

# Why are Banks Exposed to Monetary Policy?

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## Abstract

We propose a model of banks' exposure to movements in interest rates and their role in the transmission of monetary shocks. Since bank deposits provide liquidity, higher interest rates allow banks to earn larger spreads on deposits. Therefore, if risk aversion is higher than one, banks' optimal dynamic hedging strategy is to take losses when interest rates rise. This risk exposure can be achieved by a traditional maturity-mismatched balance sheet, and amplifies the effects of monetary shocks on the cost of liquidity. The model can match the level, time pattern, and cross-sectional pattern of banks' maturity mismatch.

Keywords: Monetary policy, bank deposits, interest rate risk, maturity mismatch

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# 1 Introduction

Banks typically have maturity-mismatched balance sheets, with long-duration nominal assets (like fixed-rate mortgages) and short-duration nominal liabilities (like deposits). This means that an increase in interest rates leads to a fall in banks' net worth, measured in mark-to-market terms. In this paper we argue that banks choose this exposure deliberately as part of a dynamic hedging strategy. We propose and quantify a model of risk sharing between banks and households and show that it can successfully explain banks' average maturity mismatch, its time-series properties, and the cross-sectional evidence. It provides a laboratory to understand how monetary policy determines banks' risk taking decisions.

Our baseline model is a flexible-price monetary economy where the only source of shocks is monetary policy. The economy is populated by banks and households. The distinguishing feature of banks is that they are able to provide liquidity by issuing deposits that are close substitutes to currency, up to a leverage limit. Importantly, because markets are complete, banks are able to choose their exposure to risk independently of their liquidity provision. In particular, we don't make any assumptions about what kind of securities banks hold. We show that if relative risk aversion is high (larger than one) banks optimally choose to sustain mark-to-market losses when interest rates rise. This exposure to risk can be achieved with a portfolio of long-duration nominal assets and short-duration nominal liabilities, as in a traditional bank balance sheet.

The mechanism works as follows. Because deposits provide liquidity services, banks earn the spread between the nominal interest rate on illiquid bonds and the lower interest rate on deposits. If nominal interest rates rise, the opportunity cost of holding currency goes up, so agents substitute towards deposits. This drives up the equilibrium spread between the nominal interest rate and the interest rate on deposits, increasing banks' return on equity. Because risk aversion is higher than one, banks want to transfer wealth from states of the world with high return on equity to states of the world with low return on equity. They are willing to take capital losses when interest rates rise because spreads going forward will be high. Since the supply of deposits is tied to banks' net worth, the cost of liquidity rises further, amplifying the effects of the monetary shock.

We calibrate the model to match the observed behavior of interest rates, deposit spreads, bank leverage and other macroeconomic variables. We find that an increase in the short-term interest rate of 100 basis points produces losses of around 30% of banks' mark-to-market net worth. These mark-to-market losses need not show up in banks' financial statements, which do not mandate marking to market for assets in the banking book. We show that standard

accounting can disguise *both* the fall in mark-to-market net worth and the increase in spreads that result from a rise in interest rates, so that book values and measured net interest margins will be almost constant. This is consistent with the recent findings by Drechsler et al. (2018), who document that the response of banks' net interest margins to changes in interest rates is approximately zero.

We find that the endogenous response of banks' net worth amplifies the effects of monetary shocks on the cost of liquidity. An increase of 100 *bp* in interest rates has a direct effect on deposit spreads of 62 *bp*, and an additional indirect effect through lower bank net worth and deposit supply of 15 *bp* (an amplification by a factor of 1.25). The effect is non-linear, however. The amplification through banks' net worth is larger when banks' net worth or interest rates are high.

The model can match the level, time pattern, and cross-sectional pattern of banks' maturity mismatch. First, the model can account for the average maturity mismatch. In the model, banks' exposure to movements in interest rates can be implemented with an average maturity mismatch between assets and liabilities of 3.9 years. We compare this to banking data following the approach of English et al. (2018) and find an average maturity mismatch for the median bank of 4.4 years. Second, the model reproduces the time pattern in the data. Banks' maturity mismatch rises during periods of low interest rates. While this may look like "reaching for yield", it is actually what our model predicts. The sensitivity of deposit spreads to interest rate movements is higher when interest rates are low, so banks' dynamic hedging motive is larger at these times. The time-series correlation between the model and the data is 0.77. Third, the model successfully accounts for the cross-sectional evidence. Banks with a higher deposit-to-net-worth ratio should have a stronger dynamic hedging motive and therefore choose a greater maturity mismatch. In the model, increasing the deposit-to-net-worth ratio of a bank by one unit leads to an increase in the maturity mismatch of 0.42 years. In the data, it leads to an increase of 0.43 years.

The baseline model with only monetary shocks is intended as a benchmark to examine the mechanisms at play. We also study an arguably more realistic setting where the central bank follows an inflation targeting policy. The economy is hit by real shocks that move the equilibrium real interest rate and force the central bank to adjust the nominal interest rate in order to hit its inflation target. The quantitative results are similar to the benchmark model, with an average maturity mismatch of 4.7 years, a time-series correlation with the data of 0.51, and an increase of 0.55 years in additional maturity mismatch per unit of deposit-to-net-worth ratio.

There are several alternative explanations for banks' exposure to interest rate risk. First, one could conjecture that a maturity-mismatched balance sheet is inherent to the banking business and the resulting interest rate risk is an inevitable side effect. However, banks can easily adjust the interest rate exposure of their assets without changing their maturity, for example, by using adjustable-rate mortgages. Moreover, there are deep and liquid markets for interest rate derivatives that banks can use to hedge their interest rate risk. In fact, Begenau et al. (2015) show that banks hold positions in these derivatives, but they use them to *amplify* their exposure. Here we assume complete markets. Banks' maturity mismatch, and the resulting interest rate risk, is an endogenous choice.

Second, a traditional view is that maturity mismatch is a way for banks to take advantage of the term premium. But in general equilibrium this explanation is insufficient, because households are also able to take advantage of the term premium, for example, by investing in long-term bonds. Our model does produce a term premium, and both banks and households have incentives to take advantage of it. But in equilibrium the term premium simply reflects the fact that periods with high interest rates are bad for everyone because liquidity is expensive. Banks have a natural hedge against this risk because they earn the deposit spread, so in equilibrium they take interest rate risk and earn the term premium, while households pay the term premium to insure themselves against high interest rates.

Third, interest rate risk could be evidence of risk-seeking behavior, which regulators should be concerned about. Our findings suggest an alternative, more benign, interpretation. Banks are essentially insuring against the underlying risk in their deposit taking business. Our model provides a quantitative benchmark to assess whether banks are engaging in risk-seeking. Large deviations from this benchmark *in either direction* would be indicative of risk-seeking. In particular, if banks did not expose their balance sheet to interest rates at all (for instance by having no maturity mismatch) they would in fact be taking on a large amount of risk due to the sensitivity of deposit spreads to interest rates. Our quantitative results show no evidence of risk-seeking by the aggregate banking sector: the size of banks' exposure to interest rate risk is consistent with a dynamic hedging strategy by highly risk averse agents.<sup>1</sup>

More generally, our theory provides a lens to understand banks' risk exposures beyond interest rate risk. It predicts that banks will choose exposure to risks that are correlated with their investment opportunities. While in this paper we focus on banks' role as providers

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<sup>1</sup>Of course, banks may very well be engaging in risk-seeking behavior on other dimensions. Also, the aggregate evidence does not rule out risk-seeking by individual banks.

of liquidity, banks are also involved in the origination and collection of loans and earn the spread between risky and safe bonds. The same logic implies that they should be willing to take losses when this spread goes up because they expect a higher return on wealth looking forward. In fact, Begenau et al. (2015) report that banks are highly exposed to credit-spread risk: they face large losses when the spread between BBB and safe bonds rises. In contrast, when we add TFP shocks to the model, we find that these are shared proportionally by banks and households. Our model therefore provides a theory not only of how much, but also what type of risk banks take.

Our paper fits into the literature that studies the role of the financial sector in the propagation and amplification of aggregate shocks (Brunnermeier and Sannikov (2014), He and Krishnamurthy (2011), He and Krishnamurthy (2012), Di Tella (2017), Gertler and Kiyotaki (2015)). Relative to this literature, the main innovation in our paper is that we model banks as providers of liquidity through deposits. This allows us to study the role of the banking sector in the transmission of monetary policy. An important question in this literature is why the financial sector is so exposed to certain aggregate shocks. Our approach has in common with Di Tella (2017) that we allow complete markets; the equilibrium allocation of aggregate risk reflects agents' dynamic hedging of investment opportunities. The economics, however, are very different. Explicitly modeling the banking business allows us to understand banks' dynamic hedging incentives, which are different from other financial institutions, and to assess them quantitatively.

An important ingredient of the mechanism is that the equilibrium spread between illiquid bonds and deposits is increasing in the nominal interest rate. We find this stylized fact is borne out by the empirical evidence. In our data, a 100 bp increase in interest rates is associated with a 66 bp increase in the deposit spread (our model produces 62 bp). This has been observed before. Hannan and Berger (1991) and Driscoll and Judson (2013) attribute it to a form of price stickiness; Drechsler et al. (2017) attribute it to imperfect competition among bank branches; Yankov (2018) attributes it to search costs. Nagel (2016) makes a related observation: the premium on other near-money assets (besides banks deposits) also co-moves with interest rates. He attributes this, as we do, to the substitutability between money and other liquid assets. Krishnamurthy and Vissing-Jorgensen (2015) document a negative correlation between the supply of publicly issued liquid assets and the supply of liquid bank liabilities, also consistent with their being substitutes. Begenau and Landvoigt (2016) study substitution between bank deposits and shadow bank liabilities. We choose the simplest possible specification to capture this: substitution between physical currency and

deposits, but this literature suggests that the phenomenon is broader.

There is also a large theoretical literature studying the nature of bank deposits (Diamond and Dybvig (1983), Diamond and Rajan (2001), etc.) and money (Kiyotaki and Wright (1989), Lagos and Wright (2005), etc.). We make no contribution to this literature, and simply assume that currency and deposits are substitutes in the utility function. Relative to this literature, the contribution of our work is to derive the implications for equilibrium risk management in a model where the underlying risk is modeled explicitly. It is worth stressing that there is no necessary link between liquidity provision via maturity transformation and exposure to interest rate risk. A bank could, for example, issue demand deposits backed by illiquid, long-term, *variable rate* loans: maturity transformation without interest rate risk. Interest rate swaps are another way of achieving the same outcome.

Other studies have looked at different measures of banks' interest rate risk exposure. Drechsler et al. (2018) focus on book net interest margins and show that they almost don't respond to interest rates, a finding that our model replicates. Hoffmann et al. (2018) propose also looking at a marked-to-market balance sheet, but where (unlike in our exercise) deposits are also marked to market, taking into account how deposit rates co-move with interest rates. By this measure, the average bank is not exposed to interest rate risk, although there is cross-sectional heterogeneity. English et al. (2018) use high-frequency data around FOMC announcements to study how bank stock prices react to unexpected changes in the level and slope of the yield curve, and find that bank stocks fall after interest rate increases. Paul (2020) extends this analysis by decomposing the slope of the yield curve into an expectations term and a term premium term.

Also relatedly, Rampini et al. (2015) provide an alternative explanation for why banks fail to hedge the exposure to interest rate risk that arises from their traditional business. They argue that collateral-constrained banks are willing to give up hedging to increase investment, and provide empirical evidence showing that banks who suffer financial losses consequently reduce their hedging. Our model explicitly abstracts from these considerations in the sense that all banks are equally constrained and never face a tradeoff between hedging and investment. Gomez et al. (forthcoming) show cross-sectional evidence that exposure to interest rate risk has consequences for bank lending. Haddad and Sraer (2019) propose a measure of banks' exposure to interest rate risk and find that it is positively correlated with the term premium.

## 2 The Model

**Preferences and technology.** Time is continuous. There is a fixed capital stock  $k$  which produces a constant flow of consumption goods  $y_t = ak$ . There are two types of agents: households and bankers, a continuum of each. Both have identical Epstein-Zin preferences with intertemporal elasticity of substitution equal to 1, risk aversion  $\gamma$  and discount rate  $\rho$ :

$$U_t = \mathbb{E}_t \left[ \int_t^\infty f(x_s, U_s) ds \right]$$

with

$$f(x, U) = \rho(1 - \gamma)U \left( \log(x) - \frac{1}{1 - \gamma} \log((1 - \gamma)U) \right)$$

The good  $x$  is a Cobb-Douglas composite of consumption  $c$  and liquidity services from money holdings  $m$ :

$$x(c, m) = c^\beta m^{1-\beta} \tag{1}$$

Money itself is a CES composite of real currency holdings  $h$  (provided by the government) and real bank deposits  $d$ , with elasticity of substitution  $\epsilon$ :<sup>2</sup>

$$m(h, d) = \left( \alpha^{\frac{1}{\epsilon}} h^{\frac{\epsilon-1}{\epsilon}} + (1 - \alpha)^{\frac{1}{\epsilon}} d^{\frac{\epsilon-1}{\epsilon}} \right)^{\frac{\epsilon}{\epsilon-1}} \tag{2}$$

Formulation (2) captures the idea that both currency and deposits are used in transactions, so they both provide liquidity services. Substitution between these types of money will determine the behavior of deposit interest rates.

**Currency and deposits.** The government supplies nominal currency  $H$ , following an exogenous stochastic process

$$\frac{dH_t}{H_t} = \mu_{H,t} dt + \sigma_{H,t} dB_t$$

where  $B$  is a standard Brownian motion. The process  $B$  drives equilibrium dynamics. The government distributes or withdraws currency to and from agents through lump-sum transfers or taxes.

Deposits are issued by bankers. This is in fact the only difference between bankers and households. Deposits pay an equilibrium nominal interest rate  $i^d$  and also enter the utility

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<sup>2</sup>Throughout, uppercase letters denote nominal variables and their corresponding lowercase letter are real variables. Hence  $h \equiv \frac{H}{p}$  and  $d \equiv \frac{D}{p}$  where  $p$  is the price of consumption goods in terms of currency, which we take as the numeraire.

function according to equation (2). The amount of deposits bankers can issue is subject to a leverage limit. A banker whose individual wealth is  $n$  can issue deposits  $d^S$  up to

$$d^S \leq \phi n \tag{3}$$

where  $\phi$  is an exogenous parameter. Constraint (3) may be interpreted as either a regulatory constraint or a level of capitalization required for deposits to actually have the liquidity properties implied by (2). This constraint prevents bankers from issuing an infinite amount of deposits, and makes their balance sheets important for the economy.

**Monetary policy.** The government chooses a path for currency supply  $H$  to implement the following stochastic process for the nominal interest rate  $i$  on short-term, safe but illiquid bonds:

$$di_t = \mu_i(i_t) dt + \sigma_i(i_t) dB_t \tag{4}$$

where the drift  $\mu_i(\cdot)$  and volatility  $\sigma_i(\cdot)$  are functions of  $i$ . Shocks  $B$  are our representation of monetary shocks, and they are the only source of risk in the economy.

There is more than one stochastic process  $H$  that will result in (4). Let

$$\frac{dp_t}{p_t} = \mu_{p,t} dt + \sigma_{p,t} dB_t$$

be the stochastic process for the price level (which is endogenous). We assume that the government implements the unique process  $H$  such that in equilibrium (4) holds and  $\sigma_{p,t} = 0$ . Informally, this means that monetary shocks affect the rate of inflation  $\mu_p$  but the price level moves smoothly.

**Markets.** There are complete markets where bankers and households can trade capital and contingent claims. We denote the real price of capital by  $q$ , the nominal interest rate by  $i$ , the real interest rate by  $r$ , and the price of risk by  $\pi$  (so an asset with exposure  $\sigma$  to the process  $B$  will pay an excess return  $\sigma\pi$ ). All these processes are contingent on the history of shocks  $B$ .

The total real wealth of private agents in the economy includes the value of the capital stock  $qk$ , the real value of outstanding currency  $h$  and the net present value of future government transfers and taxes, which we denote by  $g$ . Total wealth is denoted by  $\omega$ :

$$\omega = qk + h + g$$



Total household wealth is denoted by  $w$  and total bankers' wealth is denoted by  $n$ , so

$$n + w = \omega \tag{5}$$

and we denote by  $z \equiv \frac{n}{\omega}$  the share of the aggregate wealth that is owned by bankers.

## 2.1 Discussion of assumptions

**Risk averse bankers.** The assumption that bankers are agents with preferences deserves some discussion. After all, many banks are publicly held and their shares are owned by diversified outside investors. Without any frictions, Modigliani-Miller implies that banks' exposure to risk is undetermined. Instead, bankers in this model represent bank insiders—management and board members—who control the bank and have an undiversified stake through share ownership or incentive contracts. This type of incentive scheme is widespread. He and Krishnamurthy (2014) report that insiders hold around 20% of banks' equity.<sup>3</sup> Di Tella and Sannikov (2016) show that the combination of a retained equity stake and a leverage constraint implement the optimal dynamic incentive contract when agents have access to hidden savings. Here we use a reduced-form approach and take bankers' lack of diversification as given. We purposefully assume that bankers and households have the same preferences, so the mechanism that governs risk exposure in the model does not depend on differential attitudes towards risk.

**Money and monetary policy.** We model money in a highly stylized way, with a simple “currency and deposits in the utility function” specification. In addition, we assume the market for deposits is perfectly competitive, but bankers are limited in their ability to supply deposits by the leverage constraint. This prevents them from competing away deposit spreads, effectively acting like market power for bankers as a whole. Our objective is not to develop a theory of money nor to account for all features of deposit contracts or the deposit market, but rather to write down the simplest framework where banks provide liquidity and deposit spreads increase with the nominal interest rate.

In this model there is no real reason for monetary policy to do anything other than follow the Friedman rule.<sup>4</sup> The choice to model random monetary policy as the only source of risk

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<sup>3</sup>Notice, however, that the size of insider's equity stake is not itself important, because the problem is linear. What matters is that bank profits are a large, undiversified share of insiders' wealth, not that insiders keep a large fraction of those profits.

<sup>4</sup>The CES formulation (2) implies that currency demand is unbounded at  $i = 0$  but the Friedman rule is

in the economy is obviously not driven by realism but by theoretical clarity. In Section 7 we instead look at a variant of the model where monetary policy follows an inflation targeting rule and only responds to real shocks that affect the equilibrium real interest rate, and show that our results also hold in this more realistic monetary policy regime.

**Complete markets.** The assumption of complete markets is theoretically important. We want to avoid mechanically assuming the result that banks are exposed to interest rate risk.<sup>5</sup> In our model, banks are perfectly able to issue deposits without any exposure to interest rate risk, for example by investing only in short term or adjustable-rate assets, or by trading interest rate swaps. More generally, banks are completely free to take any risk exposure, independently of their deposit supply. Relatedly, while we specify deposit contracts in nominal terms, this is without loss of generality because banks could trade inflation swaps. One possible concern is that in practice households may not be able to trade interest rate swaps or other derivatives that allow them to share interest rate risk with bankers. However, households can share interest rate risk with bankers by adjusting the maturity of their assets and liabilities, or using adjustable rate debt.

We also don't make any assumptions on the kind of assets banks hold: both banks and households can hold capital. In our model banks are not particularly good at holding long term fixed rate nominal loans, or any other security. Finally, with complete markets it is not necessary to specify who receives government transfers when the supply of currency changes: all those transfers are priced in and included in the definition of wealth. Notice also that while banks can go bankrupt (if their net worth reaches zero), this never happens in equilibrium. Continuous trading allows them to scale down as their net worth falls and always avoid bankruptcy.

### 3 Equilibrium

**Households' problem.** Starting with some initial nominal wealth  $W_0$ , each household solves a standard portfolio problem:

$$\max_{W,x,c,h,d,\sigma_W} U(x)$$

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optimal in a limiting sense.

<sup>5</sup>Even though markets are complete, there is no claim that the competitive allocation is efficient. Bankers' ability to produce deposits is limited by their wealth, which involves prices. A social planner would want to manipulate these prices to relax the constraint.

subject to the budget constraint:

$$\frac{dW_t}{W_t} = \left( i_t + \sigma_{W,t}\pi_t - \hat{c}_t - \hat{h}_t i_t - \hat{d}_t (i_t - i_t^d) \right) dt + \sigma_{W,t} dB_t \quad (6)$$

$$W_t \geq 0$$

and equations (1) and (2). A hat denotes the variable is normalized by wealth, i.e.  $\hat{c} = \frac{pc}{W} = \frac{c}{w}$ . The household obtains a nominal return  $i_t$  on its wealth. It incurs an opportunity cost  $i_t$  on its holdings of currency. It also incurs an opportunity cost  $(i_t - i_t^d)$  on its holdings of deposits. Let  $s_t = i_t - i_t^d$  denote the spread between the deposit rate and the market interest rate. Furthermore, the household chooses its exposure  $\sigma_W$  to the monetary shock and obtains the risk premium  $\pi\sigma_W$  in return.

Constraint (6) can be rewritten in real terms as

$$\frac{dw_t}{w_t} = \left( r_t + \sigma_{w,t}\pi_t - \hat{c}_t - \hat{h}_t i_t - \hat{d}_t s_t \right) dt + \sigma_{w,t} dB_t \quad (7)$$

where  $r_t = i_t - \mu_{p,t}$  is the real interest rate.

**Bankers' problem.** Bankers are like households, except that they can issue deposits (denoted  $d^S$ ) up to the leverage limit and earn the spread  $s_t$  on these. The banker's problem, expressed in real terms, is:

$$\max_{n,x,c,h,d,d^S,\sigma_n} U(x)$$

subject to:

$$\frac{dn_t}{n_t} = \left( r_t + \sigma_{n,t}\pi_t - \hat{c}_t - \hat{h}_t i_t + \left( \hat{d}_t^S - \hat{d}_t \right) s_t \right) dt + \sigma_{n,t} dB_t \quad (8)$$

$$\hat{d}_t^S \leq \phi$$

$$n_t \geq 0$$

and equations (1) and (2).

**Equilibrium definition** Given an initial distribution of wealth between households and bankers  $z_0$  and an interest rate process  $i$ , a competitive equilibrium is

1. a process for the supply of currency  $H$
2. processes for prices  $p, i^d, q, g, r, \pi$
3. a plan for the household:  $w, x^h, c^h, m^h, h^h, d^h, \sigma_w$

4. a plan for the banker:  $n, x^b, c^b, m^b, h^b, d^b, d^S, \sigma_n$

such that

1. Households and bankers optimize taking prices as given and  $w_0 = (1 - z_0)(q_0k + h_0 + g_0)$  and  $n_0 = z_0(q_0k + h_0 + g_0)$
2. The goods, deposit and currency markets clear:

$$\begin{aligned} c_t^h + c_t^b &= ak \\ d_t^h + d_t^b &= d_t^S \\ h_t^h + h_t^b &= h_t \end{aligned}$$

3. Wealth holdings add up to total wealth:

$$w_t + n_t = q_tk + h_t + g_t$$

4. Capital and government transfers are priced by arbitrage:

$$q_t = \mathbb{E}_t^Q \left[ a \int_t^\infty \exp \left( - \int_t^s r_u du \right) ds \right] \quad (9)$$

$$g_t = \mathbb{E}_t^Q \left[ \int_t^\infty \exp \left( - \int_t^s r_u du \right) \frac{dH_s}{p_s} \right] \quad (10)$$

where  $Q$  is the equivalent martingale measure implied by  $r$  and  $\pi$ .

5. Monetary policy is consistent

$$\begin{aligned} i_t &= r_t + \mu_{p,t} \\ \sigma_{p,t} &= 0 \end{aligned}$$

**Aggregate state variables.** We look for a recursive equilibrium in terms of two state variables: the interest rate  $i$  (exogenous), and bankers' share of aggregate wealth  $z$  (endogenous) which is important because it affects bankers' ability to issue deposits and provide liquidity. Using the definition of  $z = \frac{n}{n+w}$ , we obtain a law of motion for  $z$  from Ito's lemma

and the budget constraints:

$$\begin{aligned} \frac{dz_t}{z_t} = & \underbrace{\left( (1 - z_t) \left( (\sigma_{n,t} - \sigma_{w,t})\pi_t + \phi s_t - (\hat{x}_t^b - \hat{x}_t^h)\chi_t + \sigma_{w,t}(\sigma_{w,t} - \sigma_{n,t}) \right) - \frac{z_t}{1 - z_t} \sigma_{z,t}^2 \right)}_{\equiv \mu_{z,t}} dt \\ & + \underbrace{(1 - z_t) (\sigma_{n,t} - \sigma_{w,t})}_{\equiv \sigma_{z,t}} dB_t \end{aligned} \quad (11)$$

where  $\chi$  is the minimized cost of the bundle good  $x$ , defined by equation (13) below. The law of motion of  $i$  is given by (4). All other equilibrium objects will be functions of  $i$  and  $z$ .

**Static Decisions and Hamilton-Jacobi-Bellman equations.** We study the banker's problem first. It can be separated into a static problem (choosing  $c$ ,  $m$ ,  $h$  and  $d$  given  $x$ ) and a dynamic problem (choosing  $x$  and  $\sigma_n$ ).

Consider the static problem first. Given the form of the aggregators (1) and (2), we immediately get that the minimized cost of one unit of money  $m$  is given by  $\iota$ :

$$\iota(i, s) = (\alpha i^{1-\epsilon} + (1 - \alpha) s^{1-\epsilon})^{\frac{1}{1-\epsilon}} \quad (12)$$

the minimized cost of one unit of the good  $x$  is given by  $\chi$ :

$$\chi(i, s) = \beta^{-\beta} \left( \frac{\iota(i, s)}{1 - \beta} \right)^{1-\beta} \quad (13)$$

and the static choices of  $c$ ,  $m$ ,  $h$  and  $d$  are given by:

$$\frac{c}{x} = \beta \chi \quad (14)$$

$$\frac{m}{x} = (1 - \beta) \frac{\chi}{\iota} \quad (15)$$

$$\frac{h}{m} = \alpha \left( \frac{\iota}{i} \right)^\epsilon \quad (16)$$

$$\frac{d}{m} = (1 - \alpha) \left( \frac{\iota}{s} \right)^\epsilon \quad (17)$$

Turn now to the dynamic problem. In equilibrium it will be the case that  $i^d < i$  so

bankers' leverage constraint will always bind. This means that (8) reduces to

$$\frac{dn_t}{n_t} = \underbrace{(r_t + \sigma_{n,t}\pi_t - \chi(i_t, s_t)\hat{x}_t + \phi s_t)}_{\equiv \mu_{n,t}} dt + \sigma_{n,t}dB_t \quad (18)$$

Given the homotheticity of preferences and the linearity of budget constraints the problem of the banker has a value function of the form:

$$V_t^b(n) = \frac{(\xi_t n)^{1-\gamma}}{1-\gamma}$$

$\xi_t$  captures the value of the banker's investment opportunities, i.e. his ability to convert units of wealth into units of lifetime utility, and follows the law of motion

$$\frac{d\xi_t}{\xi_t} = \mu_{\xi,t}dt + \sigma_{\xi,t}dB_t$$

where  $\mu_{\xi,t}$  and  $\sigma_{\xi,t}$  are equilibrium objects.

The associated Hamilton-Jacobi-Bellman equation is

$$0 = \max_{x, \sigma_n, \mu_n} f(x, V_t^b) + \mathbb{E}_t [dV_t^b]$$

Using Ito's lemma and simplifying, we obtain:

$$\begin{aligned} 0 &= \max_{\hat{x}, \sigma_n, \mu_n} \rho(1-\gamma) \frac{(\xi_t n_t)^{1-\gamma}}{1-\gamma} \left[ \log(\hat{x}n_t) - \frac{1}{1-\gamma} \log((\xi_t n_t)^{1-\gamma}) \right] \\ &\quad + \xi_t^{1-\gamma} n_t^{1-\gamma} \left( \mu_n + \mu_{\xi_t} - \frac{\gamma}{2} \sigma_n^2 - \frac{\gamma}{2} \sigma_{\xi_t}^2 + (1-\gamma) \sigma_{\xi_t} \sigma_n \right) \\ \text{s.t.} \quad &\mu_n = r_t + \sigma_n \pi_t + \phi s_t - \hat{x} \chi_t \end{aligned}$$

The household's problem is similar. The only difference is that the term  $\phi s_t$  is absent from the budget constraint. The value function has the form

$$V_t^h(w) = \frac{(\zeta_t w)^{1-\gamma}}{1-\gamma}$$

where

$$\frac{d\zeta_t}{\zeta_t} = \mu_{\zeta,t}dt + \sigma_{\zeta,t}dB_t$$

and the HJB equation is

$$\begin{aligned}
0 = & \max_{\hat{x}, \sigma_w, \mu_w} \rho(1 - \gamma) \frac{(\zeta_t w_t)^{1-\gamma}}{1 - \gamma} \left[ \log(\hat{x} w_t) - \frac{1}{1 - \gamma} \log((\zeta_t w_t)^{1-\gamma}) \right] \\
& + \zeta_t^{1-\gamma} w_t^{1-\gamma} \left( \mu_w + \mu_{\zeta_t} - \frac{\gamma}{2} \sigma_w^2 - \frac{\gamma}{2} \sigma_{\zeta_t}^2 + (1 - \gamma) \sigma_{\zeta_t} \sigma_w \right) \\
s.t. & \quad \mu_w = r_t + \sigma_w \pi_t - \hat{x} \chi_t
\end{aligned}$$

**Total wealth, spreads and currency holdings.** The first order conditions for  $\hat{x}$  in the banker and household problem are both given by:

$$\hat{x}_t = \frac{\rho}{\chi_t} \quad (19)$$

Since the intertemporal elasticity of substitution is 1, both bankers and households spend their wealth at a constant rate  $\rho$  independent of prices.

Using (19) and the goods market clearing condition we can solve for total wealth:

$$\omega = \frac{ak}{\beta\rho} \quad (20)$$

Hence in this economy total wealth will be constant. This follows because the Cobb-Douglas form of the  $x$  aggregator implies that consumption is a constant share of spending (the rest is liquidity services), the rate of spending out of wealth is constant and total consumption is constant and equal to  $ak$ .

Using (15) and (17), the fact that deposit supply is  $\phi n$  and (19), the deposit market clearing condition can be written as:

$$\rho(1 - \alpha)(1 - \beta)\iota(i, s)^{\epsilon-1} s^{-\epsilon} = \phi z \quad (21)$$

Solving (21) for  $s$  implicitly defines bank spreads  $s(i, z)$  as a function of  $i$  and  $z$ . It's easy to show from (21) that the spread is increasing in  $i$  as long as  $\epsilon > 1$ . If currency and deposits are close substitutes, an increase in  $i$ , which increases the opportunity cost of holding currency, increases the demand for deposits, so the spread must rise to clear the deposit market. Likewise, (21) implies that the spread is decreasing in  $z$ . If bankers have a larger fraction of total wealth, they can supply more deposits so the spread must fall to clear the deposit market.

Finally, using (15), (16), (19) and (20), the currency market clearing condition simplifies

to:

$$h = \frac{ak}{\beta} \alpha (1 - \beta) \iota(i, s)^{\epsilon-1} i^{-\epsilon} \quad (22)$$

Having solved for  $s(i, z)$ , (22) immediately gives the level of real currency holdings  $h(i, z)$ .

**Risk sharing.** The first order conditions for bankers' choice of  $\sigma_n$  and households' choice of  $\sigma_w$  are, respectively:

$$\sigma_{n,t} = \frac{\pi_t}{\gamma} + \frac{1 - \gamma}{\gamma} \sigma_{\xi,t} \quad (23)$$

$$\sigma_{w,t} = \frac{\pi_t}{\gamma} + \frac{1 - \gamma}{\gamma} \sigma_{\zeta,t} \quad (24)$$

The first term in each of (23) and (24) relates exposure to  $B$  to the risk premium  $\pi$ ; this is the myopic motive for choosing risk exposure: a higher premium will induce higher exposure. The second term captures the dynamic hedging motive, which depends on an income and a substitution effect. If the agent is sufficiently risk averse ( $\gamma > 1$ ), then the income effect dominates. The agent will want to have more wealth when his investment opportunities (captured by  $\xi$  and  $\zeta$  respectively) are worse.

Figure 1 illustrates the basic dynamic hedge from the point of view of a banker, based on the calibrated model. It plots the evolution of a banker's net worth after a 100 *bp* movement in interest rates (up or down) starting from the steady state, under two possible strategies. The dotted line represents a zero exposure strategy,  $\sigma_n = 0$ . The banker's net worth is not affected by the shock, but it's return on wealth is because the deposit spread moves with the interest rate. So after an increase in interest rates the banker's net worth starts to grow; after a fall in interest rates it starts to shrink. As a result, the banker's net worth is dynamically very volatile, and therefore so is his consumption. In contrast, the solid line represents the optimal strategy. The bank realizes a financial loss when interest rates go up, which is subsequently made up with higher returns. In exchange, after interest rates go down the bank realizes a financial gain. Overall, the banker's net worth and consumption is less volatile.



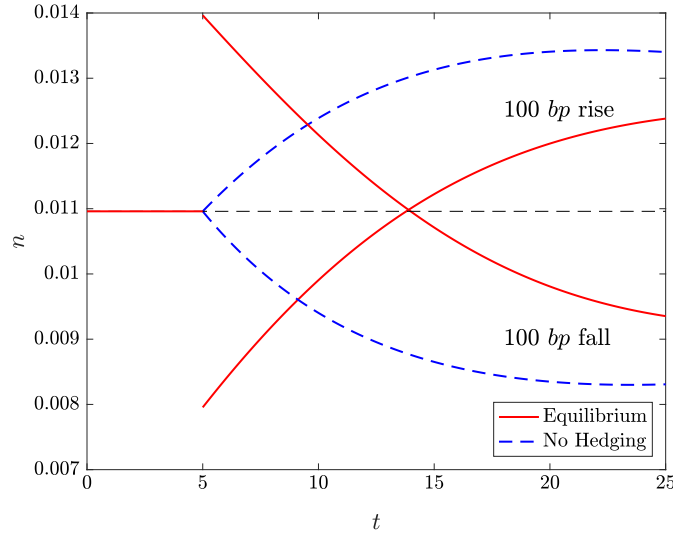


Figure 1: Impulse response of bank net worth to a 100 *bp* change in interest rates, with and without dynamic hedging.

From (23) and (24) we obtain the following expression for  $\sigma_z$ .<sup>6</sup>

$$\sigma_{z,t} = (1 - z_t) \frac{1 - \gamma}{\gamma} (\sigma_{\xi,t} - \sigma_{\zeta,t}) \quad (25)$$

The object  $\sigma_z$  measures how the bankers' share of wealth responds to the aggregate shock. The term  $\sigma_{\xi,t} - \sigma_{\zeta,t}$  in (25) captures the relative sensitivity of bankers' and households' investment opportunities to the aggregate shock. How this differential sensitivity feeds into changes in the wealth share depends on income and substitution effects. If agents are highly risk averse ( $\gamma > 1$ ) they will shift aggregate wealth towards bankers after shocks that worsen their investment opportunities relative to households', i.e.  $\frac{\xi}{\zeta}$  goes down. Notice that the premium  $\pi$  does not appear in equation (25). Households and bankers are equally able to earn any term premium, and have the same incentives to do so, so the level of the premium does not affect their relative exposure.

It is worth stressing that we cannot understand banks' risk taking behavior in isolation. Some other agent needs to take the other side (households in our model), so what matters is how monetary shocks affect their investment opportunities *relative* to households, as equation (25) shows. In other words, it is perfectly possible that neither banks nor households

<sup>6</sup>At this level of generality, this condition for aggregate risk sharing is analogous to the one in Di Tella (2017). However, the economic mechanism underlying the response of relative investment opportunities to aggregate shocks is specific to each setting.

prefer losses after interest rates increases (liquidity is scarcer and the economic environment therefore worse for all agents), but banks dislike this less than households.

We can use Ito's lemma to obtain an expression for  $\sigma_\xi - \sigma_\zeta$ :

$$\sigma_\xi - \sigma_\zeta = \left( \frac{\xi_z}{\xi} - \frac{\zeta_z}{\zeta} \right) \sigma_z z + \left( \frac{\xi_i}{\xi} - \frac{\zeta_i}{\zeta} \right) \sigma_i \quad (26)$$

Notice that  $\sigma_z$  enters the expression for  $\sigma_\xi - \sigma_\zeta$ : the response of relative investment opportunities to aggregate shocks depends in part on aggregate risk sharing decisions as captured by  $\sigma_z$ . This is because in equilibrium investment opportunities depend on the distribution of wealth  $z$ , so we must look for a fixed point. Replacing (26) into (25) and solving for  $\sigma_z$ :

$$\sigma_z = \frac{(1-z)^{\frac{1-\gamma}{\gamma}} \left( \frac{\xi_i}{\xi} - \frac{\zeta_i}{\zeta} \right)}{1-z(1-z)^{\frac{1-\gamma}{\gamma}} \left( \frac{\xi_z}{\xi} - \frac{\zeta_z}{\zeta} \right)} \sigma_i \quad (27)$$

**Implementation.** With complete markets, there is more than one way to attain the exposure dictated by equations (23) and (24). As long as  $\sigma_n$  is always negative the desired exposure can be implemented with a “traditional” banking balance sheet: long-term nominal assets, deposits as the only liability, and no derivatives. To be concrete, we'll imagine a banker's balance sheet with net worth  $n$ ,  $\phi n$  deposits as the only liability and  $(1+\phi)n$  nominal zero-coupon bonds that mature in  $T$  years as the only asset.

In the model, the price  $p^B(i, z; T)$  of a zero-coupon nominal bond of maturity  $T$  obeys the following partial differential equation:

$$\underbrace{\frac{p_i^B \mu_i + p_z^B \mu_z z + \frac{1}{2} [p_{ii}^B \sigma_i^2 + p_{zz}^B \sigma_z^2 z^2 + 2p_{iz}^B \sigma_i \sigma_z z]}{p^B}}_{\text{Nominal Capital gain}} - \frac{p_T^B}{p^B} - i = \pi \underbrace{\frac{p_i^B \sigma_i + p_z^B \sigma_z z}{p^B}}_{\text{Risk Premium}} \quad (28)$$

with boundary condition  $p^B(i, z, 0) = 1$  for all  $i, z$ . We use equation (28) to price bonds of all maturities at every point in the state space. The exposure to  $B$  of a traditional bank whose assets have maturity  $T$  is

$$\begin{aligned} \sigma_n &= (1+\phi) \sigma_{p^B} \\ &= (1+\phi) \frac{p_i^B(i, z; T) \sigma_i + p_z^B(i, z; T) \sigma_z z}{p^B(i, z; T)} \end{aligned} \quad (29)$$

We then find  $T(i, z)$  for each point in the state space by solving (29) for  $T$ , taking  $\sigma_n$

from the equilibrium of the model. We also consider an alternative implementation with geometric-coupon bonds.

## 4 Calibration

We make two minor changes to the baseline model to obtain quantitative results. First, we let productivity follow a geometric Brownian motion:

$$\frac{da_t}{a_t} = \mu_a dt + \tilde{\sigma}_a d\tilde{B}_t$$

where  $\tilde{B}_t$  is a standard Brownian motion, independent of  $B_t$ .<sup>7</sup> The economy scales with  $a$  so this change does not introduce a new state variable. The main effect of this change is to lower the equilibrium real interest rate. Second, in order to obtain a stationary wealth distribution we add tax on bankers' wealth at a rate  $\tau$  that is redistributed to households as a wealth subsidy. This tax can represent the administrative cost of running a bank.

We solve for the recursive equilibrium by mapping it into a system of partial differential equations for the equilibrium objects and solve them numerically using a finite difference scheme. Appendix A explains the modifications to the model and the numerical procedure in detail.

**Parameter values.** Table 1 summarizes the parameter values we use. We set the risk aversion parameter  $\gamma = 10$ , consistent with the asset pricing literature (see for instance Bansal and Yaron (2004)). We also perform a sensitivity analysis with different values of  $\gamma$ . EIS is 1 in our setting, in the interest of theoretical clarity and tractability, as explained above. It is also close to values used in the asset pricing literature. We choose the rest of the parameter values so that the model economy matches some key features of the US economy. The details of the data we use are in Appendix D.

We assume interest rates follow the Cox et al. (1985) stochastic process, so that  $\mu_i(i) = -\lambda(i - \bar{i})$  and  $\sigma_i(i) = \sigma\sqrt{i}$ . The concept of  $i$  in the model corresponds to a short term rate on an instrument that does not have the liquidity properties of bank deposits. We take the empirical counterpart to this to be the 6-month LIBOR rate in US dollars. We choose  $\bar{i} = 3.5\%$  to match the average LIBOR rate between 1990 and 2014. Estimating the

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<sup>7</sup>We assume that monetary policy is carried out so that the price level is also not sensitive to  $\tilde{B}$ , i.e.  $\tilde{\sigma}_p = 0$ .

Parameter	Meaning	Value
$\gamma$	Risk aversion	10
$\bar{i}$	Mean interest rate	3.5%
$\sigma$	Volatility of $i$	0.044
$\lambda$	Mean reversion of $i$	0.056
$\rho$	Discount rate	0.055
$\phi$	Leverage	8.77
$\alpha$	CES weight on currency	0.95
$\beta$	Cobb-Douglas weight on consumption	0.93
$\epsilon$	Elasticity of substitution between currency and deposits	6.6
$\mu_a$	Average growth rate of TFP	0.01
$\tilde{\sigma}_a$	Volatility of TFP	0.073
$\tau$	Tax on bank equity	0.195

Table 1: Parameter values

persistence parameter  $\lambda$  in a short sample has well known econometric difficulties (Phillips and Yu 2009). This parameter is very important in the model, for two related reasons. First, more persistence means that a change in interest rates has a long-lasting effect on bank spreads, which drive bankers’ relative desire to hedge. Second, more persistence means that a change in interest rates will have a large effect on the prices of long-term bonds, so the maturity  $T$  needed to implement any desired  $\sigma_n$  shortens. We set  $\lambda = 0.056$  and  $\sigma = 0.044$  to match the standard deviation of the LIBOR rate (2.4%) and 10-year Treasury yields (1.8%) for the period 1990-2014.

We use equation (20) to choose a value for the discount rate  $\rho$ . The Flow of Funds reports a measure of aggregate wealth. To be consistent with our model which has no labor, we adjust this measure by dividing by 0.35 (the approximate capital share of GDP) in order to obtain a measure of wealth that capitalizes labor income. We then compute an average consumption-to-adjusted-wealth ratio between 1990 and 2014, taking consumption as consumption of nondurables and services from NIPA data. This results in  $\frac{ak}{\omega} = 5.1\%$ , which, given the value of  $\beta$  set below, leads to  $\rho = 0.055$ .

We use data on bank balance sheets from the Flow of Funds to set a value of the leverage parameter  $\phi$ . In the model there is only one kind of liquid bank liability (“deposits”) whereas in reality banks have many type of liabilities of varying degrees of liquidity, so any sharp line between “deposits” and “not deposits” involves a certain degree of arbitrariness. We choose the sum of checking and savings deposits as the empirical counterpart of the model’s deposits, leaving out time deposits since these are less liquid and the spreads that banks

obtain on them are much lower. We set  $\phi = 8.77$  to match the average ratio of deposits to bank net worth between 1990 and 2014.

We construct a time series for  $z$  using data on banking sector net worth and total wealth from the Flow of Funds (total wealth is divided by 0.35 as before to account for labor income). The Flow of Funds data uses book values, which is the right empirical counterpart for  $n$  in the model (market value of banks' equity includes the value of investment opportunities which is not part of  $n$ ). We then use the data from Drechsler et al. (2017) on interest rates paid on checking and savings deposits and weight them by their relative volumes from the Flow of Funds to obtain a time series for the average interest rate paid on deposits.<sup>8</sup> We subtract this from LIBOR to obtain a measure of spreads. We set  $\beta$  (the Cobb-Douglas weight on consumption as opposed to money),  $\alpha$  (the CES weight on currency as opposed to deposits), and  $\epsilon$  (the elasticity of substitution between currency and deposits) jointly to minimize the sum of squared distances between the spreads predicted by equation (21), given the measured time series for  $i$  and  $z$ , and the measured spreads. The data seem to prefer very high values of  $\alpha$  so we, somewhat arbitrarily, fix  $\alpha = 0.95$  (letting  $\alpha$  take even higher values does not improve the fit very much). Minimizing over  $\beta$  and  $\epsilon$  leads to  $\beta = 0.93$  and  $\epsilon = 6.6$ .

We set the growth rate of productivity  $\mu_a = 0.01$  and its volatility  $\tilde{\sigma}_a = 0.073$  for the model to match the average real interest rate between 1990 and 2014, which was 1%. This value of  $\tilde{\sigma}_a$  is close to that used by He and Krishnamurthy (2012), who use  $\tilde{\sigma}_a = 0.09$ .

Finally, we set the tax rate of bank capital to  $\tau = 0.195$  for the average value of  $z$  in the model to match the average value in the data between 1990 and 2014, which is 0.56%. Given  $\phi = 8.77$ , this is equivalent to 2% of assets, equal to the administrative expenses ratio reported by Drechsler et al. (2018).

**Spreads.** Since the behavior of deposit spreads plays a central role in the mechanism, it is worth checking how our model accounts for them. Figure 2 shows the spread as a function of  $i$  and  $z$  for our parameter values. As we know from equation (21), it is increasing in  $i$  and decreasing in  $z$ . Furthermore, it is concave in  $i$ . When  $i$  is high, agents are already holding very little currency, so further increases in  $i$  do not generate as much substitution into deposits and therefore don't lead to large increases in spreads.

These properties of  $s(i, z)$  are consistent with the data. Table 2 shows the results of regressing spreads on interest rates and banks' share of total wealth. The first column,

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<sup>8</sup>The data ranges from 1999 to 2008. We thank Philipp Schnabl for kindly sharing this data with us.

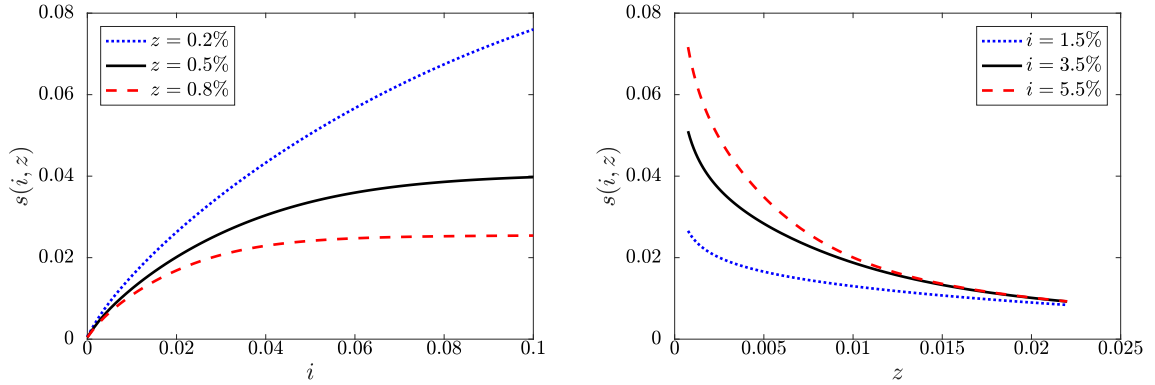


Figure 2: Spreads in the model as a function of  $i$  and  $z$ .

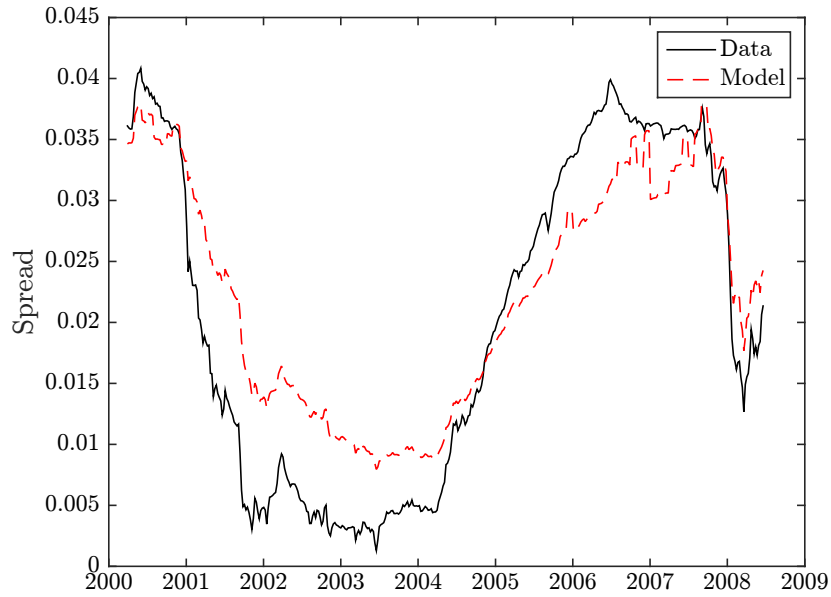


Figure 3: Spreads in the data compared to spreads implied by the  $s(i, z)$  function given our parameter values and the measured time series of  $i$  and  $z$ .

without a quadratic term, shows that a one percentage point increase in LIBOR is associated with a 66 basis points increase in bank spreads, while a one percentage point increase in banks' share of total wealth is associated with a 99 basis points fall in bank spreads. The second column, including a quadratic term, shows that there is indeed evidence that bank spreads flatten out slightly as  $i$  increases.

	(1)	(2)
Constant	0.3%	-0.3%
	(0.22%)	(0.44%)
$i$	0.66	0.98
	(0.028)	(0.17)
$i^2$	-	-4.07
		(2.02)
$z$	-0.99	-0.71
	(0.25)	(0.32)
$R^2$	0.89	0.90
$N$	430	430

Note: The dependent variable is the spread. Newey and West (1987) standard errors are in parentheses.

Table 2: Spreads, interest rates and banks' share of aggregate wealth.

The model is able to match the time series behavior of spreads quite closely. Figure 3 compares the time series for  $s(i, z)$  produced by the model with the time series of measured spreads from the Drechsler et al. (2017) data.

## 5 Exploring the Mechanism

In this section we describe, quantitatively, how the model works. Bankers' mark-to-market net worth is exposed to interest rate risk as part of equilibrium risk-sharing. After an increase in interest rates, bankers' net worth falls, which forces them to reduce their supply of deposits, amplifying the effects of the monetary shock on the cost of liquidity.

**Aggregate risk sharing.** Bankers' mark-to-market net worth is highly exposed to movements in interest rates. The top panels of Figure 4 show bankers' exposure. If the nominal interest rate rises by 100 basis points, bankers' net mark-to-market net worth changes by  $\frac{\sigma_n}{\sigma_i}$  percent. It is always negative, so banks face falls in net worth after an increase in nominal interest rates. Quantitatively, the effect is quite large. At the mean levels of  $i$  and  $z$ , if interest rates rise by 100 basis points, banks lose about 30% percent of their mark-to-market

net worth. As discussed in Section 6, this magnitude is consistent with what we observe in bank balance sheets.

To understand the mechanism, note that because aggregate wealth is insensitive to  $B$ ,  $\sigma_n = \sigma_z$  so movements in bankers' net worth and in their share of total net worth are equivalent. We know from (21) that an increase in the nominal interest rate raises the spread  $s$ . Since bankers earn this spread and households don't, bankers' relative investment opportunities  $\frac{\xi}{\zeta}$  improve when the interest rate  $i$  rises, as shown in the middle-left panel of Figure 4. Equation (25) implies that  $z$  must fall in response, which further raises the spread  $s$ , amplifying the effect of monetary shocks on the cost of liquidity. As a result, bankers' relative investment opportunities  $\frac{\xi}{\zeta}$  improve even more, as shown in the middle-right panel of Figure 4, which amplifies bankers' incentives to choose a negative  $\sigma_n$  (this is the reason the denominator in equation (27) is less than one).

The hedging motive weakens at higher levels of  $i$  and  $z$ ;  $\frac{\sigma_n}{\sigma_i}$  is greater (in absolute value) for low  $i$  and  $z$ . This reflects the behavior of spreads. As shown in Figure 2, the spread flattens out for higher  $i$  and  $z$ . As a result, relative investment opportunities are less sensitive to  $i$  when  $i$  or  $z$  are high, so bankers choose lower exposure. To see the link between flattening spreads and lower exposure, we re-solved the banker's problem replacing the equilibrium  $s(i, z)$  by the *linear* form  $s(i, z) = 0.3\% + 0.66i - 0.99z$ , which is the best linear approximation to the data, as shown on Table 2. Since the sensitivity of spreads to  $i$  is constant in this experiment, the banker's exposure  $\frac{\sigma_n}{\sigma_i}$  is almost constant as a function of  $i$  and  $z$ .<sup>9</sup>

While bankers choose a large exposure to interest rate risk, TFP shocks  $\tilde{B}$  are shared proportionally by both banks and households:  $\tilde{\sigma}_n = \tilde{\sigma}_w = \tilde{\sigma}_a$ . The reason for this is that these TFP shocks don't affect the investment opportunities of banks relative to households, so there is no relative hedging motive as in equation (26). Our theory therefore provides not only an explanation for why banks are exposed to risk in general, but also why they are exposed to interest rate risk in particular. A similar line of argument indicates that if banks also earn a credit spread, dynamic hedging motives would explain why they choose to be exposed to changes in this spread, as documented by Begenau et al. (2015).

**Maturity mismatch.** We can implement banks desired exposure to interest rate risk  $\sigma_n$  with a traditional maturity-mismatched balance sheet as explained in Section 3. The resulting maturity mismatch is shown on the third row of Figure 4. At the mean levels of  $i$  and  $z$ , the maturity mismatch  $T$  needed to implement the desired exposure  $\frac{\sigma_n}{\sigma_i}$  is 3.6 years.

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<sup>9</sup>Available upon request.



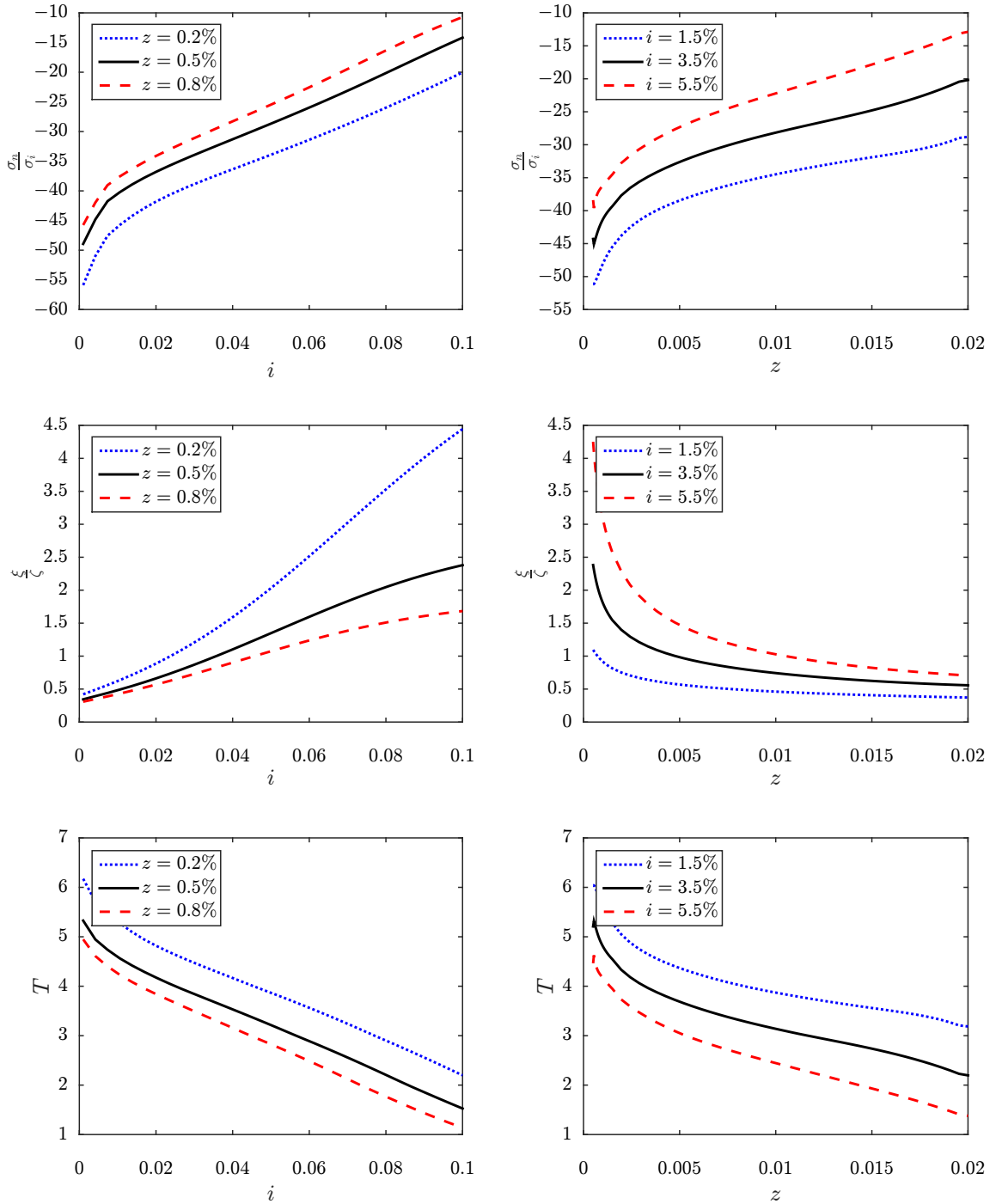


Figure 4: Aggregate risk sharing.

$T$  is decreasing in both  $i$  and  $z$ . This reflects the higher desired exposure when  $i$  and  $z$  are low, which in turn results from the higher sensitivity of spreads to  $i$  in this region.<sup>10</sup> Notice that the increase in the desired maturity mismatch when interest rates are low may look like “reaching for yield”, but it’s precisely what risk averse agents should do to insure against greater deposit spread risk.

**Amplification.** The endogenous response of bankers’ net worth to movements in interest rates amplifies the effect of monetary shocks on the cost of liquidity. Low net worth limits bankers’ ability to supply liquidity, and drives up the equilibrium spread on deposits. In other words, equilibrium risk exposures imply that the quantity of deposits falls in response to interest rate increases, as documented by Drechsler et al. (2017). As a result, an increase in nominal interest rates has a direct effect on deposit spreads and an indirect effect through weaker bank balance sheets  $z$ . We can decompose the response of spreads to monetary shocks as follows:

$$\frac{\sigma_s}{\sigma_i} = \underbrace{\frac{\partial s}{\partial i}}_{\text{Direct Effect}} + \underbrace{\frac{\partial s}{\partial z} \frac{\sigma_z z}{\sigma_i}}_{\text{Indirect Effect}}$$

If interest rates go up by 100 *bp*, the direct effect on the deposit spread is  $\frac{\partial s}{\partial i} \times 100 \text{ bp}$ , and the indirect effect is  $\frac{\partial s}{\partial z} \frac{\sigma_z z}{\sigma_i} \times 100 \text{ bp}$ . In the calibrated model, the average (over the stationary distribution) effect of a 100 *bp* increase in interest rates is an increase in deposit spreads of 77 *bp*. This is decomposed into a direct effect of interest rates of 62 *bp* and an indirect effect through the the endogenous response of banker’s net worth of 15 *bp*.<sup>11</sup> Banks’ exposure to interest rate risk therefore amplifies the effect of monetary shocks on deposit spreads by a factor of 1.25 on average.

The endogenous amplification is non-linear, however, as shown in Figure 5. The direct effect of interest rates on deposit spreads is decreasing in both  $i$  and  $z$  because in both cases currency becomes a smaller fraction of total liquidity. The strength of the indirect effect is governed by two opposing forces. First, when  $i$  or  $z$  is high, deposits are a larger fraction of total liquidity, so changes in their supply have a bigger effect on spreads. On the other hand, banks’ hedging motive weakens for high  $i$  or  $z$ , so  $z$  itself is less sensitive to  $i$ . On balance, the indirect effect is increasing in  $i$  and slightly decreasing in  $z$ . The amplification

<sup>10</sup> $T$  depends on both the desired exposure  $\sigma_n$  and the sensitivity of bond prices  $\sigma_{pB}$  for each maturity; the latter does vary with  $i$  but not by much, so the movement in  $T$  reflects mostly the movement in  $\frac{\sigma_n}{\sigma_i}$ .

<sup>11</sup>The direct effect can be compared directly to the OLS coefficient on  $i$  from Table 2, which is 66 basis points. The indirect effect can be compared to the OLS coefficient on  $z$ , which is  $-99$  basis points, times the average value of  $\frac{\sigma_z z}{\sigma_i}$ , which is  $-0.16$ .

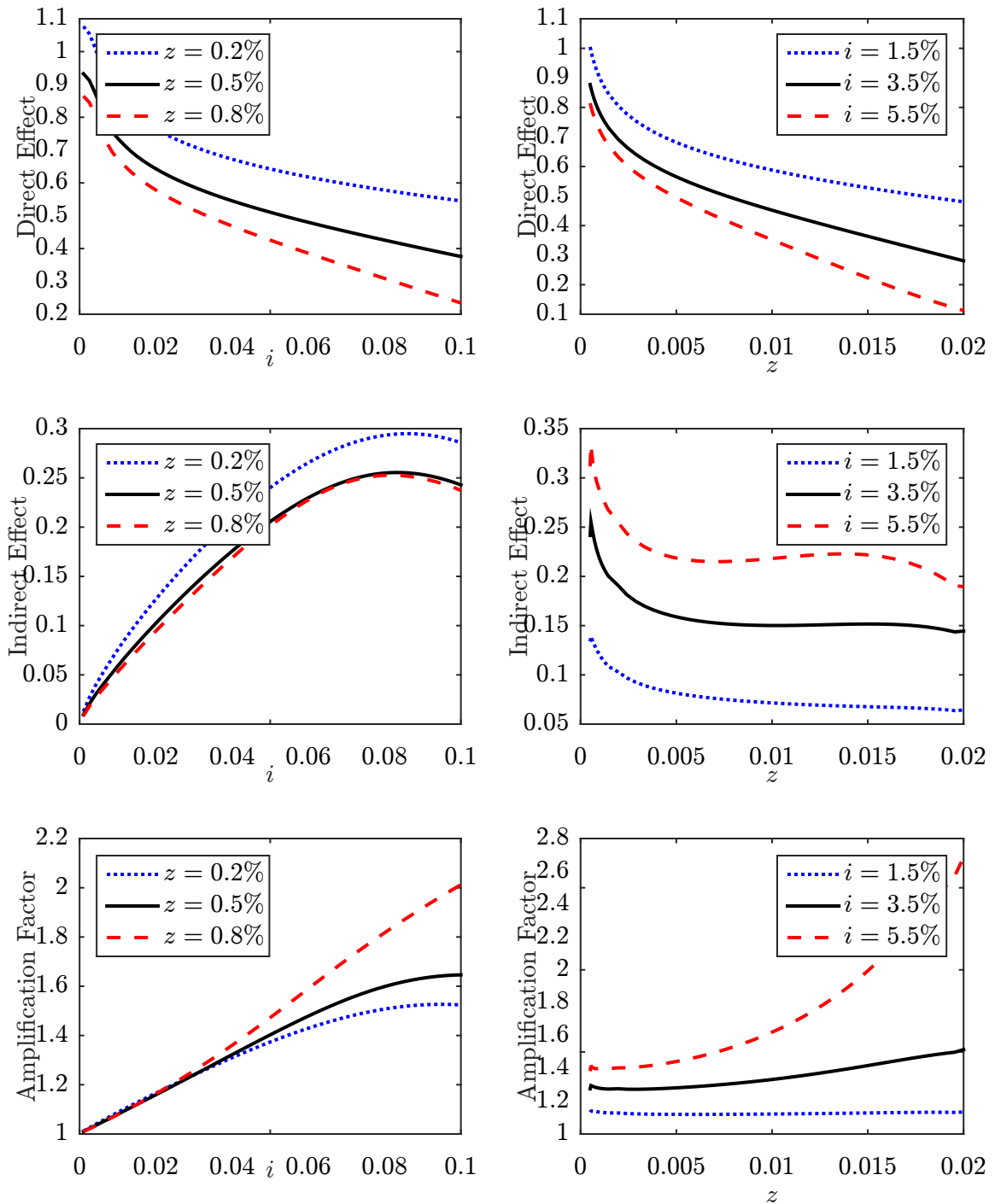


Figure 5: Sensitivity of deposit spreads to interest rates through endogenous response of banks' net worth.

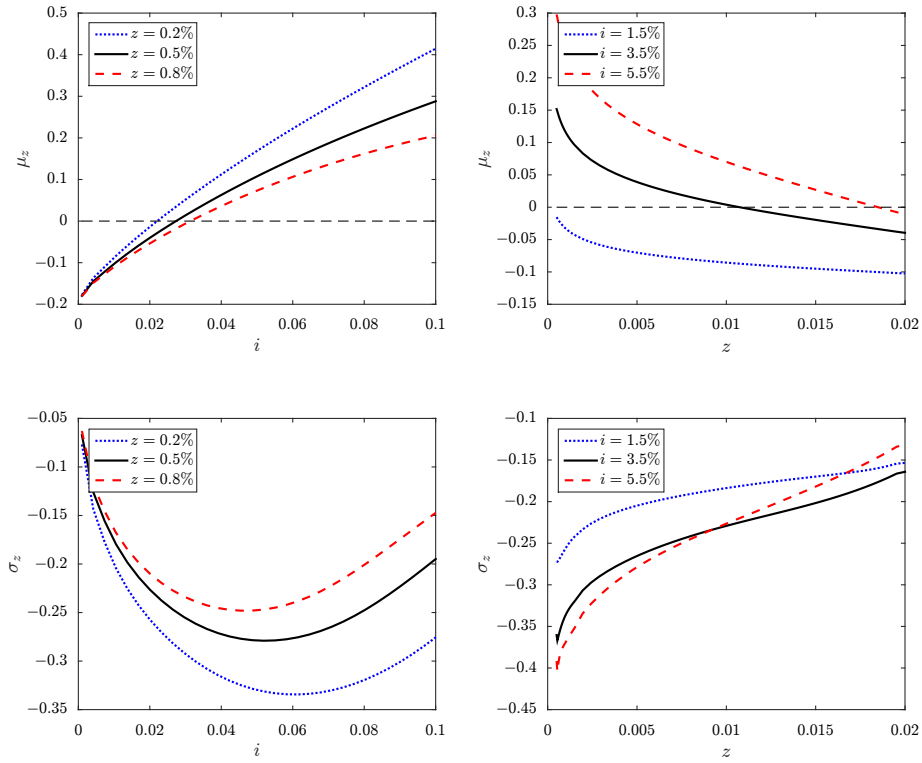


Figure 6: The drift of  $z$ ,  $\mu_z$  (upper panels) and its volatility  $\sigma_z$  (lower panels).

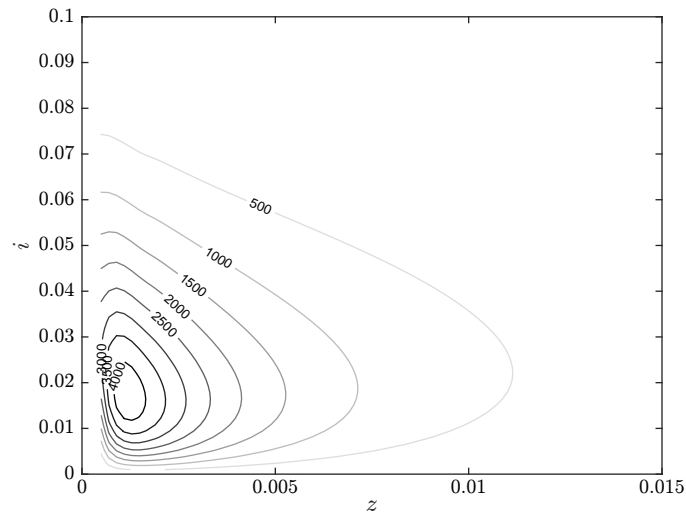


Figure 7: Stationary distribution over  $(i, z)$ .

factor is computed simply as the ratio of the total effect to the direct effect, and is increasing in both  $i$  and  $z$ . For example, when nominal interest rates are 5.5% and  $z = 1\%$ , a 100 *bp* increase in interest rates has a total effect on deposit spreads of 57 *bp*, of which the direct effect of interest rates is 35 *bp* and the indirect effect through banks' balance sheets is 22 *bp*. The amplification factor in this case is 1.62.

**Dynamics.** Agents' endogenous exposure to interest rate risk leads to the equilibrium dynamics shown in Figure 6. The upper panels show the drift of bankers' share of aggregate wealth  $z$  and the bottom panels its sensitivity to  $B$ . The drift of  $z$  is positive for small  $z$  and high  $i$ , because in this region the spread is high.

On impact, banks take losses when interest rates rise. Since total wealth  $\omega$  is fixed, their share of aggregate wealth  $z$  falls. This is reflected in the bottom panels of Figure 6, where  $\sigma_z$  is negative. Over time, higher interest rates mean higher spreads and bank balance sheets strengthen. The resulting stationary distribution is shown in Figure 7.

## 6 Quantitative Evaluation

In this section we test the model by comparing its quantitative predictions to the empirical evidence. The model can successfully account for the level, time pattern, and cross-sectional pattern of banks' maturity mismatch.

**Measuring banks' maturity mismatch.** Following the methodology of English et al. (2018), we construct an empirical measure of banks' maturity gap. Using Call Reports data, we record the contractual maturity (in case of fixed-rate contracts) or repricing maturity (in case of floating-rate contracts) for each line of the balance sheet, take weighted averages of assets and liabilities, and subtract. We then compute an aggregate measure by taking the asset-weighted median across banks for each quarter.<sup>12</sup> See Appendix D for details.

**Time-series evidence.** The model closely matches the empirical behavior of banks' maturity mismatch. Figure 8 compares the time series for  $T$  predicted by the model with the

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<sup>12</sup>The maturity gap measures the on-balance-sheet exposure, not the exposure through derivatives. English et al. (2018) show that for the majority of banks, this makes no difference since they do not trade derivatives. However, the evidence in Begenau et al. (2015) indicates that, especially for the largest banks, derivatives amplify interest rate exposure, so just measuring the on-balance-sheet positions underestimates the maturity mismatch. On the other hand, the option to refinance fixed-rate mortgages lowers their effective maturity. Using the contractual maturity therefore overestimates the maturity mismatch.

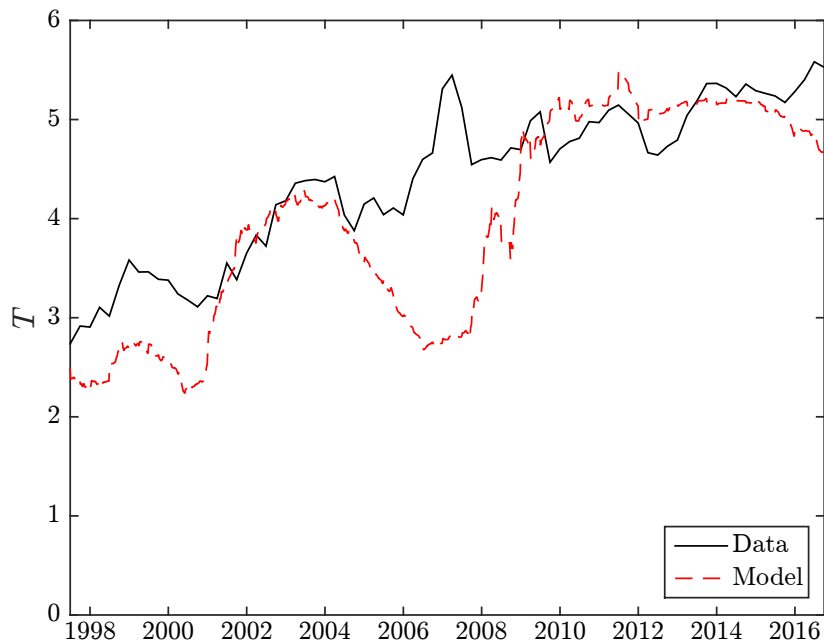


Figure 8: Maturity mismatch of banks in the data and in the model.

data. For the model values, we simply plug in the measured time series of  $i$  and  $z$  into the function  $T(i, z)$  produced by the model. This is a way to test the goodness of fit of the function  $T(i, z)$  over the range of  $i$  and  $z$  in the data. The average  $T$  in the data is 4.4 years; in the model, it's 3.9 years.<sup>13,14</sup> The model is less successful at the beginning of the financial crisis in 2007, where it underpredicts  $T$ .

The model also reproduces the time pattern in the data. The correlation between the model and the data is 0.77. To understand this time pattern, recall from Figure 4 that the model predicts that banks' maturity mismatch  $T$  should be larger during periods of low interest rates because deposits spreads are more sensitive to movements in interest rates. This basic correlation is borne out by the data. Table 3 shows the results of an OLS regression of banks' maturity mismatch  $T$  on  $i$  and  $z$ . A 100 bp increase in  $i$  is associated with decrease in  $T$  of 0.12 years; a 100 bp increase in  $z$  is associated with a decrease in  $T$  of 0.019 years.

<sup>13</sup> $T$  and  $\frac{\sigma_n}{\sigma_i}$  are alternative metrics for risk exposure. For example, at the mean of the stationary distribution, where the model predicts  $T = 3.6$ , a 100 bp increase in interest rates results in a 3.2% fall in the price of a 3.6-year bond. Amplified by leverage of  $\phi = 8.77$ , this results in a fall in bank net worth of 31%, consistent with Figure 4.

<sup>14</sup>4.4 years is the time-series average of the asset-weighted cross-sectional median. The time-series average of the asset-weighted cross-sectional mean is 4.5.

Higher maturity mismatch during periods of low interest rates may look like “reaching for yield”. But it’s precisely what the dynamic hedging mechanism calls for.

	(1)
Constant	4.4
	(0.1)
$i$	-11.7
	(6.8)
$z$	-1.9
	(0.4)
$R^2$	0.63
$N$	78

Note: The dependent variable is the asset-weighted cross-sectional median maturity mismatch  $T$ . Newey and West (1987) standard errors are in parentheses.  $i$  and  $z$  are demeaned.

Table 3: Maturity mismatch of banks, interest rates, and banks’ share of aggregate wealth.

**Cross-sectional evidence.** We can use cross-sectional data to further test the mechanism. Although the model features banks with identical deposit-to-net-worth ratios, its logic implies that banks with a larger deposit base (relative to net worth) should have a larger maturity mismatch in order to hedge their larger exposure. This prediction is borne out by the cross-sectional data.

First we compute the quantitative relationship between the deposit-to-net-worth ratio  $\phi$  and the maturity mismatch  $T$  predicted by the model. We re-solve the individual banker’s problem for different values of the deposit-to-net-worth ratio  $\phi$ , taking the model’s equilibrium prices as given. For each value of  $\phi$  we then compute the time series of the maturity mismatch  $T$  as above. We then compute a time-series average  $T$  for each  $\phi$ . We find that a unit increase in  $\phi$  (i.e. increasing deposits by one time net worth) is associated with an increase in average  $T$  of 0.42 years. The relationship is almost linear in the range  $\phi \in [4, 13]$ .

We then measure the same relationship in the data. For each bank in our sample, we compute the time-series average of the maturity mismatch and the deposit-to-net-worth ratio. We then run an (asset-weighted) median regression of the maturity mismatch on the deposit-to-net-worth ratio, on the cross-section of banks. The results are reported in Table 4. A unit increase in the deposit-to-net-worth ratio is associated with an increase in the weighted median maturity mismatch of 0.43 years, which coincides almost exactly with the quantitative prediction of the model. Table 4 also reports the results of an OLS regression; a one unit increase in the deposit-to-net-worth ratio is associated with an increase in the weighted average maturity mismatch of 0.26 years. The difference between the median and

	Median	OLS
Constant	2.6 (0.0023)	3.6 (0.063)
$\phi$	0.43 (0.0004)	0.26 (0.013)
$N$	10, 351	10, 351

Note: The dependent variable is the time-averaged maturity mismatched. Standard errors are in parentheses.

Table 4: Maturity mismatch of banks and deposit-to-net-worth ratio

OLS regressions is evidence that the distribution of the maturity mismatch is more right-skewed for banks with higher deposit-to-net-worth ratio. Our model cannot speak to this interesting fact, since we assume all banks with the same deposit-to-net-worth ratio to be identical.

**Term premium.** A traditional argument for why banks take interest rate risk is that they do so to take advantage of the term premium. In general equilibrium, this argument is incomplete: the fact that there is a term premium cannot explain why banks choose to earn it and households choose to pay it. Households could earn the term premium just as easily as bankers (for example, by investing in long-term bonds), but in equilibrium choose not to.

Our model does produce a term premium, and banks do earn it, but it does not play any role in determining their risk exposure. The equilibrium term premium reflects that states of the world with high interest rates are bad for everyone because liquidity, which is part of the consumption bundle, is expensive. Bankers' deposit-taking business is a natural hedge, so high interest rate states of the world are relatively better for bankers than for households. This creates a relative hedging motive, so banks end up taking risk and earning the premium while households are willing to pay the premium to partially insure themselves against high interest rates.

This is reflected in FOCs (23) and (24). The premium  $\pi$  does appear on both of these equations: both agents want to take advantage of the term premium. But the premium is endogenous, the market has to clear. As a result  $\pi$  drops out of equation (25) for  $\sigma_z$  and only the relative hedging motive matters.

Quantitatively, the premium in the model is significant but lower than in the data. We can compute the excess return on a long term nominal bond simply as

$$ER = \sigma_{pB} \pi$$



Piazzesi and Schneider (2007) report an average excess return on 5-year treasuries of 99 basis points. In the model, this excess return is 22 basis points, so the forces in the model explain about a fifth of the term premium.<sup>15</sup>

**The role of risk aversion and risk-seeking.** Since the mechanism in this model is related to dynamic hedging, and more broadly to asset pricing, we use a value for the coefficient of relative risk aversion  $\gamma = 10$  in the range that has been found useful in matching asset pricing data, as in Bansal and Yaron (2004) or Bansal et al. (2009). However, there is no consensus in the literature on the appropriate value for this parameter, and lower values are more typical in macroeconomic models. We therefore perform a sensitivity analysis with  $\gamma = 3$ ,  $\gamma = 6$  and  $\gamma = 20$ . In each case, we set the rest of the parameter values to match the same targets and re-compute the time series for  $T$  predicted by the model. The results are shown in Table 5. Even with  $\gamma = 3$ , we get a significant maturity mismatch  $T = 2.6$ . Note that the maturity mismatch *increases* with risk aversion: banks take interest rate risk to insure against their stochastic investment opportunities. The effect is therefore stronger the higher risk aversion is. This can be seen in equation (25) (in particular, with  $\gamma = 1$  we would get  $T = 0$ ).

A widespread concern among regulators is banks' potentially risk-seeking behavior. Our quantitative results show no evidence of risk-seeking by the aggregate banking sector, at least with respect to interest rate risk. Banks' exposure to interest rate risk is consistent with a dynamic hedging strategy by highly risk averse agents. The main insight is that banks have a large underlying exposure to movements in interest rates arising from the deposit spread. If banks did not expose their balance sheet to interest rates at all (for instance by having no maturity mismatch) they would in fact be taking on a large amount of risk due to the sensitivity of deposit spreads to interest rates, as illustrated in Figure 1. A back-of-the-envelope calculation shows that a 100 *bp* fall in interest rates would make banks lose in net present value of future deposit spreads about 49% of their net worth.<sup>16</sup> Their maturity mismatch leads them to recover about 30% of net worth up front. This partially offsets the underlying risk.

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<sup>15</sup>See Haddad and Sraer (2019) and Paul (2020) for further discussion of banks and term premia.

<sup>16</sup>A fall in interest rates of 100 *bp* results in a fall in spreads of about 66 *bp* and a fall in the return on equity of  $\phi \times 66 = 579$  *bp*. The NPV of 49% follows from applying a mean reversion rate of  $\lambda = 5.6\%$  and a discount rate of 6.4%.

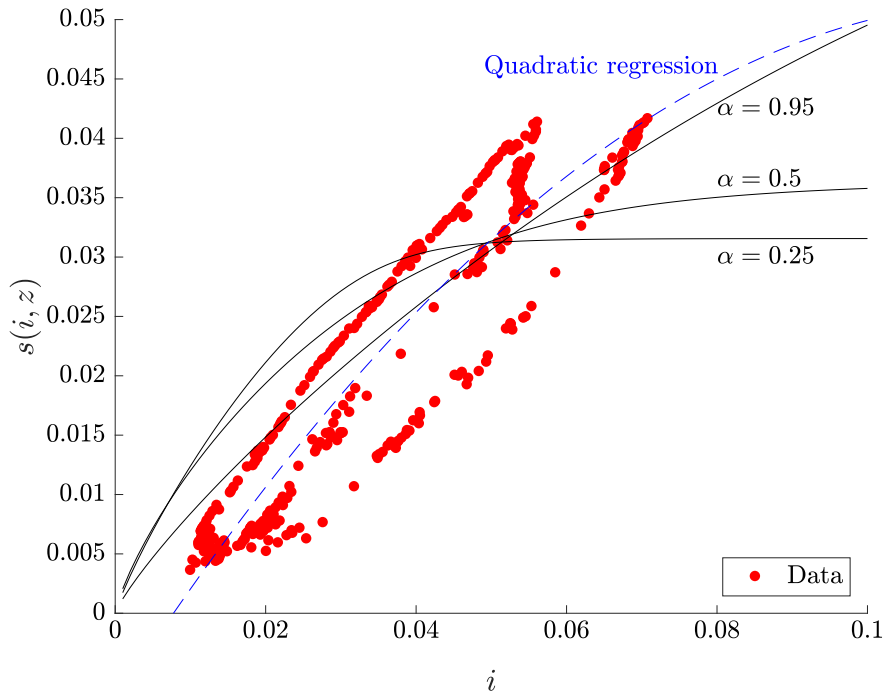


Figure 9: Spreads resulting from different values of  $\alpha$ , fixing  $z$  at the sample mean.

**The role of currency.** In order to account quantitatively for the positive co-movement of spreads and interest rates, the model requires a high value of  $\epsilon$  (i.e. a high elasticity of substitution between currency and deposits). Furthermore, unless  $\alpha$  is high (i.e. there is a strong preference for currency), the co-movement weakens rapidly as the interest rate rises. The reason is that as interest rates rise, there is less and less currency to substitute away from. Therefore unless there is a lot of currency to begin with, substitution between currency and deposits weakens, and with it the co-movement of interest rates and deposit spreads. Hence the better fit of the model with high values of  $\alpha$ .

Figure 9 shows the function  $s(i, z)$  that results from re-calibrating  $\epsilon$  and  $\beta$  to best match the spread data while fixing different values of  $\alpha$ , together with a scatterplot of spreads against interest rates (after controlling for  $z$ ). For lower values of  $\alpha$ , the model-produced  $s(i, z)$  function can still match observed spreads fairly well at low interest rates, but the relationship flattens out too much relative to the data at interest rates higher than about 5%.<sup>17</sup>

One consequence of setting  $\alpha = 0.95$  is that the model produces high and variable

<sup>17</sup>This implies that setting a lower value of  $\alpha$  in the model would make banks' maturity mismatch too sensitive to interest rates relative to the data, and too low at higher interest rates.

currency holdings, much more so than in the data. The model just cannot simultaneously match the behavior of spreads and the quantity of currency. Overall, we conclude that our microeconomic model of bank spreads is probably too simplistic and the observed co-movement of interest rates and bank spreads is also driven by imperfect competition between banks (Drechsler et al. 2017), stickiness in deposit rates (Hannan and Berger 1991, Driscoll and Judson 2013), search costs (Yankov 2018), or other factors.

For the purposes of bank risk management, the exact microeconomic mechanism that drives spreads is not so essential. What matters is how these co-move with interest rates and banks' share of wealth, and the model matches this quite well. Furthermore, the model also matches the quantity of deposits because we target bank leverage, and consumption-to-wealth and bank-equity-to-wealth ratios in our calibration.

**Alternative maturity structure.** The model pins down banks' exposure to interest rate risk, but this exposure can be implemented in many ways. The baseline implementation has a single zero-coupon bond of maturity  $T$ . This is a simple portfolio structure that is easy to interpret. Another simple implementation uses bonds with a geometric maturity structure. A bond with coupon payments  $\delta e^{-\delta t}$  has average maturity  $1/\delta$ . We can compute the value of such a bond as a bundle of zero-coupon bonds,

$$p^{GB}(i, z; \delta) = \int_0^\infty \delta e^{-\delta T} p^B(i, z; T) dT$$

Then we find the  $\delta(i, z)$  that implements the desired exposure to interest rate risk,  $\sigma_n = (1 + \phi)\sigma_{p^{GB}(\delta)}$ , for every  $(i, z)$ .

Table 5 summarizes the average maturity  $1/\delta(i, z)$  for different values of  $\gamma$ . The results are very similar, but the average maturity mismatch is a little higher with geometric bonds. The reason is that the duration of geometric bonds is lower than the average maturity of their coupon payments, so a longer average maturity is necessary to obtain the same risk exposure (with zero-coupon bonds, duration and maturity coincide). For example, with a constant nominal interest rate, the duration of a geometric bond is  $1/(i + \delta)$  which is less than its maturity  $1/\delta$ .

**Non mark-to-market accounting and net interest margins.** Our model is cast entirely in mark-to-market terms. However, accounting rules do not require marking to market for the majority of long term assets held in the "banking" book, and these accounting rules determine how financial statements will respond to interest rate shocks.

$\gamma$	3	6	10	20	data
Average $T$ (zero-coupon)	2.6	3.5	3.9	4.3	4.4
Average $T$ (geometric)	3.1	4.5	5.1	5.8	4.4

Table 5: Average maturity mismatch for different values of risk aversion and zero-coupon vs. geometric bonds

Suppose there is an increase in interest rates. In mark-to-market terms, banks experience an immediate capital loss.<sup>18</sup> Afterwards, they accrue interest on long-term assets at the *new* rate, so net interest margins increase. Under non-mark-to-market accounting, banks keep their long-term assets on their books at par, record no capital loss, and continue to accrue interest on them at the *old* rate. On the liability side, since they pass on part of the rate increase to depositors, interest expenses increase. Therefore, the measured net interest margin would initially fall. As long-term assets mature and are replaced by new assets yielding the new rate, the measured net interest margin would start rising, eventually exceed its initial level and then level off as the interest rate mean-reverts.

The exact time pattern of this impulse response is indeterminate. Banks in the model can implement the equilibrium risk exposure with different maturity profiles, such as zero-coupon bonds, geometric bonds, combinations of long and short bonds, etc. Each of these would yield a different time path for the book net interest margin. Figure 10 shows the impulse response of the net interest margin under a particular implementation, which approximates banks' observed maturity profile. We consider a bank which invests 37% of its assets in an instantaneous bond (this is the fraction of bank assets that mature within one quarter in our data) and the remainder in a geometric bond with  $\delta = 0.127$  (or 7.9 years maturity). This portfolio produces the equilibrium risk exposure at the rest point of the state space. The figure shows the impulse response of book net interest margins after a 100 b.p. increase in the interest rate (see Appendix C for details of the calculation). The figure shows that, even initially, net interest margins barely respond. The fraction of assets held in short-term bonds (whose book yields respond immediately) is very close to the pass-through of interest rates to deposit rates, so even in the short run the book yields of assets and liabilities respond similarly. Over time, as the geometric bond matures and is reinvested, book yields on the

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<sup>18</sup>When we mark to market, we treat deposits at face value because that's what the bank would have to pay if it were liquidated. Hoffmann et al. (2018) propose an alternative marking-to-market approach. They estimate deposits' effective duration, which is longer than zero because deposit rates respond less than one for one to market rates. This means that the present value of payments on deposits falls with interest rates, and deposits are market to market accordingly. On the basis on this calculation for European banks, Hoffmann et al. (2018) find that the average the interest rate risk exposure of banks' net worth is close to zero, which is consistent with a fall in net worth under our mark-to-market measure.

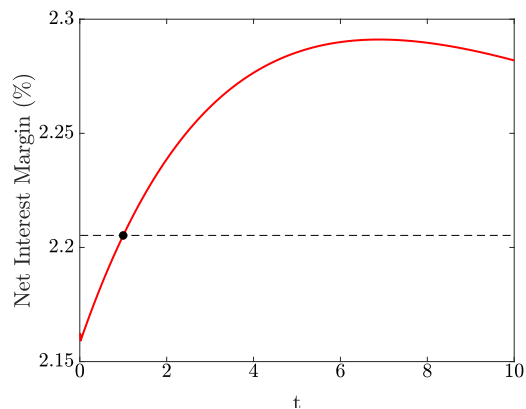


Figure 10: Impulse response of net interest margins to a 100 b.p. increase in interest rates

bank's assets start to increase, and the net interest margin is back to its original level after precisely one year. Book yields on assets then continue to increase for some time, before falling again as interest rates mean revert. Overall, the net interest margin is never more than 9 b.p. away from its starting point.

Drechsler et al. (2018) estimate impulse responses of changes in book net interest margins to changes in interest rates and report the cumulative response over four quarters (i.e. the change in net interest margins after one year). They find it to be almost exactly zero, which is precisely in line with the predictions from our model under this particular implementation. Other implementations produce slightly different but qualitatively similar impulse responses. For instance, if the bank uses a single geometric bond, the net interest margin falls more initially and rises more slowly, returning to its original level after 2.1 years.

## 7 Real Shocks under Inflation Targeting

Up to this point we have assumed that the only source of changes in interest rates is monetary policy shocks. This benchmark was useful to examine the mechanisms at play without confounding factors. In this section we look at the opposite case, where monetary policy follows an inflation targeting rule and changes in interest rates are the result of real shocks. There are many possible real shocks that could have an effect on equilibrium real interest rates. We will focus on a simple case, where the only shock is a change in the expected growth rate of TFP.

In particular, we assume that the growth rate of productivity  $\mu_a$  is stochastic and follows:

$$d\mu_{a,t} = -\lambda(\mu_{a,t} - \bar{\mu}_a) dt + \sigma\sqrt{\mu_{a,t} - \mu_a^{min}}dB_t$$

This is a Cox et al. (1985) stochastic process for the growth  $\mu_{a,t} - \mu_a^{min}$ .  $\bar{\mu}_a$  is the mean growth rate and  $\mu_a^{min}$  is a, possibly negative, lower bound.

Monetary policy consists of targeting a constant rate of inflation  $\bar{\mu}_p$  (keeping  $\sigma_p = \tilde{\sigma}_p = 0$  as before). Therefore the nominal interest rate is just:

$$i_t = r_t + \bar{\mu}_p$$

where  $r_t$  is the endogenous real interest rate.

Instead of shocks to monetary policy, changes in the interest rate reflect the central bank's endogenous response to changes in the equilibrium real interest rate, which is driven by shocks to the expected growth rate. During booms when growth rates are high, equilibrium real interest rates must be high to clear the goods market. In order to maintain constant inflation, the central bank raises nominal interest rates.

The model can be solved along the same lines as the baseline model. The main difference is that the state variables are now  $\mu_a$  (exogenous) and  $z$  (endogenous). Appendix B shows the details of the solution method.

**Parameter values.** We maintain most of the parameter values from the baseline model. In particular, we keep the same values for  $\gamma$ ,  $\rho$ ,  $\phi$ ,  $\alpha$ ,  $\beta$  and  $\epsilon$ . We again set  $\tilde{\sigma}_a$  to match average real interest rates, which results in  $\tilde{\sigma}_a = 0.073$  and we set  $\tau = 0.146$  to match the average level of  $z$ . We set the inflation target  $\bar{\mu}_p = 2.53\%$  to match average inflation for 1990-2014. We set  $\bar{\mu}_a = 0.01$  as in the baseline model and set  $\lambda = 0.013$  and  $\sigma = 0.024$  to match the standard deviation of LIBOR and the 10-year bond yield. This implies a very persistent and volatile process for the expected growth of the economy, much more so than the data. The goal is to match movements in both short and long interest rates that we observe, and which are central to the mechanism; a full theory of why equilibrium real interest rates move so much is beyond the scope of this exercise. The lower bound  $\mu_a^{min} = \gamma\tilde{\sigma}_a^2 - \beta\rho - \bar{\mu}_p = -0.023$  is set to ensure that it's always possible to attain the inflation target.<sup>19</sup>

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<sup>19</sup>Equilibrium requires a positive nominal interest rate  $i = r + \bar{\mu}_p$ . If the expected growth rate  $\mu_a$  is very negative, the required equilibrium real interest rate could be too negative,  $r < -\bar{\mu}_p$  for some  $(i, z)$ , which would force the central bank to miss its inflation target.

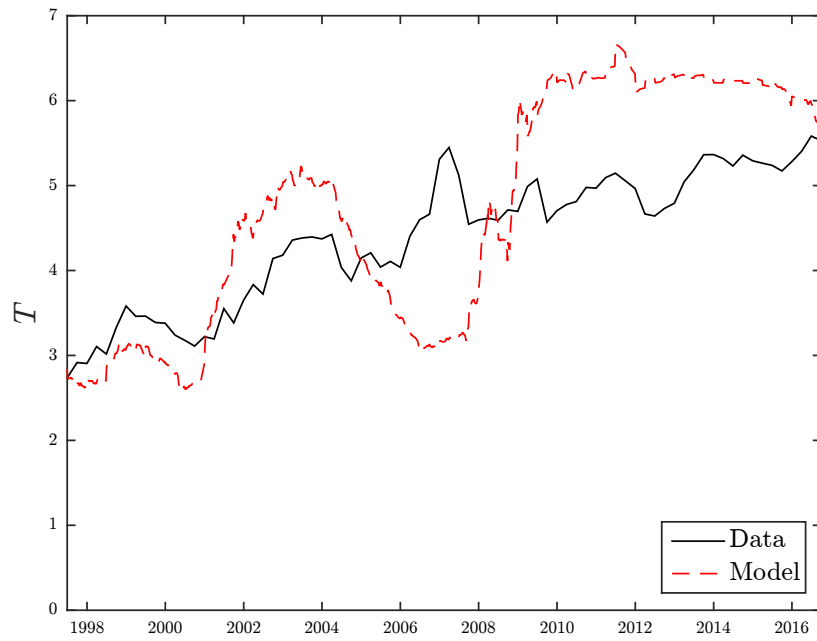


Figure 11: Maturity mismatch of banks in the English et al. (2018) data and in the model, under inflation targeting.

**Results.** Lower growth rates lead to lower equilibrium real rates and, since inflation is constant, to lower nominal rates. Since holding currency is always an option, the nominal interest rate is always positive. Banks' exposure is always negative and quite large, as in the baseline model. At the average values of  $\mu_a$  and  $z$ , a change in the growth rate that induces a 100 basis point rise in the nominal interest rate results in banks losing about 35% of their net worth. The underlying mechanism is the same as in the baseline model and the magnitude of the effect is similar.

Figure 11 compares the time series for  $T$  predicted by the model with the data. For the model values, we take the measured time series for  $i$  and  $z$  and back out the level of  $\mu_a$  in the model that would generate the observed  $i$  given the observed  $z$ . We then plug in the imputed  $\mu_a$  and the measured  $z$  into the function  $T(\mu_a, z)$  produced by the model. Again, the model matches the behavior of  $T$  quite closely, both in terms of average levels and in terms of time pattern. The average  $T$  in the data is 4.4 years; in the model, it's 4.7 years. The correlation between the model and the data is 0.51. We also compare the model predictions with the cross-sectional evidence. In the model an increase of the deposit-to-net-worth ratio of a bank is associated with an increase in the maturity mismatch of 0.55 years. In the data it's 0.43 years.

We conclude from this that the explanatory power of the mechanism does not depend on random monetary policy being the driver of interest rates. Real shocks under an inflation targeting regime have approximately the same effect, as long as they imply similar movement in nominal interest rates.

## 8 Conclusion

Banks' mark-to-market net worth is highly exposed to movements in interest rates and this plays an important role in the transmission of monetary shocks. We propose an explanation for banks' exposure to interest rate risk based on their role as providers of liquidity. Since the spread between (liquid) deposits and (illiquid) bonds rises after the interest rate increases, their exposure to interest rate risk is part of a dynamic hedging strategy. Banks are willing to take large losses after interest rates increase because they expect better investment opportunities looking forward (relative to households). This risk exposure can be achieved with a traditional banking balance sheet with a maturity mismatch between assets and liabilities. Since banks' supply of deposits depends on their net worth, the endogenous response of banks' balance sheets amplifies the effects of monetary shocks on the cost of liquidity.



When we calibrate the model to US data, we find an average maturity mismatch of 3.9 years, compared to 4.4 years in the data. The model also reproduces the time and cross-sectional patterns in the data. The maturity mismatch is larger during periods of low interest rates, and for banks with higher deposit-to-net-worth ratios. This is true both when interest rates are driven by monetary policy shocks and when they are driven by real shocks under an inflation targeting regime. Seen through the lens of our model, banks' exposure to interest rate risk does not constitute risk seeking, but rather a form of insurance, and *increases* with risk aversion.

More generally, our theory has implications for banks' risk exposure beyond interest rate risk. Banks will choose exposure to risks that are correlated with their investment opportunities. The approach in this paper can therefore be useful in studying not only how much, but also what type of risks banks take.

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## Appendix A: Modified Model and Solution Method

**Modified model with taxes and stochastic productivity.** Let  $\tilde{\sigma}$  denote exposure to the productivity shock  $\tilde{B}_t$  and let  $\tilde{\pi}_t$  denote the risk premium for exposure to this shock. Since the model scales linearly with the level of  $a$  we redefine  $\omega$  as total wealth divided by  $a$  and likewise for  $h$ ,  $g$ ,  $k$  and  $q$ .

If  $\tau$  is the tax rate on bankers' wealth, the government budget implies that  $\tau \frac{z_t}{1-z_t}$  is the subsidy rate on households' wealth. The budget constraints thus become, respectively:

$$\begin{aligned} \frac{dn_t}{n_t} &= (r_t - \tau + \sigma_{n,t}\pi_t + \tilde{\sigma}_{n,t}\tilde{\pi}_t - \chi_t\hat{x}_t + \phi s_t) dt + \sigma_{n,t}dB_t + \tilde{\sigma}_{n,t}d\tilde{B}_t \\ \frac{dw_t}{w_t} &= \left( r_t + \tau \frac{z_t}{1-z_t} + \sigma_{n,t}\pi_t + \tilde{\sigma}_{n,t}\tilde{\pi}_t - \chi_t\hat{x}_t + \phi s_t \right) dt + \sigma_{n,t}dB_t + \tilde{\sigma}_{n,t}d\tilde{B}_t \end{aligned}$$

and the HJB equations are, respectively:

$$\begin{aligned} 0 &= \max_{\hat{x}, \sigma_n, \tilde{\sigma}_n, \mu_n} \rho(1-\gamma) \frac{(\xi_t n_t)^{1-\gamma}}{1-\gamma} \left[ \log(\hat{x}n_t) - \frac{1}{1-\gamma} \log((\xi_t n_t)^{1-\gamma}) \right] \\ &+ \xi_t^{1-\gamma} n_t^{1-\gamma} \left( \mu_n + \mu_{\xi,t} - \frac{\gamma}{2} \sigma_n^2 - \frac{\gamma}{2} \sigma_{\xi,t}^2 + (1-\gamma)\sigma_{\xi,t}\sigma_n - \frac{\gamma}{2} \tilde{\sigma}_n^2 - \frac{\gamma}{2} \tilde{\sigma}_{\xi,t}^2 + (1-\gamma)\tilde{\sigma}_{\xi,t}\tilde{\sigma}_n \right) \end{aligned}$$

$$s.t. \quad \mu_n = r_t - \tau + \sigma_n\pi_t + \tilde{\sigma}_n\tilde{\pi}_t + \phi s_t - \hat{x}\chi_t$$

and:

$$\begin{aligned}
0 &= \max_{\hat{x}, \sigma_w, \tilde{\sigma}_w, \mu_w} \rho(1-\gamma) \frac{(\zeta_t w_t)^{1-\gamma}}{1-\gamma} \left[ \log(\hat{x} w_t) - \frac{1}{1-\gamma} \log((\zeta_t w_t)^{1-\gamma}) \right] \\
&\quad + \zeta_t^{1-\gamma} w_t^{1-\gamma} \left( \mu_w + \mu_{\zeta,t} - \frac{\gamma}{2} \sigma_w^2 - \frac{\gamma}{2} \sigma_{\zeta,t}^2 + (1-\gamma) \sigma_{\zeta,t} \sigma_w - \frac{\gamma}{2} \tilde{\sigma}_w^2 - \frac{\gamma}{2} \tilde{\sigma}_{\zeta,t}^2 + (1-\gamma) \tilde{\sigma}_{\zeta,t} \tilde{\sigma}_w \right) \\
s.t. \quad \mu_w &= r_t + \tau \frac{z_t}{1-z_t} + \sigma_w \pi_t + \tilde{\sigma}_w \tilde{\pi}_t - \hat{x} \chi_t
\end{aligned}$$

The first order conditions (19), (23) and (24) are unaffected so formula (27) still applies. The first order conditions for  $\tilde{\sigma}_n$  and  $\tilde{\sigma}_w$  are:

$$\begin{aligned}
\tilde{\sigma}_n &= \frac{\tilde{\pi}_t}{\gamma} + \frac{1-\gamma}{\gamma} \tilde{\sigma}_{\xi,t} \\
\tilde{\sigma}_w &= \frac{\tilde{\pi}_t}{\gamma} + \frac{1-\gamma}{\gamma} \tilde{\sigma}_{\zeta,t}
\end{aligned}$$

The same steps that lead to (27) imply:

$$\tilde{\sigma}_z = \frac{(1-z) \frac{1-\gamma}{\gamma} \left( \frac{\xi_i}{\xi} - \frac{\zeta_i}{\zeta} \right)}{1-z(1-z) \frac{1-\gamma}{\gamma} \left( \frac{\xi_z}{\xi} - \frac{\zeta_z}{\zeta} \right)} \tilde{\sigma}_i$$

and since, by definition,  $\tilde{\sigma}_i = 0$ , this implies  $\tilde{\sigma}_z = 0$ .

Its easy to see from the market clearing conditions that

$$\omega = \frac{k}{\beta \rho} \tag{30}$$

$$h = \frac{k}{\beta} \alpha (1-\beta) l^{\epsilon-1} i^{-\epsilon} \tag{31}$$

and condition (21) still applies.

Using Ito's lemma,

$$\tilde{\sigma}_{\xi} = \frac{\xi_i}{\xi} \tilde{\sigma}_i + \frac{\xi_z}{\xi} z \tilde{\sigma}_z$$

This implies that  $\tilde{\sigma}_{\xi} = 0$  and similarly  $\tilde{\sigma}_{\zeta} = 0$ , so

$$\tilde{\sigma}_n = \tilde{\sigma}_w = \frac{\tilde{\pi}_t}{\gamma}$$

And since  $n = za\omega$ , then  $\tilde{\sigma}_n = \tilde{\sigma}_z + \tilde{\sigma}_a + \tilde{\sigma}_\omega = \tilde{\sigma}_a$ . Therefore:

$$\tilde{\pi}_t = \gamma\tilde{\sigma}_a \quad (32)$$

Replacing the first order conditions in the HJB equations and simplifying, these reduce to:

$$\rho \log(\xi_t) = \rho \log\left(\frac{\rho}{\chi_t}\right) + r_t - \tau - \rho + \phi s_t + \mu_{\xi,t} - \frac{\gamma}{2}\sigma_{\xi,t}^2 + \frac{\gamma}{2}\sigma_{n,t}^2 + \frac{\gamma}{2}\tilde{\sigma}_a^2 \quad (33)$$

$$\rho \log(\zeta_t) = \rho \log\left(\frac{\rho}{\chi_t}\right) + r_t + \tau \frac{z_t}{1-z_t} - \rho + \mu_{\zeta,t} - \frac{\gamma}{2}\sigma_{\zeta,t}^2 + \frac{\gamma}{2}\sigma_{w,t}^2 + \frac{\gamma}{2}\tilde{\sigma}_a^2 \quad (34)$$

The price of capital  $q$  follows the stochastic process:

$$\frac{dq_t}{q_t} = \mu_{q,t}dt + \sigma_{q,t}dB_t + \tilde{\sigma}_{q,t}d\tilde{B}_t \quad (35)$$

but  $\tilde{\sigma}_{q,t} = 0$ , because the TFP shock affects neither  $i$  nor  $z$ . Likewise for  $g$  and  $\psi$  below, we have  $\tilde{\sigma}_{g,t} = 0$  and  $\tilde{\sigma}_{\psi,t} = 0$ .

Arbitrage pricing implies:

$$1 + \mu_a q_t + \mu_{q,t} q_t - r_t q_t = \pi_t \sigma_{q,t} q_t + \tilde{\pi}_t \tilde{\sigma}_a q_t \quad (36)$$

Similarly, the value of government transfers  $g$  follows the stochastic process:

$$dg_t = \mu_{g,t}dt + \sigma_{g,t}dB_t + \tilde{\sigma}_{g,t}d\tilde{B}_t \quad (37)$$

The real flow of transfers is  $\frac{dH_t}{p_t}$  and since  $h_t \equiv \frac{H_t}{a_t p_t}$  and  $\sigma_p = \tilde{\sigma}_p = 0$ , arbitrage pricing of  $g$  implies:

$$h_t (\mu_{h,t} + \mu_a + i_t - r_t) + (\mu_{g,t} + g_t \mu_a) - g_t r_t = \pi_t (\sigma_{h,t} h_t + \sigma_{g,t}) + \tilde{\pi}_t (h_t + g_t) \tilde{\sigma}_a \quad (38)$$

where

$$\frac{dh_t}{h_t} = \mu_{h,t}dt + \sigma_{h,t}dB_t + \tilde{\sigma}_{h,t}d\tilde{B}_t \quad (39)$$

is the stochastic process followed by  $h$ .

Let  $\psi \equiv q + g$  follow the stochastic process:

$$d\psi_t = \mu_{\psi,t}dt + \sigma_{\psi,t}dB_t + \tilde{\sigma}_{\psi,t}d\tilde{B}_t \quad (40)$$

Adding (36) and (38) and rearranging:<sup>20</sup>

$$[1 + h_t (\mu_{h,t} + \mu_a + i_t - r_t)] + [\mu_{\psi,t} + \psi_t \mu_a] - r_t \psi_t = \pi_t [\sigma_{\psi,t} + \sigma_{h,t} h_t] + \tilde{\pi}_t \tilde{\sigma}_a \omega \quad (41)$$

**Solution procedure.** The solution method finds endogenous objects as functions of state variables. We divide the equilibrium objects into three groups. The first are the objects that we can find statically before knowing the value functions:  $s$ ,  $h$ ,  $r$  and  $\psi$ . The second group consists of  $\pi$ ,  $\sigma_n$  and  $\sigma_w$ . These variables can be solved statically if we know  $\xi$  and  $\zeta$ . The last group consists of the two value functions  $\xi$  and  $\zeta$ . We'll express these as a system of differential equations and solve it backwards.

**Objects solved statically.**  $s(i, z)$  comes from (21).  $h(i, z)$  comes from (31). By definition,  $\psi = q + g = \omega - h$ , so  $\psi(i, z)$  follows from subtracting  $h(i, z)$  from (30). Finally, rearranging (41), using  $\psi = \omega - h$  and using (32) to replace  $\tilde{\pi}$  we obtain  $r$ :

$$r = \frac{1 + hi}{\omega} + \mu_a - \gamma \tilde{\sigma}_a^2 \quad (42)$$

**Solving for  $\pi$ ,  $\sigma_n$  and  $\sigma_w$  given  $\xi$ ,  $\zeta$ ,  $s$ ,  $h$ ,  $r$  and  $\psi$**  Suppose we had found any function  $X(i, z)$  that is a function of  $i$  and  $z$ . By Ito's Lemma it follows that the law of motion of  $X$  is:

$$dX(i, z) = \mu_X(i, z) dt + \sigma_X(i, z) dB \quad (43)$$

where the drift and volatility are

$$\begin{aligned} \mu_X(i, z) &= X_z(i, z) \mu_z(i, z) + X_i(i, z) \mu_i(i) \\ &\quad + \frac{1}{2} [X_{zz}(i, z) \sigma_z^2(i, z) z^2 + X_{ii}(i, z) \sigma_i^2(i) + 2X_{zi}(i, z) \sigma_i(i) z \sigma_z(i, z)] \\ \sigma_X &= X_z(i, z) \sigma_z(i, z) z + X_i(i, z) \sigma_i(i) \end{aligned}$$

or, in geometric form:

$$\frac{dX(i, z)}{X(i, z)} = \mu_X(i, z) dt + \sigma_X(i, z) dB \quad (44)$$

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<sup>20</sup>Note that in expressions (36), (38) and (41), the stochastic processes for  $g$  and  $\psi$  are expressed in absolute terms, as set out by (37) and (40). The stochastic processes for  $q$  and  $h$  are expressed in geometric terms, as set out by (35) and (39).

where the drift and volatility are

$$\begin{aligned}\mu_X(i, z) &= \frac{X_z(i, z)}{X(i, z)}\mu_z(i, z) + \frac{X_i(i, z)}{X(i, z)}\mu_i(i) \\ &\quad + \frac{1}{2} \left[ \frac{X_{zz}(i, z)}{X(i, z)}\sigma_z^2(i, z)z^2 + \frac{X_{ii}(i, z)}{X(i, z)}\sigma_i^2(i) + 2\frac{X_{zi}(i, z)}{X(i, z)}\sigma_i(i)z\sigma_z(i, z) \right] \\ \sigma_X &= \frac{X_z(i, z)}{X(i, z)}\sigma_z(i, z)z + \frac{X_i(i, z)}{X(i, z)}\sigma_i(i)\end{aligned}$$

Hence if we know  $\mu_z(i, z)$  and  $\sigma_z(i, z)$  and we know the functions  $\xi$ ,  $\zeta$ ,  $s$ ,  $h$  and  $\psi$  and their derivatives, we know their drifts and volatilities at every point of the state space. Numerically, we approximate the derivatives with finite-difference matrices  $D_i$ ,  $D_z$ ,  $D_{ii}$  and  $D_{zz}$  such that for any set of values of  $\xi$  on a grid, the values of the derivatives on the grid are:

$$\begin{aligned}\xi_i &\approx D_i\xi \\ \xi_z &\approx \xi D_z \\ \xi_{ii} &\approx D_{ii}\xi \\ \xi_{zz} &\approx \xi D_{zz} \\ \xi_{iz} &\approx D_i\xi D_z\end{aligned}$$

The variables  $\pi$ ,  $\sigma_n$  and  $\sigma_w$  can be found as follows. First, in order to apply formulas (43) or (44) we need to know  $\mu_z(i, z)$  and  $\sigma_z(i, z)$ . We get  $\sigma_z(i, z)$  from equation (27). Since  $n = za\omega$  and  $\omega$  is a constant and  $\sigma_a = 0$ , we have that  $\sigma_n = \sigma_z$ . Using the FOC (23), we can solve for

$$\pi = \gamma\sigma_n - (1 - \gamma)\sigma_\xi$$

and using the FOC (24) we can solve for  $\sigma_w$ . Now, to obtain  $\mu_z$ , note that

$$\frac{z}{1 - z} = \frac{n}{w}$$

and therefore

$$\mu_z = (1 - z) [\mu_n - \mu_w + \sigma_w(\sigma_w - \sigma_n)] - \frac{z}{(1 - z)}\sigma_z^2$$



which, using the FOCs, reduces to

$$\mu_z = (1 - z) \left[ (\sigma_n - \sigma_w) \pi + \phi s - \frac{\tau}{1 - z} + \sigma_w (\sigma_w - \sigma_n) \right] - \frac{z}{(1 - z)} \sigma_z^2$$

**Solving for  $\xi$  and  $\zeta$ .** We need to solve (33) and (34). To do so, we define time derivatives such that the equations hold exactly:

$$\dot{\xi} = - \left[ \rho \log \left( \frac{\rho}{\chi} \right) + r - \tau - \rho + \phi s + \mu_\xi - \frac{\gamma}{2} \sigma_\xi^2 + \frac{\gamma}{2} \sigma_n^2 + \frac{\gamma}{2} \tilde{\sigma}_a^2 - \rho \log (\xi) \right] \xi \quad (45)$$

$$\dot{\zeta} = - \left[ \rho \log \left( \frac{\rho}{\chi} \right) + r + \tau \frac{z}{1 - z} - \rho + \mu_\zeta - \frac{\gamma}{2} \sigma_\zeta^2 + \frac{\gamma}{2} \sigma_w^2 + \frac{\gamma}{2} \tilde{\sigma}_a^2 - \rho \log (\zeta) \right] \zeta \quad (46)$$

The algorithm for finding  $\xi$  and  $\zeta$  is as follows.

1. Guess values for  $\xi$  and  $\zeta$  at every point in the state space
2. Compute the derivatives with respect to  $i$  and  $z$  by a finite difference approximation
3. Compute values for  $\pi$ ,  $\sigma_n$  and  $\sigma_w$  at every point in the state space given the guess for  $\xi$  and  $\zeta$ .
4. Compute  $\dot{\xi}$  and  $\dot{\zeta}$  at every point in the state space using (45) and (46)
5. Take a time-step backwards to define a new guess for  $\xi$  and  $\zeta$ . We use a Runge-Kutta 4 procedure.
6. Repeat steps 1-5 until  $\dot{\xi} \approx \dot{\zeta} \approx 0$ .

The condition  $\dot{\xi} \approx \dot{\zeta} \approx 0$  is equivalent to saying that equilibrium conditions hold.

**Finding the stationary distribution.** Once we solve for the equilibrium, this defines drifts and volatilities for the two state variables:  $\mu_i(i, z)$ ,  $\sigma_i(i, z)$ ,  $\mu_z(i, z)$ ,  $\sigma_z(i, z)$ . The density  $f(i, z)$  of the steady state distribution is the solution to the stationary Kolmogorov Forward Equation:

$$0 = - \frac{\partial}{\partial i} [\mu_i(i, z) f(i, z)] - \frac{\partial}{\partial z} [\mu_z(i, z) f(i, z)] \quad (47)$$

$$+ \frac{1}{2} \left( \frac{\partial^2}{\partial i^2} [\sigma_i(i, z)^2 f(i, z)] + \frac{\partial^2}{\partial z^2} [\sigma_z(i, z)^2 f(i, z)] + 2 \frac{\partial^2}{\partial i \partial z} [\sigma_i(i, z) \sigma_z(i, z) f(i, z)] \right)$$

We solve this equation by rewriting it in matrix form.<sup>21</sup> The first step is to discretize the state space into a grid of  $N_i \times N_z$  points and then convert it to a  $N_i N_z \times 1$  vector. Let  $vec(\cdot)$  be the operator that does this conversion. We then convert the differentiation matrices so that they are properly applied to vectors:

$$\begin{aligned}
D_i^{vec} &\equiv I_{N_i} \otimes D_i \\
D_{ii}^{vec} &\equiv I_{N_i} \otimes D_{ii} \\
D_z^{vec} &\equiv M' (I_{N_z} \otimes D_z) M \\
D_{zz}^{vec} &\equiv M' (I_{N_z} \otimes D_{zz}) M \\
D_{iz}^{vec} &\equiv D_i^{vec} D_z^{vec}
\end{aligned}$$

where  $\otimes$  denotes the Kroenecker product and  $M$  is the vectorized transpose matrix such that  $Mvec(A) = vec(A')$ .

Now rewrite (47):

$$\begin{aligned}
& -D_i^{vec} \cdot (diag(vec(\mu_i)) vec(f)) - D_z^{vec} (diag(vec(\mu_z)) vec(f)) \\
& + \frac{1}{2} \left[ \begin{aligned} & D_{ii}^{vec} \cdot (diag(vec(\sigma_i^2)) vec(f)) + D_{zz}^{vec} (diag(vec(\sigma_z^2)) vec(f)) \\ & + 2D_{iz}^{vec} (diag(vec(\sigma_i)) diag(vec(\sigma_z)) vec(f)) \end{aligned} \right] = 0
\end{aligned}$$

and therefore

$$Avec(f) = 0 \tag{48}$$

where

$$\begin{aligned}
A &= -D_i^{vec} \cdot diag(vec(\mu_i)) - D_z^{vec} \cdot diag(vec(\mu_z)) \\
& + \frac{1}{2} [D_{ii}^{vec} \cdot diag(vec(\sigma_i^2)) + D_{zz}^{vec} \cdot diag(vec(\sigma_z^2)) + 2D_{iz}^{vec} \cdot (diag(vec(\sigma_i)) diag(vec(\sigma_z)))]
\end{aligned}$$

Equation (48) defines an eigenvalue problem. We solve it by imposing the additional condition that  $f$  integrates to 1.

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<sup>21</sup>See Achdou et al. (2014) for details on this procedure.

## Appendix B: Solution of the Model with Shocks to the Growth Rate

The steps that lead to equations (30) -(41) are unchanged, except for two differences. First, all objects are functions of  $\mu_a$  and  $z$  instead of  $i$  and  $z$ . For instance:

$$\sigma_z = \frac{(1-z)^{\frac{1-\gamma}{\gamma}} \left( \frac{\xi_{\mu_a}}{\xi} - \frac{\zeta_{\mu_a}}{\zeta} \right)}{1-z(1-z)^{\frac{1-\gamma}{\gamma}} \left( \frac{\xi_z}{\xi} - \frac{\zeta_z}{\zeta} \right)} \sigma_{\mu_a}$$

Second, since  $i$  is not a state variable, we need to solve for it. However, it can simply be replaced by  $i = r + \mu_p$ .

The solution procedure is also to divide the equilibrium objects into three groups. The objects we can solve statically are:  $s$ ,  $h$ ,  $\psi$  and  $r$ . We can then find  $\pi$ ,  $\sigma_n$  and  $\sigma_w$  if we know  $\xi$  and  $\zeta$ . We find  $\xi$  and  $\zeta$  by solving a system of differential equations.

We use the market clearing condition for deposits (21), the market clearing condition for currency (31), equation (42) for  $r$  and the definition  $\omega = \psi + h$  to solve for  $h$ ,  $s$ ,  $r$  and  $\psi$  simultaneously.  $\pi$ ,  $\sigma_n$  and  $\sigma_w$  are obtained in the same way as in the baseline model and the differential equations (45) and (46) still apply and can be solved the same way.

## Appendix C: Projection of Book Values and Net Interest Margins

We assume the bank invests a fraction  $\kappa$  of its assets in instantaneous bonds and a fraction  $(1 - \kappa)$  in geometric bonds of maturity  $\frac{1}{\delta}$ . The yield-to-maturity of a geometric bond bought in state  $(i, z)$  is:

$$i_\delta(i, z) = \delta \left( \frac{1}{p^{GB}(i, z, \delta)} - 1 \right)$$

In the bank's income statement, the bond will accrue this yield for its entire lifetime. If the bond was bought at time  $s$ , its remaining book value at time  $t > s$  is:

$$e^{-\delta(t-s)} p^{GB}(i_s, z_s, \delta)$$

Note that  $p^{GB}(i_s, z_s, \delta)$  is the price at which the bond was purchased, not its current market price.

Let  $b_0^\delta$  be a bank's initial stock of geometric bonds,  $n_t^\delta$  its rate of purchases of these bonds per unit of time and  $b_t^I$  its stock of instantaneous bonds (which are always valued at par). It follows that the book value of the bank's assets will be:

$$a_t^B = b_t^I + e^{-\delta t} b_0^\delta p^{GB}(i_0, z_0, \delta) + \int_0^t e^{-\delta(t-s)} n_s^\delta p^{GB}(i_s, z_s, \delta) ds \quad (49)$$

The bank's total interest income, as shown on its income statement, will be:

$$r_t^B = b_t^I i_t + e^{-\delta t} b_0^\delta p^{GB}(i_0, z_0, \delta) i_\delta(i_0, z_0) + \int_0^t e^{-\delta(t-s)} n_s^\delta p^{GB}(i_s, z_s, \delta) i_\delta(i_s, z_s) ds \quad (50)$$

so the book yield on assets will be:

$$i_t^B = \frac{r_t^B}{a_t^B} \quad (51)$$

In order to compute  $r_t^B$  and  $a_t^B$ , we need to keep track of purchases of geometric bonds over time. Since the bank always invests a fraction  $\kappa$  in short term bonds, we have that the stocks of instantaneous and geometric bonds at time  $t$  are:

$$b_t^S = \kappa (1 + \phi) z_t \quad (52)$$

$$b_t^\delta = (1 - \kappa) \frac{1}{p^{GB}(i_t, z_t, \delta)} (1 + \phi) z_t \quad (53)$$

where  $z_t$  is the bank's mark-to-market net worth, so  $(1 + \phi) z_t$  is the mark-to-market value of the bank's assets. Taking the time derivative of (53):

$$\dot{b}_t^\delta p^{GB}(i_t, z_t, \delta) + b_t^\delta \dot{p}^{GB}(i_t, z_t, \delta) = (1 - \kappa) (1 + \phi) \dot{z}_t \quad (54)$$

and, since geometric bonds mature at a rate  $\delta$ :

$$\dot{b}_t^\delta = -\delta b_t^\delta + n_t^\delta \quad (55)$$

Replacing (55) into (54) and rearranging:

$$n_t^\delta = \frac{(1 - \kappa) (1 + \phi) \dot{z}_t - b_t^\delta \dot{p}^{GB}(i_t, z_t, \delta)}{p^{GB}(i_t, z_t, \delta)} + \delta b_t^\delta \quad (56)$$

The purchase rate  $n_t^\delta$  can be computed directly from an impulse response of  $i_t$  and  $z_t$  (which in turn imply impulse responses for  $p^{GB}(i_t, z_t, \delta)$  and, using (54), for  $b_t^\delta$ ). Replacing (56) in (49), (50) and (51), one can compute the book yield on assets  $i_t^B$ . The net interest margin is just the difference between this book yield and the deposit rate.

## Appendix D: Data Sources

- We take the weekly 6-month LIBOR rate (series USD6MTD156N) and 10-year bond yields (series DGS10) from FRED.
- The wealth measure is “All sectors; U.S. wealth” (series Z1/Z1/FL892090005.Q) from the quarterly Flow of Funds.
- Consumption is “Personal Consumption Expenditures, Nondurable Goods” plus “Personal Consumption Expenditures, Services” from NIPA.
- Total checking deposits and total savings deposits are, respectively “Private depository institutions; checkable deposits; liability” (series Z1/Z1/FL703127005.Q) and “Private depository institutions; small time and savings deposits; liability” (series Z1/Z1/FL703131005.Q) from the quarterly Flow of Funds.
- Total net worth of banks is the difference between “Private depository institutions; total liabilities and equity” (series Z1/Z1/FL704194005.Q) and “Private depository institutions; total liabilities” (series Z1/Z1/FL704190005.Q) from the quarterly Flow of Funds.
- The interest rates on checking and savings deposits are, respectively, the average rates on “Interest checking 0k” and “Money market deposit 10k” reported by Drechsler et al. (2017).
- Inflation, used to construct real interest rates, is the CPI inflation from the BLS

**Maturity mismatch data** In order to construct the measures of maturity mismatch, we take data from the quarterly Reports of Condition and Income (“Call Reports”) filed by banks regulated by the Federal Reserve System, Federal Deposit Insurance Corporation, and the Comptroller of the Currency (almost all U.S. commercial banks) from 2001:Q1 to 2016:Q3. We then follow the procedure used by **English et al. (2018)**:

- The maturity mismatch measure  $T_{jt}$  for bank  $j$  in quarter  $t$ , is:

$$T_{jt} = \Xi_{jt}^A - \Xi_{jt}^L$$

- $\Xi_{jt}^A$  is the weighted average asset repricing/maturity period, calculated according to:

$$\Xi_{jt}^A = \frac{\sum_k m_A^k A_{jt}^k}{\sum_k A_{jt}^k}$$

where  $k$  indexes the 26 interest-earning asset categories with information about maturity or repricing (RCFDA549-RCFDA562, RCONA549-RCONA562, RCONA564-RCONA575).  $A_{jt}^k$  is the value of asset category  $k$  and  $m_A^k$  denotes the estimated average maturity of that category. These 26 asset categories account, on average for 92.5% of the total interest-earning assets per bank. The underlying assumption on this calculation is that the average maturity of those assets for which no repricing or maturity information is available is the same as the average maturity of the 26 categories.

- $\Xi_{jt}^L$  is the weighted average liability repricing/maturity period, calculated according to:

$$\Xi_{jt}^L = \frac{\sum_k m_L^k L_{jt}^k}{L_{jt}^{IE}}$$

where  $k$  indexes the 11 interest-earning liability categories with information about maturity or repricing (RCON6810, RCON0352 RCONA579-RCONA582, RCONA584-RCONA587, RCON2215).  $L_{jt}^k$  is the value of liability category  $k$  and  $m_L^k$  denotes the estimated average maturity of that category. These 11 asset categories account, on average for 92.9% of total liabilities per bank.

- The estimated maturity of each asset and liability category ( $m_L^k$  and  $m_A^k$ ) is set to the midpoint of that category’s maturity or repricing range specified on the Call Report. For example, RCFDA550, the asset category of US Treasuries from 3 months to 12 months, is assigned a maturity of 7.5 months. Long-term assets with no endpoint are assigned the same maturity as English et al. (2018). That is, asset and liability categories labeled as “three years and more” are assigned a duration of 5 years, and those labeled as “fifteen years and more” are assigned a

duration of 20 years.

- From the original 475,220 bank-quarter observations extracted from the Call Reports, all bank-quarter observations with zero loans and leases or zero liabilities were eliminated, leaving us 467,620 observations that form our bank-quarter data set.
- For the period 2001 to 2016, we compute an asset-weighted median  $T_t$  for each quarter from 2001 to 2016. For 1997-2000 (where we don't have detailed bank-level data), we take the asset-weighted median measure directly from English et al. (2018), who kindly shared their data with us.
- We construct the deposit leverage of each bank  $\phi_j$  in the following manner:
  - Bank's net worth ( $n_{jt}$ ) is the dollar value of total assets (RCON2170 for banks that don't have a foreign office and RCFD2170 for those that do) minus total liabilities (RCON2948 and RCFD2948 respectively).
  - Bank's deposits ( $d_{jt}$ ) is the sum of transaction accounts (RCON2215) and saving accounts (*RCON6810* and *RCON0352*)
  - Deposit leverage  $\phi_{jt}$  is just  $\frac{d_{jt}}{n_{jt}}$
  - An extra filter was used to screen for extreme observations. We trimmed the variable  $\phi_{jt}$  above the 99th percentile and below the 1st percentile of its distribution over the entire set of observations.
  - For each of the 10,351 banks in the sample, we construct  $\phi_j$  and  $T_j$  by taking the time average over the period 2001-2016.