.0011907 |

## Linear models

We can visualize the different responses of subjects (four subjects shown):


## Linear models

Given these differences between subjects, you could fit a separate linear model for each subject, collect together the intercepts and slopes for each subject, and then check if the intercepts and slopes are significantly different from zero.
We will fit the model using log reading times because we want to make sure we satisfy model assumptions (e.g., normality of residuals).

## Linear models

There is a function in the package lme4 that computes separate linear models for each subject: lmList.
library(lme4)
\#\# Loading required package: Matrix
lmlist.fm1<-lmList(logrt~solsubject, dat)

## Linear models

Intercept and slope estimates for three subjects:
lmlist.fm1\$1`\$coefficients

| \#\# (Intercept) | so |
| :---: | :---: |
| \#\# 5.769617 | 0.043515 |
| lmlist.fm1\$`28`\$coefficien |  |
| \#\# (Intercept) | so |
| \#\# 6.10021 | 0.44814 |
| lmlist.fm1\$`37`\$coefficient |  |
| (Intercept) | O |
| \#\# 6.61699 | 0.35537 |

## Linear models

One can plot the individual lines for each subject, as well as the linear model m0's line (this shows how each subject deviates in intercept and slope from the model m0's intercept and slopes).

## Linear models



## Linear models

To find out if there is an effect of RC type, you can simply check whether the slopes of the individual subjects' fitted lines taken together are significantly different from zero.

## Linear models

```
t.test(coef(lmlist.fm1)[2])
```

\#\#
\#\# One Sample t-test
\#\#
\#\# data: coef(lmlist.fm1) [2]
\#\# t = 2.81, df = 41, p-value = 0.0076
\#\# alternative hypothesis: true mean is not equal to 0
\#\# 95 percent confidence interval:
\#\# 0.0174490 .106585
\#\# sample estimates:
\#\# mean of $x$
\#\# 0.062017

## Linear models

The above test is exactly the same as the paired t-test and the varying intercepts linear mixed model on aggregated data:

```
t.test(logrt~condition, bysubj,paired=TRUE)$statistic
## t
## 2.8102
## also compare with linear mixed model:
summary(lmer(logrt~condition+(1|subject),
                                    bysubj))$coefficients[2,]
```

```
## Estimate Std. Error t value
## -0.124033 0.044137 -2.810207
```


## Linear models

- The above ImList model we fit is called repeated measures regression. We now look at how to model unaggregated data using the linear mixed model.
- This model is now only of historical interest, and useful only for understanding the linear mixed model, which is the modern standard approach.


## Linear mixed models

- The linear mixed model does something related to the above by-subject fits, but with some crucial twists, as we see below.
- In the model shown in the next slide, the statement (1|subject) adjusts the grand mean estimates of the intercept by a term (a number) for each subject.


## Linear mixed models

Notice that we did not aggregate the data here.
m0.lmer<-lmer (logrt~so+(1|subject), dat)
Abbreviated output:
Random effects:

| Groups | Name | Variance | Std.Dev. |
| :--- | :--- | :--- | :--- |
| subject | (Intercept) | 0.09983 | 0.3160 |
| Residual | 0.14618 | 0.3823 |  |

Number of obs: 672, groups: subject, 42

Fixed effects:
Estimate Std. Error t value

| (Intercept) | 5.88306 | 0.05094 | 115.497 |
| :--- | :--- | :--- | ---: |
| so | 0.06202 | 0.01475 | 4.205 |

## Linear mixed models

One thing to notice is that the coefficients (intercept and slope) of the fixed effects of the above model are identical to those in the linear model m0 above.
The varying intercepts for each subject can be viewed by typing:

```
ranef(m0.lmer)$subject[,1] [1:10]
## [1] -0.1039283 0.0771948 -0.2306209 0.2341978 0.0088
## [7] -0.2055713 -0.1553708 0.0759436 -0.3643671
```


## Visualizing random effects

Here is another way to summarize the adjustments to the grand mean intercept by subject. The error bars represent 95\% confidence intervals.
library(lattice)
print(dotplot(ranef(m0.lmer, condVar=TRUE)))

## Visualizing random effects

\#\# \$subject
subject


## Linear mixed models

The model m0.Imer above prints out the following type of linear model. i indexes subject, and j indexes items.
Once we know the subject id and the item id, we know which subject saw which condition:

```
subset(dat,subject==1 & item == 1)
```

\#\# subject item condition rawRT so logrt

| \#\# 6 | 1 | 1 | objgap | 320 | 1 | 5.7683 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |

$$
\begin{equation*}
y_{i j}=\beta_{0}+u_{0 i}+\beta_{1} \times s o_{i j}+\epsilon_{i j} \tag{1}
\end{equation*}
$$

The only new thing here is the by-subject adjustment to the intercept.

## Linear mixed models

- Note that these by-subject adjustments to the intercept $u_{0 i}$ are assumed by Imer to come from a normal distribution centered around 0 :

```
u}0i~Normal(0,\mp@subsup{\sigma}{u0}{}
```

- The ordinary linear model m 0 has one intercept $\beta_{0}$ for all subjects, whereas the linear mixed model with varying intercepts m 0 .Imer has a different intercept $\left(\beta_{0}+u_{0 i}\right)$ for each subject $i$.
- We can visualize the adjustments for each subject to the intercepts as shown below.

Lecture 6
$L_{\text {Linear mixed models }}$
-Model type 1: Varying intercepts models

## Linear mixed models

Formal statement of varying intercepts linear mixed model i indexes subjects, j items.

$$
\begin{equation*}
y_{i j}=\beta_{0}+u_{0 i}+\left(\beta_{1}\right) \times s o_{i j}+\epsilon_{i j} \tag{2}
\end{equation*}
$$

Variance components:

- $u_{0} \sim \operatorname{Normal}\left(0, \sigma_{u 0}\right)$
- $\epsilon \sim \operatorname{Normal}(0, \sigma)$


## Linear mixed models

Note that, unlike the figure associated with the Imlist.fm1 model above, which also involves fitting separate models for each subject, the model m0.Imer assumes different intercepts for each subject but the same slope.
We can have Imer fit different intercepts AND slopes for each subject.

## Linear mixed models

Varying intercepts and slopes by subject
We assume now that each subject's slope is also adjusted:

$$
\begin{equation*}
y_{i j}=\beta_{0}+u_{0 i}+\left(\beta_{1}+u_{1 i}\right) \times s o_{i j}+\epsilon_{i j} \tag{3}
\end{equation*}
$$

That is, we additionally assume that $u_{1 i} \sim \operatorname{Normal}\left(0, \sigma_{u 1}\right)$.
m1.lmer<-lmer (logrt~so+(1+sol|subject), dat)

Random effects:

| Groups | Name | Variance | Std.Dev. |
| :--- | :--- | :--- | :--- |
| subject | (Intercept) | 0.1006 | 0.317 |
| subject. | so | 0.0121 | 0.110 |
| Residual | 0.1336 | 0.365 |  |

Number of obs: 672, groups: subject, 42

Fixed effects:

## Linear mixed models

These fits for each subject are visualized below (the red line shows the model with a single intercept and slope, i.e., our old model m 0 ):
varying intercepts and slopes for each subject


## Linear mixed models

Comparing ImList model with varying intercepts model
Compare this model with the Imlist.fm1 model we fitted earlier:

## ordinary linear model


condition
varying intercepts and slopes


## Visualizing random effects

print(dotplot(ranef(m1.lmer, condVar=TRUE)))
$L_{\text {Linear mixed models }}$
$\square_{\text {Model type 2: Varying intercepts and slopes model (no correlation) }}^{\text {n }}$

## Visualizing random effects

\#\# \$subject


Formal statement of varying intercepts and varying slopes linear mixed model
i indexes subjects, j items.

$$
\begin{equation*}
y_{i j}=\beta_{0}+u_{0 i}+\left(\beta_{1}+u_{1 i}\right) \times s o_{i j}+\epsilon_{i j} \tag{4}
\end{equation*}
$$

Variance components:

- $u_{0} \sim \operatorname{Normal}\left(0, \sigma_{u 0}\right)$
- $u_{1} \sim \operatorname{Normal}\left(0, \sigma_{u 1}\right)$
- $\epsilon \sim \operatorname{Normal}(0, \sigma)$


## Shrinkage in linear mixed models

- The estimate of the effect by participant is smaller than when we fit a separate linear model to the subject's data.
- This is called shrinkage in linear mixed models: the individual level estimates are shunk towards the mean slope.
- The less data we have from a given subject, the more the shrinkage.


## Shrinkage in linear mixed models



Subject 36's estimates


Subject 37's estimates


## Shrinkage in linear mixed models

The effect of missing data on estimation in LMMs
Let's randomly delete some data from one subject:

```
set.seed(4321)
## choose some data randomly to remove:
rand<-rbinom(1,n=16,prob=0.5)
```


## Shrinkage in linear mixed models

The effect of missing data on estimation in LMMs

```
dat[which(dat$subject==37),]$rawRT
## [1] 770 536 686 578 457 487 2419 884 3365 233
## [15] 1081 971
dat$deletedRT<-dat$rawRT
dat[which(dat$subject==37),]$deletedRT<-
    ifelse(rand,NA,
    dat[which(dat$subject==37),]$rawRT)
```


## Shrinkage in linear mixed models

The effect of missing data on estimation in LMMs
Now subject 37 's estimates are going to be pretty wild:
subset(dat,subject==37)\$deletedRT
\#\# [1] 770 NA 686578 NA NA NA NA 3365233
\#\# [15] NA 971

## Shrinkage in linear mixed models

The effect of missing data on estimation in LMMs

```
## original no pooling estimate:
lmList.fm1_old<-lmList(log(rawRT)~solsubject,dat)
coefs_old<-coef(lmList.fm1_old)
intercepts_old<-coefs_old[1]
colnames(intercepts_old)<-"intercept"
slopes_old<-coefs_old[2]
## subject 37's original estimates:
intercepts_old$intercept[37]
## [1] 6.617
slopes_old$so[37]
## [1] 0.35537
```


## Shrinkage in linear mixed models

The effect of missing data on estimation in LMMs

```
## on deleted data:
lmList.fm1_deleted<-lmList(log(deletedRT) ~solsubject,dat)
coefs<-coef(lmList.fm1_deleted)
intercepts<-coefs[1]
colnames(intercepts)<-"intercept"
slopes<-coefs[2]
## subject 37's new estimates on deleted data:
intercepts$intercept [37]
## [1] 6.6879
slopes$so[37]
```

\#\# [1] 0.38843

## Shrinkage in linear mixed models

The effect of missing data on estimation in LMMs

Subject 37's estimates


SR
OR
$38 / 49$

## Shrinkage in linear mixed models

The effect of missing data on estimation in LMMs

- What we see here is that the estimates from the hierarchical model are barely affected by the missingness, but the estimates from the no-pooling model are heavily affected.
- This means that linear mixed models will give you more robust estimates (think Type $M$ error!) compared to no pooling models.
- This is one reason why linear mixed models are such a big deal.


## Crossed subjects and items in LMMs

Subjects and items are fully crossed:
head(xtabs(~subject+item,dat))

| \#\# | item |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| \#\# subject | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 1 | 12 | 13 | 14 | 15 | 16 |  |
| \#\# | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| \#\# | 2 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| \#\# | 3 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| \#\# | 4 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| \#\# | 5 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| \#\# | 6 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |

## Linear mixed models

Linear mixed model with crossed subject and items random effects.
m2.lmer<-lmer (logrt~so $+(1+$ sol| subject $)+(1+$ sol|item $)$, dat $)$

## Linear mixed models

Random effects:

| Groups | Name | Variance | Std.Dev. |
| :--- | :--- | :--- | :--- |
| subject | (Intercept) | 0.10090 | 0.3177 |
| subject.1 | so | 0.01224 | 0.1106 |
| item | (Intercept) | 0.00127 | 0.0356 |
| item.1 | so | 0.00162 | 0.0402 |
| Residual |  | 0.13063 | 0.3614 |

Number of obs: 672, groups: subject, 42 ; item, 16

Fixed effects:
Estimate Std. Error t value

| (Intercept) | 5.8831 | 0.0517 | 113.72 |
| :--- | ---: | ---: | ---: |
| so | 0.0620 | 0.0242 | 2.56 |

## Lecture 6

Linear mixed models

- Varying intercepts and slopes model, with crossed random effects for subjects and for items


## Visualizing random effects

## subject



## Lecture 6

Linear mixed models

- Varying intercepts and slopes model, with crossed random effects for subjects and for items


## Visualizing random effects

item


## Linear mixed models

Linear mixed model with crossed subject and items random effects.

```
m3.lmer<-lmer(logrt~so+(1+so|subject)+(1+so|item),
    dat)
## boundary (singular) fit: see ?isSingular
```


## Linear mixed models

Linear mixed model with crossed subject and items random effects.
Random effects:

| Groups | Name | Variance | Std.Dev. Corr |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| subject | (Intercept) | 0.10103 | 0.3178 |  |
|  | so | 0.01228 | 0.1108 | 0.58 |
| item | (Intercept) | 0.00172 | 0.0415 |  |
|  | so | 0.00196 | 0.0443 | 1.00 <= degenerate |
| Residual | 0.12984 | 0.3603 |  |  |

Number of obs: 672, groups: subject, 42; item, 16

Fixed effects:

|  | Estimate Std. Error | t value |  |
| :--- | ---: | ---: | ---: |
| (Intercept) | 5.8831 | 0.0520 | 113.09 |
| so | 0.0620 | 0.0247 | 2.51 |

Formal statement of varying intercepts and varying slopes linear mixed model with correlation
i indexes subjects, j items.

$$
\begin{equation*}
y_{i j}=\alpha+u_{0 i}+w_{0 j}+\left(\beta+u_{1 i}+w_{1 j}\right) * s o_{i j}+\varepsilon_{i j} \tag{5}
\end{equation*}
$$

where $\varepsilon_{i j} \sim \operatorname{Normal}(0, \sigma)$ and

$$
\begin{gather*}
\Sigma_{u}=\left(\begin{array}{cc}
\sigma_{u 0}^{2} & \rho_{u} \sigma_{u 0} \sigma_{u 1} \\
\rho_{u} \sigma_{u 0} \sigma_{u 1} & \sigma_{u 1}^{2}
\end{array}\right) \quad \Sigma_{w}=\left(\begin{array}{cc}
\sigma_{w 0}^{2} & \rho_{w} \sigma_{w 0} \sigma_{w 1} \\
\rho_{w} \sigma_{w 0} \sigma_{w 1} & \sigma_{w 1}^{2}
\end{array}\right)  \tag{6}\\
\binom{u_{0}}{u_{1}} \sim \mathcal{N}\left(\binom{0}{0}, \Sigma_{u}\right), \quad\binom{w_{0}}{w_{1}} \sim \mathcal{N}\left(\binom{0}{0}, \Sigma_{w}\right) \tag{7}
\end{gather*}
$$

$L_{\text {Linear mixed models }}$
-Model type 3: Varying intercepts and varying slopes, with correlation

## Visualizing random effects

subject


## Lecture 6

Linear mixed models
-Model type 3: Varying intercepts and varying slopes, with correlation

## Visualizing random effects

These are degenerate estimates


