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HT's

t-Distributions

Shaft  
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A Sample Size  
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df and  
Satterthwaite

Data Desk  
View of  
Creativity

# Math 1710 Class 32

t - Inference incl. 2-Sample and MP  
Dr. Back

Nov. 11, 2009

# Tomorrow's Lab

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Because of the calendar (prelim next Th followed by Thanksgiving the following Th), tomorrow's lab will be the next to last one you hand in.

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We'll try and use today's class to give you an introduction to all the concepts coming up in the lab.

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Consequently tomorrow's lab includes elements of chapters 23, 24, and 25.

We'll try and use today's class to give you an introduction to all the concepts coming up in the lab.

In some cases we'll skip some details and come back to them over the next few classes.

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A coffee vending machine dispenses coffee into a paper cup. You're supposed to get 10 ounces of coffee., but the amount varies slightly form cup to cup. Here are the amounts measured in a random sample of 20 cups.

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9.9	9.7	10.0	10.1
9.9	9.6	9.8	9.8
10.0	9.5	9.7	10.1
9.9	9.6	10.2	9.8
10.0	9.9	9.5	9.9



# Coffee Machine

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Is there evidence that the machine is shortchanging customers?

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Is there evidence that the machine is shortchanging customers?

A natural HT situation.

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Is there evidence that the machine is shortchanging customers?

We'll summarize the data by its mean  $\bar{x} = 9.845$  and its standard deviation  $s = .1986$ .

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**Notation:** Let  $\mu$  denote the mean amount of coffee in a dispensed cup.

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**Notation:** Let  $\mu$  denote the mean amount of coffee in a dispensed cup.

**Hypotheses:**

- $H_0: \mu = 10$  (or  $\mu \geq 10$ )
- $H_a: \mu < 10$

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Is there evidence that the machine is shortchanging customers?

Recall by the CLT that the sampling distribution of  $\bar{x}$  is

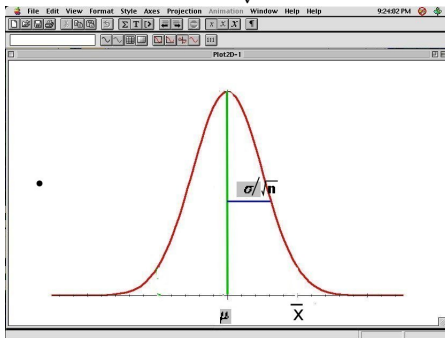
$$N\left(\mu, \frac{\sigma}{\sqrt{n}}\right)$$

when  $n$  is large.

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Is there evidence that the machine is shortchanging customers?

$$N\left(\mu, \frac{\sigma}{\sqrt{n}}\right)$$



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Is there evidence that the machine is shortchanging customers?

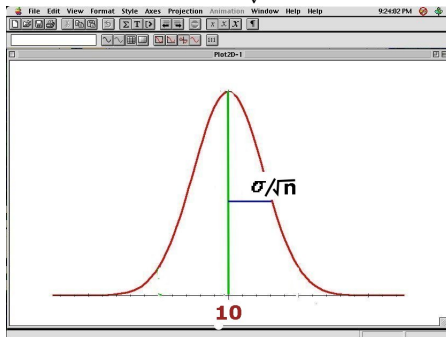
As usual with HT's, we are interested in whether the observed statistic of  $\bar{x} = 9.845$  is reasonably consistent with the sampling distribution assuming  $H_0$  is true.



# Coffee Machine

Is there evidence that the machine is shortchanging customers?

$$N(10, \frac{\sigma}{\sqrt{n}})$$



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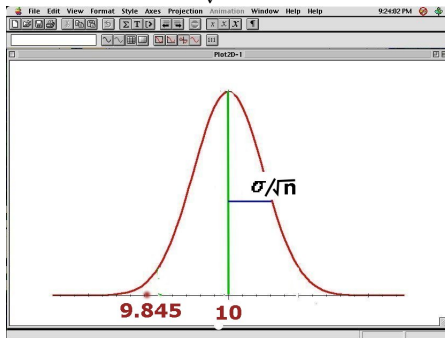
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# Coffee Machine

Is there evidence that the machine is shortchanging customers?

$$N(10, \frac{\sigma}{\sqrt{n}}) \text{ with } \bar{x}$$



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Is there evidence that the machine is shortchanging customers?

We know  $n = 20$ , but the major catch is not knowing  $\sigma$ .

What is the obvious approximation?

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Answer: Use  $s = .1986$  instead of  $\sigma$ .

# Coffee Machine

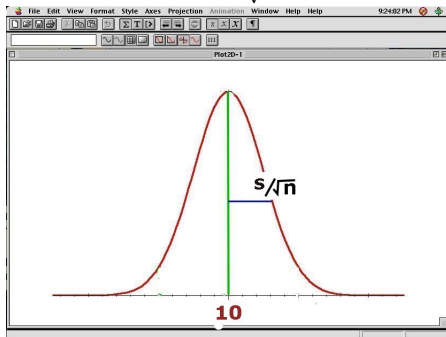
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We know  $n = 20$ , but the major catch is not knowing  $\sigma$ .

What is the obvious approximation?

Answer: Use  $s = .1986$  instead of  $\sigma$ .

$$N\left(10, \frac{s}{\sqrt{n}}\right)$$



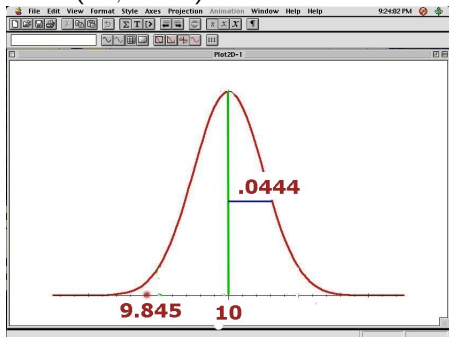
# Coffee Machine

Is there evidence that the machine is shortchanging customers?

Since the standard error is

$$SE(\bar{x}) = \frac{s}{\sqrt{n}} = \frac{.1986}{\sqrt{20}} = .0444$$

$N(10, .0444)$  with  $\bar{x} = 9.845$ .



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Is there evidence that the machine is shortchanging customers?

If  $s = \sigma$ , we'd look at a Z-statistic

$$Z = \frac{\bar{x} - \mu_0}{\frac{\sigma}{\sqrt{n}}} = \frac{9.845 - 10}{\frac{.1986}{\sqrt{20}}}$$

where we've written  $H_0$  more abstractly as  $\mu = \mu_0$ ,  $\mu_0$  being the hypothesized value, 10 in this case.

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Is there evidence that the machine is shortchanging customers?

Because  $s$  will not exactly match  $\sigma$ , we actually get a bit of extra error here. This is compensated for by viewing

$$t = \frac{\bar{x} - \mu_0}{\frac{s}{\sqrt{n}}} = \frac{9.845 - 10}{\frac{.1986}{\sqrt{20}}} = \frac{-.155}{.0444} = -3.49.$$

as a *t-Statistic*.



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as a *t-Statistic*.

Since the error in approximating  $\sigma$  by  $s$  varies with the sample size, there is a different t-distribution for each sample size.

These are labeled by the “degrees of freedom” which for a 1-sample t-test is:

$$df = n - 1.$$

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Is there evidence that the machine is shortchanging customers?

One tail probability		0.10	0.05	0.025	0.01	0.005
Table T Values of $t_\alpha$	df					
	1	3.078	6.314	12.706	31.821	63.657
	2	1.886	2.920	4.303	6.965	9.925
	3	1.638	2.353	3.182	4.541	5.841
	4	1.533	2.132	2.776	3.747	4.604
	5	1.476	2.015	2.571	3.362	4.045
	6	1.440	1.943	2.449	3.143	3.707
	7	1.415	1.895	2.365	2.998	3.501
	8	1.393	1.858	2.306	2.898	3.358
	9	1.383	1.833	2.262	2.821	3.251
10	1.375	1.812	2.228	2.750	3.177	
11	1.368	1.794	2.199	2.693	3.119	
12	1.362	1.779	2.174	2.648	3.071	
13	1.357	1.766	2.152	2.604	3.028	
14	1.353	1.754	2.132	2.562	2.989	
15	1.349	1.743	2.114	2.522	2.953	
16	1.346	1.734	2.098	2.484	2.919	
17	1.343	1.726	2.084	2.448	2.887	
18	1.340	1.719	2.071	2.414	2.857	
19	1.328	1.729	2.093	2.539	2.861	
20	1.325	1.725	2.086	2.528	2.845	
21	1.323	1.721	2.080	2.518	2.831	
22	1.321	1.718	2.075	2.510	2.819	

These are all critical values  $t^*$ .

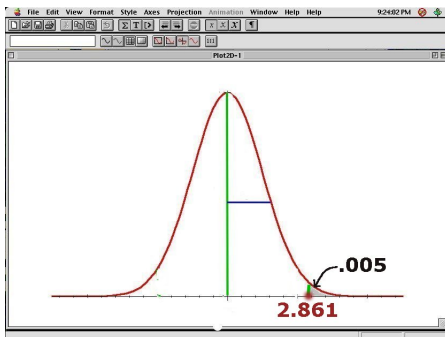
For example  $P(t > 1.328) = .10$  for the t distribution with 19 df.

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Is there evidence that the machine is shortchanging customers?

Our t-statistic of  $-3.49$  is more extreme than any on the  $df=19$  row of the table.

The picture



shows what the critical value  $t^* = 2.861$  for a tail prob. of  $.005$

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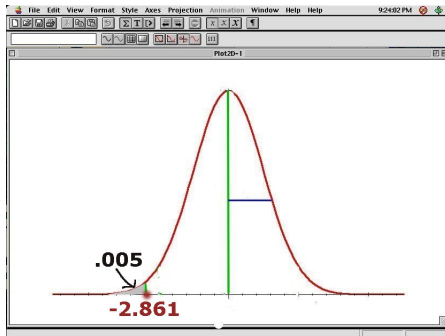
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Is there evidence that the machine is shortchanging customers?

So by symmetry

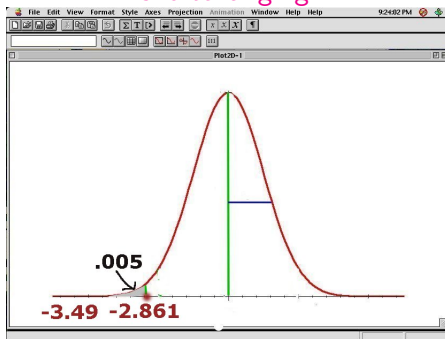


$P(T < -2.861) = .005$  as well.

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Is there evidence that the machine is shortchanging customers?

So our tail probability and p-value are both less than  $.005$  and we reject the null. The machine does appear to be shortchanging.



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Sps. our t-statistic had been 2.00 with the same 1-sided hypotheses

- $H_0: \mu = 10$  (or  $\mu \geq 10$ )
- $H_a: \mu < 10$

What P-value would we report?

# Coffee Machine

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What P-value would we report?

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**Answer:** A tail probability and P-value of between .025 and .05.

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Sps. instead our t-statistic had been 2.00 with 2-sided hypotheses

- $H_0: \mu = 10$
- $H_a: \mu \neq 10$

What P-value would we report?



# Coffee Machine

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**Answer:** Our tail probability is still between .025 and .05 but our P-value is now between .05 and .10.

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A t-hypothesis test with

$$H_0 : \mu = m\mu_0$$

is resolved using a t-statistic of

$$t = \frac{\bar{X} - \mu_0}{\frac{s}{\sqrt{n}}}.$$

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$H_a$  can be any of the three

1  $\mu \neq \mu_0$

2  $\mu > \mu_0$

3  $\mu < \mu_0$

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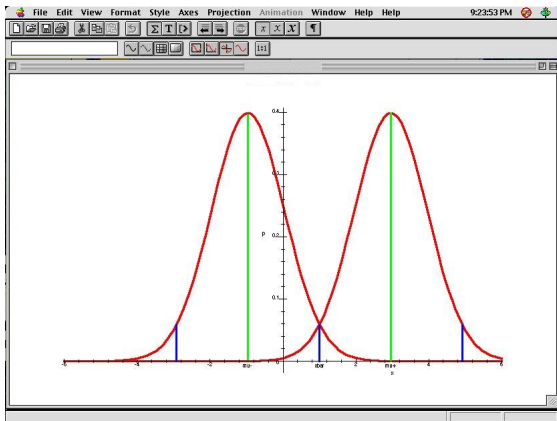
A t confidence interval for  $\mu$  is

$$\bar{x} \pm t^* \frac{s}{\sqrt{n}}$$

for the same reason as the corresponding formula in the proportion case:

# t-CI's and HT's

The two sampling distributions at the border of the CI.  
The blue lines delimit central regions on the sampling distributions.  
The CI is between the green lines on the horizontal axis.



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In the coffee machine example, this gives a 95% CI for  $\mu$  of

$$9.845 \pm 2.093(.0444) = 9.845 \pm .093 = (9.752, 9.938)$$

. You did expect it to not include 10, didn't you?

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Let  $X_0, X_1, \dots, X_d$  be  $d + 1$  independent standard normal random variables.

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Let  $X_0, X_1, \dots, X_d$  be  $d + 1$  independent standard normal random variables.

The t-distribution with  $d$  degrees of freedom is defined to be the random variable

$$\frac{X_0}{\sqrt{X_1^2 + X_2^2 + \dots + X_d^2}}.$$



# t-Distributions

The t-distribution with  $d$  degrees of freedom is defined to be the random variable

$$\frac{X_0}{\sqrt{X_1^2 + X_2^2 + \dots + X_d^2}}.$$

Think of this as keeping track of the t-statistic

$$\frac{\bar{x} - \mu_0}{\frac{s}{\sqrt{n}}} = \frac{\frac{\bar{x} - \mu_0}{\sigma}}{\frac{s}{\sigma}}$$

with the top  $X_0$  keeping track of the numerator and the denominator  $\sqrt{X_1^2 + X_2^2 + \dots + X_d^2}$  keeping track of the ratio of  $s$  to  $\sigma$ .

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Just as the normal distribution has the density formula

$$f(x) = \frac{e^{-\frac{(x-\mu)^2}{2\sigma^2}}}{\sqrt{2\pi\sigma}}$$

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Methods of calculus show that the density formula for the t-distribution with  $d$  degrees of freedom is

$$f(x) = \frac{\Gamma(\frac{d+1}{2})}{\Gamma(\frac{d}{2})} \frac{1}{\sqrt{d\pi}} \frac{1}{(1 + \frac{x^2}{d})^{\frac{d+1}{2}}}$$

where  $d$  is the number of degrees of freedom and

$$\Gamma(n) = \int_0^{\infty} t^{n-1} e^{-t} dt$$

is the *gamma* function, a generalized factorial.

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It will be important when we think about 2-sample tests to realize that the degrees of freedom is just a parameter in the above formula.

$$f(x) = \frac{\Gamma(\frac{d+1}{2})}{\Gamma(\frac{d}{2})} \frac{1}{\sqrt{d\pi}} \frac{1}{(1 + \frac{x^2}{d})^{\frac{d+1}{2}}}$$

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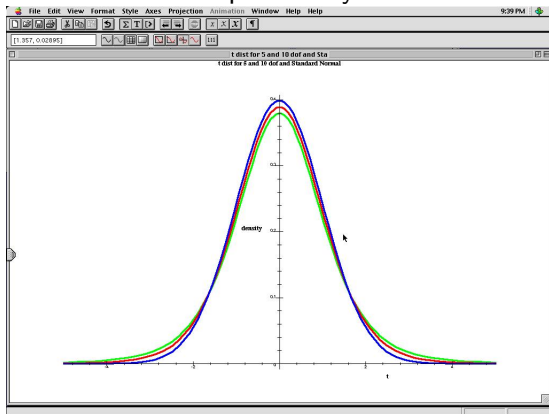
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## A little more probability in the tails



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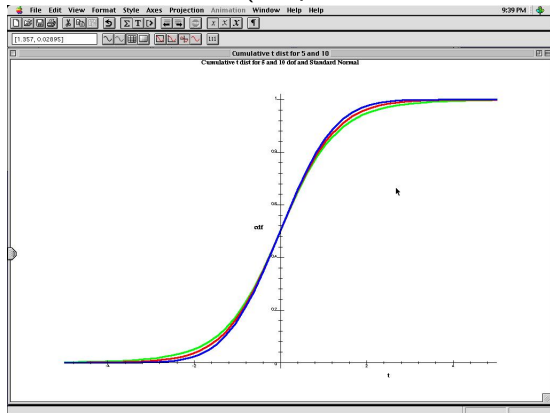
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## Cumulative Dist. Fcn (Graphical Form of Table Z)



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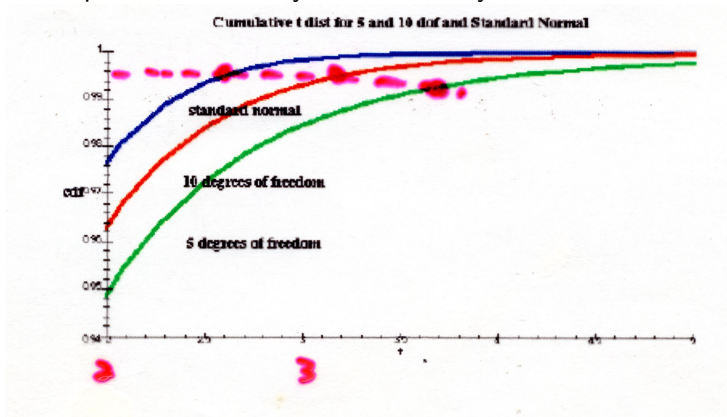
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This picture shows why  $t^*$  can be very different from  $z^*$ .



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t distribution critical value

one tail test tail	0.20	0.10	0.05	0.02	0.01
two tail	0.40	0.20	0.025	0.01	0.005
df					
1	3.078	6.314	12.706	31.821	63.657
2	1.886	2.920	4.303	6.965	9.025
3	1.638	2.353	3.182	4.541	5.841
4	1.533	2.132	2.776	3.747	4.604
5	1.476	2.015	2.571	3.365	4.032
6	1.440	1.943	2.447	3.143	3.707
7	1.415	1.893	2.365	2.998	3.499
8	1.397	1.861	2.306	2.896	3.355
9	1.383	1.833	2.262	2.821	3.250
10	1.372	1.812	2.228	2.764	3.169
11	1.363	1.796	2.201	2.718	3.106
12	1.356	1.782	2.179	2.681	3.055
13	1.350	1.771	2.160	2.650	3.012
14	1.345	1.761	2.145	2.624	2.977
15	1.341	1.753	2.131	2.602	2.949
16	1.337	1.746	2.120	2.583	2.923
17	1.333	1.740	2.110	2.567	2.896
18	1.330	1.734	2.101	2.552	2.879
19	1.328	1.729	2.093	2.539	2.861
20	1.325	1.725	2.086	2.528	2.845
21	1.323	1.721	2.080	2.518	2.831
22	1.321	1.717	2.074	2.509	2.819
23	1.319	1.714	2.069	2.500	2.807
24	1.318	1.711	2.064	2.492	2.797
25	1.316	1.708	2.060	2.485	2.787
26	1.315	1.706	2.056	2.479	2.779
27	1.314	1.703	2.052	2.473	2.771
28	1.313	1.701	2.048	2.467	2.763
29	1.311	1.699	2.043	2.462	2.756
30	1.310	1.697	2.042	2.457	2.750
32	1.309	1.694	2.037	2.448	2.738
35	1.306	1.690	2.030	2.438	2.723
40	1.303	1.684	2.021	2.423	2.704
45	1.301	1.679	2.014	2.412	2.689
50	1.299	1.676	2.009	2.403	2.678
60	1.296	1.671	2.000	2.390	2.660
75	1.293	1.665	1.992	2.377	2.643
100	1.290	1.660	1.984	2.364	2.626
120	1.289	1.658	1.980	2.358	2.617
140	1.288	1.656	1.977	2.353	2.611
160	1.288	1.653	1.973	2.347	2.605
200	1.285	1.649	1.969	2.341	2.596
300	1.284	1.646	1.966	2.336	2.588
400	1.282	1.644	1.962	2.331	2.581
500	1.282	1.643	1.960	2.328	2.575
∞					
90%	0.05	0.10	0.20	0.40	0.50
95%	0.025	0.05	0.10	0.20	0.25
99%	0.005	0.01	0.025	0.05	0.01



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# Haiku Poetry: Intrinsic or Extrinsic Motivation Different in Effectiveness?

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**haiku:** A Japanese poem composed of three unrhymed lines of five, seven, and five syllables. Haiku often reflect on some aspect of nature.

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How can you study this?

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## How can you study this?

Randomized Experiment on Creativity in poetry:

Grades on creativity recorded. (Higher is better.)

Both groups filled out questionnaires before writing poetry.

- Intrinsic group first did questionnaire emphasizing intrinsic motivations.
- Extrinsic group first did questionnaire emphasizing extrinsic motivations.

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How can you study this?

People then wrote Haiku poems which were graded by experienced poets.

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How can you study this?

## The Results

Group	n	$\bar{x}$	s
Intrinsic	24	19.88	4.44
Extrinsic	23	15.73	5.25

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How can you study this?

**Notation:** Let  $\mu_i$  and  $\mu_e$  denote the respective mean haiku scores for the intrinsically and the extrinsically motivated groups.



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**Notation:** Let  $\mu_i$  and  $\mu_e$  denote the respective mean haiku scores for the intrinsically and the extrinsically motivated groups.

**Hypotheses:**

- $H_0: \mu_i = \mu_e$
- $H_a: \mu_i \neq \mu_e$

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How can you study this?

Hypotheses:

- $H_0: \mu_i = \mu_e$
- $H_a: \mu_i \neq \mu_e$

Under  $H_0$ , the sampling distribution of  $\bar{x}_i - \bar{x}_e$  has mean 0 and approx. std deviation

$$SE(\bar{x}_i - \bar{x}_e) = \sqrt{\frac{s_i^2}{n_i} + \frac{s_e^2}{n_e}}$$

# Haiku Poetry: Intrinsic or Extrinsic Motivation Different in Effectiveness?

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$$SE(\bar{x}_i - \bar{x}_e) = \sqrt{\frac{s_i^2}{n_i} + \frac{s_e^2}{n_e}}$$

$$SE(\bar{x}_i - \bar{x}_e) = \sqrt{\frac{4.44^2}{24} + \frac{5.25^2}{23}} = 1.42.$$

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# Haiku Poetry: Intrinsic or Extrinsic Motivation Different in Effectiveness?

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$$SE(\bar{x}_i - \bar{x}_e) = \sqrt{\frac{4.44^2}{24} + \frac{5.25^2}{23}} = 1.42.$$

Our t-statistic is

$$t = \frac{19.88 - 15.73}{1.42} = 2.92.$$

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Our t-statistic is

$$t = \frac{19.88 - 15.73}{1.42} = 2.92.$$

For homework and hand computation on exams, we suggest you use df the smaller of 24-1 and 23-1 to interpret the t-statistic. *To be explained later.*

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How can you study this?

Hypotheses:

- $H_0: \mu_i = \mu_e$
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With  $df = 22$ , 2.92 is more extreme than the biggest  $t^*$  in table T (which is 2.819) so our tail probability is less than .005 and our P-value is less than .01.

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We reject  $H_0$ . Intrinsic vs. extrinsic motivation does seem to make a difference.

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How can you study this?

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- $H_0: \mu_i = \mu_e$

- $H_a: \mu_i \neq \mu_e$

We reject  $H_0$ . Intrinsic vs. extrinsic motivation does seem to make a difference.

A 95% CI for  $\mu_i \neq \mu_e$  would be

$$19.88 - 15.73 \pm 2.074 \cdot 1.42 = 4.15 \pm 2.95 = (1.20, 7.10).$$



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How can you study this?

Hypotheses:

- $H_0: \mu_i = \mu_e$
- $H_a: \mu_i \neq \mu_e$

Scope of the Inference:

- Causal relationship seems to be established.
- Doesn't say much about a population since not an SRS.

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In the 2-independent sample case, even if the individual data is normal, the difference  $\bar{x} - \bar{y}$  does *not* follow a t-distribution unless  $\sigma_x = \sigma_y$ .

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In the 2-independent sample case, even if the individual data is normal, the difference  $\bar{x} - \bar{y}$  does *not* follow a t-distribution unless  $\sigma_x = \sigma_y$ .

But the sampling distribution of  $\bar{x} - \bar{y}$  can be well approximated by a t-distribution with the following degrees of freedom:

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In the 2-independent sample case, even if the individual data is normal, the difference  $\bar{x} - \bar{y}$  does *not* follow a t-distribution unless  $\sigma_x = \sigma_y$ .

$$df = \frac{\left(\frac{s_x^2}{n_x} + \frac{s_y^2}{n_y}\right)^2}{\frac{1}{n_x-1} \left(\frac{s_x^2}{n_x}\right)^2 + \frac{1}{n_y-1} \left(\frac{s_y^2}{n_y}\right)^2}$$

# Satterthwaite

In the 2-independent sample case, even if the individual data is normal, the difference  $\bar{x} - \bar{y}$  does *not* follow a t-distribution unless  $\sigma_x = \sigma_y$ .

$$df = \frac{\left(\frac{s_x^2}{n_x} + \frac{s_y^2}{n_y}\right)^2}{\frac{1}{n_x-1} \left(\frac{s_x^2}{n_x}\right)^2 + \frac{1}{n_y-1} \left(\frac{s_y^2}{n_y}\right)^2}$$

Let  $\min(n_x - 1, n_y - 1)$  denote the smallest of  $n_x - 1$  and  $n_y - 1$ . This df is always between  $\min(n_x - 1, n_y - 1)$  and  $n_x + n_y - 2$ .

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Special Case  $s_x = s_y = s$ ,  $n_x = n_y = n$ .

$$df = \frac{\left(\frac{2s^2}{n}\right)^2}{\frac{1}{n-1} \frac{2s^4}{n^2}} = 2(n-1).$$

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Special Case  $\frac{s_x^2}{n_x} \ll \frac{s_y^2}{n_y}$

$$df = \frac{\left(\frac{s_y^2}{n_y}\right)^2}{\frac{1}{n_y-1} \left(\frac{s_y^2}{n_y}\right)^2} = n_y - 1.$$

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Calculators and computers will use this df formula. *For hand computation on HW and exams, we suggest you never use this formula. Instead use the conservative  $\min(n_x - 1, n_y - 1)$ .*

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Calculators and computers will use this df formula. *For hand computation on HW and exams, we suggest you never use this formula. Instead use the conservative  $\min(n_x - 1, n_y - 1)$ .* Your  $t^*$  will be a little too big so your CI will be a little too wide and your HT a little harder to show statistical significance.

**But this is not so bad.**

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In your published papers and lab reports, use a computer which will likely automatically use this formula.

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$$df = \frac{\left(\frac{s_x^2}{n_x} + \frac{s_y^2}{n_y}\right)^2}{\frac{1}{n_x-1} \left(\frac{s_x^2}{n_x}\right)^2 + \frac{1}{n_y-1} \left(\frac{s_y^2}{n_y}\right)^2}$$

Satterthwaite derived this formula using the fact that the exact mean and variance of the sampling distribution of  $\bar{x} - \bar{y}$  is easy to calculate. He asked, “What df in the t-distribution would give the right  $\mu$  and  $\sigma$ .” The answer is the above.

# Satterthwaite

In the 2-independent sample case, even if the individual data is normal, the difference  $\bar{x} - \bar{y}$  does *not* follow a t-distribution unless  $\sigma_x = \sigma_y$ .

$$df = \frac{\left(\frac{s_x^2}{n_x} + \frac{s_y^2}{n_y}\right)^2}{\frac{1}{n_x-1} \left(\frac{s_x^2}{n_x}\right)^2 + \frac{1}{n_y-1} \left(\frac{s_y^2}{n_y}\right)^2}$$

$$f(x) = \frac{\Gamma\left(\frac{d+1}{2}\right)}{\Gamma\left(\frac{d}{2}\right)} \frac{1}{\sqrt{d\pi}} \frac{1}{\left(1 + \frac{x^2}{d}\right)^{\frac{d+1}{2}}}$$

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Summary of **INTRINSIC:SCORE**  
No Selector

Percentile 25

<b>Count</b>	24
<b>Mean</b>	19.8833
<b>Median</b>	20.4
<b>StdDev</b>	4.43951
<b>Range</b>	17.7
<b>IntQRange</b>	5.05
<b>Lower ith %tile</b>	17.35
<b>Upper ith %tile</b>	22.4

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Summary of **EXTRINSIC:SCORE**  
No Selector

Percentile 25

<b>Count</b>	23
<b>Mean</b>	15.7391
<b>Median</b>	17.2
<b>StdDev</b>	5.2526
<b>Range</b>	19
<b>IntQRange</b>	7
<b>Lower ith %tile</b>	12.075
<b>Upper ith %tile</b>	19.075

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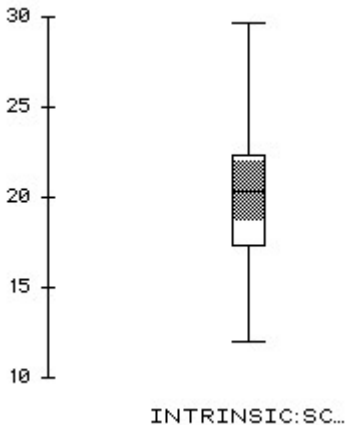
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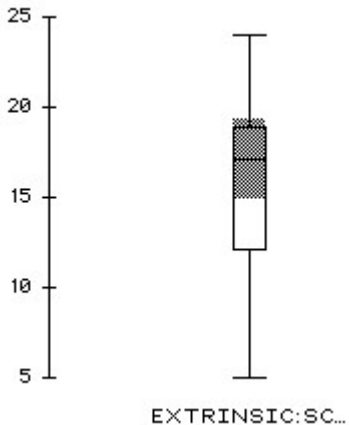
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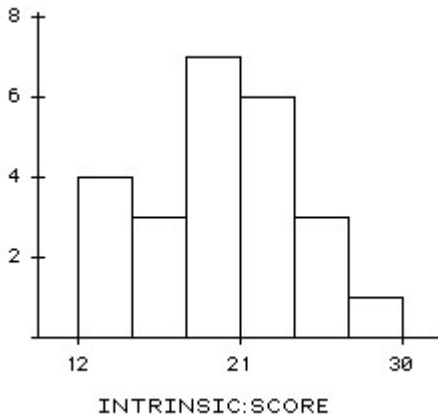
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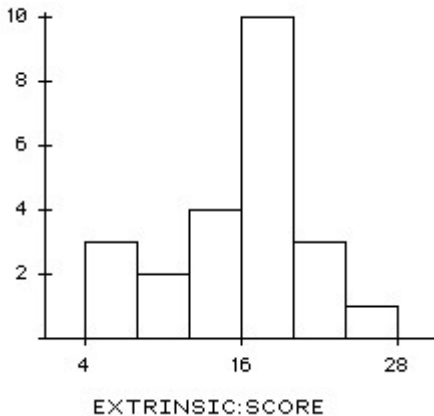
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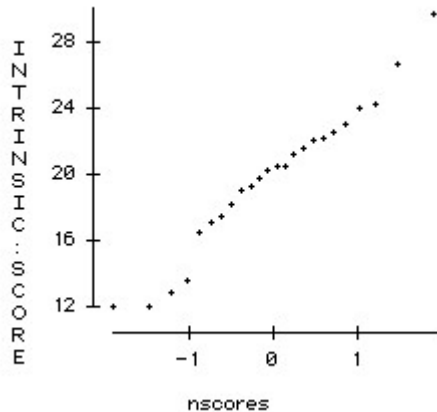
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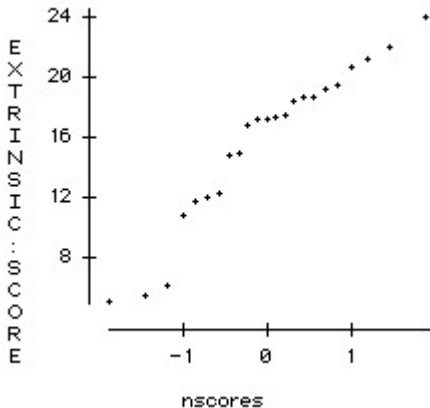
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t-Interval for Individual  $\mu$ 's

No Selector

Individual Confidence 95.00%

Bounds: Lower Bound  $< \mu <$  Upper Bound

With 95.00% Confidence,  $18.008691 < \mu(\text{INTRINSIC:SCORE}) < 21.757975$

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t-Interval for Individual  $\mu$ 's

No Selector

Individual Confidence 95.00%

Bounds: Lower Bound  $< \mu <$  Upper Bound

With 95.00% Confidence, 13.467738  $< \mu$ (EXTRINSIC:SCORE)  $<$  18.010523

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2-Sample t-Interval for  $\mu_1 - \mu_2$

No Selector

Individual Confidence 95.00%

Bounds: Lower Bound  $< \mu_1 - \mu_2 <$  Upper Bound

With 95.00% Confidence,  $-7.0108029 < \mu(\text{EXTRINSIC:SCORE}) - \mu(\text{INTRINSIC:SCORE}) < -1.2776029$

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2-Sample t-Test of  $\mu_1 - \mu_2$

No Selector

Individual Alpha Level 0.05

Ho:  $\mu_1 - \mu_2 = 0$  Ha:  $\mu_1 - \mu_2 \neq 0$

**EXTRINSIC:SCORE - INTRINSIC:SCORE:**

Test Ho:  $\mu(\text{EXTRINSIC:SCORE}) - \mu(\text{INTRINSIC:SCORE}) = 0$  vs Ha:  $\mu(\text{EXTRINSIC:SCORE}) - \mu(\text{INTRINSIC:SCORE}) \neq 0$

Difference Between Means = -4.1442029 t-Statistic = -2.915 w/43 df

Reject Ho at Alpha = 0.05

p = 0.0056



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**Does regular exercise reduce resting pulse rates?** 10 volunteers do 20 minutes of exercise 3 times a week for 6 weeks. Their resting pulse rates before and after were measured.  
(Beats/min.)

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Subject	Before	After
Allen	73	73
Brandon	83	79
Carlos	85	81
David	87	86
Edwin	91	87
Franco	99	91
Graeme	87	84
Hans	85	83
Ivan	83	84
Jorge	79	76

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Allen	73	73
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Carlos	85	81
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Edwin	91	87
Franco	99	91
Graeme	87	84
Hans	85	83
Ivan	83	84
Jorge	79	76

Notice that 8 out of 10 subjects had reduced RPR's. There is a non-parametric test called the sign test that could be used to conclude that RPR's tend to decrease as hypothesized. **Can you see how binomial distribution ideas could be used for this?**

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**Notation:** Let  $\mu_b$  be the mean resting pulse rate before the exercise program and  $\mu_a$  the mean resting pulse rate after.

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**Notation:** Let  $\mu_b$  be the mean resting pulse rate before the exercise program and  $\mu_a$  the mean resting pulse rate after.

**Hypotheses:**

- $H_0: \mu_b = \mu_a$  (or  $\mu_b - \mu_a \leq 0$  or ...)
- $H_a: \mu_b > \mu_a$

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Since individuals have important characteristics besides their exercise programs which determine their RPR's, statistical tests based on

$$\text{Var}(\bar{x}_b - \bar{x}_a) = \text{Var}(\bar{x}_b) + \text{Var}(\bar{x}_a)$$

would not be appropriate here.

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In MP, we look at the differences and apply 1-sample ideas to the differences.

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Differences  $d_i$

Subject	Before - After
Allen	0
Brandon	4
Carlos	4
David	1
Edwin	4
Franco	8
Graeme	3
Hans	2
Ivan	-1
Jorge	3



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The mean  $\bar{d} = 2.8$  with a standard deviation  $s$  of 2.8.

Note you need to directly calculate  $s$  for the differences; it is not determined by  $s_a$  and  $s_b$ .

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The mean  $\bar{d} = 2.8$  with a standard deviation  $s$  of 2.8.

$$t = \frac{2.8}{\frac{2.53}{\sqrt{10}}} = -3.50$$

with  $df=9$ . Our P-value is  $< .005$  and we reject  $H_0$ . **Exercise does seem to reduce RPR's.**

# Matched Pairs vs. Indep. 2-Sample

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<b>Group1</b>	<b>Group 2</b>	
$x_1$	$y_1$	
$x_2$	$y_2$	
$\dots$	$\dots$	
$x_n$	$\dots$	
	$y_m$	
$\bar{x}$	$\bar{y}$	<b>mean</b>
$s_x$	$s_y$	<b>std. dev</b>

# Matched Pairs vs. Indep. 2-Sample

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Group1	Group 2	
$x_1$	$y_1$	
$x_2$	$y_2$	
$\dots$	$\dots$	
$x_n$	$\dots$	
	$y_m$	
$\bar{x}$	$\bar{y}$	<b>mean</b>
$s_x$	$s_y$	<b>std. dev</b>

## Independent 2-Sample

Each  $x_i$  and each  $y_j$  are (except to the extent that they come from one of the groups) independent of each other.

The sample sizes  $m$  and  $n$  can be different.

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Group1	Group 2	
$x_1$	$y_1$	
$x_2$	$y_2$	
$\dots$	$\dots$	
$x_n$	$\dots$	
	$y_m$	
$\bar{x}$	$\bar{y}$	<b>mean</b>
$s_x$	$s_y$	<b>std. dev</b>

## Matched Pairs

The  $i$ 'th observation  $x_i$  in group 1 is more closely related (in a relevant way) to the  $i$ 'th observation  $y_i$  in group 2 than a general  $x_i$  is to a general  $y_j$ .

The sample sizes  $m$  and  $n$  must be the same.